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Medicaid Discontinuity among Adults with Cardiovascular Disease or High Risk Conditions: Associations with Medication Adherence and Health Care Utilization

Abstract

BACKGROUND: Medicaid coverage among adults is often characterized by discontinuity – loss of Medicaid coverage, and churning, or entering and exiting Medicaid – over short durations. Little is known about the impact of having discontinuous Medicaid coverage on access to care, preventive care, pharmacotherapy, primary care, and hospitalizations, among non-elderly adults with cardiovascular disease (CVD) or high-risk conditions.

OBJECTIVES: This dissertation employed a three empirical research papers approach to pursue the following aims: (1) characterize the adult subpopulations with CVD or conditions placing them at high risk for CVD who lack continuous Medicaid coverage, and examine the characteristics associated with Medicaid discontinuity, (2) examine associations between Medicaid discontinuity, medication adherence and medication utilization, and (3) examine the associations between Medicaid discontinuity, access to care, preventive care, primary care visits, and hospitalizations.

METHODS: This was a retrospective comparative analysis of the 2002–2011 Medical Expenditure Panel Survey employing a repeated cross-sectional study design. Study sample included adults aged 18–64 years diagnosed with ≥ 1 CVD (defined as acute myocardial infarction, coronary artery disease, congestive heart failure, peripheral and visceral atherosclerosis, or stroke) or high-risk conditions for CVD (defined as hypertension, lipid disorders, diabetes, or chronic kidney disease) who reported having Medicaid coverage any time during survey year. Individuals having continuous, full-year Medicaid coverage (Continuous Medicaid) were compared to those with (Discontinuous–Uninsured), and separately to those with (Discontinuous–Insured). Associations between Medicaid discontinuity, access to care, and preventive care were estimated using multivariate logistic regression. Medication adherence, measured as medication possession ratio (MPR) with adequate adherence being considered at $MPR > 0.8$, was estimated using multivariate logistic regression. Medication utilization, measured as the number of all-cause, and disease-specific prescription drug fills, was estimated using multivariate negative binomial regression. Four health care services utilization outcomes – inpatient, emergency room (ER), hospital outpatient, and office-based physician visits – measured as both, number of all-cause and number of disease-specific visits, were estimated using either zero-inflated negative binomial regression or negative binomial regression depending on the distribution of the outcome of interest.

RESULTS: Overall, 31.8% of adults with CVD or high-risk conditions for CVD had coverage, majority of whom (23.5%) belonged to the Discontinuous–Uninsured group. Of those who had Medicaid at the beginning of the year, only 21.9% of the Discontinuous–Uninsured, and 8% of the Discontinuous–Insured still had Medicaid by the year end. Male gender, minority race/ethnicity, receiving disability benefits or participating in a federal assistance program, Medicaid managed care enrollment, and diagnosis of respiratory illnesses were the characteristics associated with lower odds of Medicaid discontinuity, whereas being married, residing in the South, having higher income, or education, being employed, and having fair to poor perceived health status were associated with higher odds of Medicaid discontinuity. Overall adherence to commonly prescribed therapeutic medication classes, measured as average MPR, was not significantly different between the Continuous Medicaid and the two discontinuous coverage groups, whereas examination of class-specific adherence yielded mixed results. Discontinuous Medicaid coverage was associated with significantly lower all-cause and disease-specific prescription drug utilization among both the discontinuous Medicaid groups. Medicaid discontinuity was associated with poor access to care, and higher disease-specific inpatient and ER hospitalizations among both the discontinuous Medicaid coverage groups. Additionally, among the Discontinuous–Uninsured, Medicaid discontinuity was associated with lower odds of routine medical checkup, lower all-cause primary care

office visits, and higher disease-specific hospital outpatient visits.

CONCLUSION: Among non-elderly adults with CVD or high-risk conditions, for CVD having discontinuous Medicaid coverage was found to be associated with poor access to care and preventive care, poor adherence to certain medication classes and lower utilization of prescription medications, higher hospitalizations for CVD or associated conditions, and lower primary care office visits. Disruptions in and loss of Medicaid coverage among adults with CVD or high-risk conditions may lead to negative health outcomes due to the disruptions in continuity of care and inability to appropriately manage these disease conditions. This research provides strong support for implementation of policies to stabilize Medicaid coverage and reduce Medicaid discontinuity among individuals with CVD or high-risk conditions for CVD, such as the 12 month continuous Medicaid eligibility provisions currently in place for low-income children. Such policies will greatly improve some of the adverse access to care, preventive care, pharmacotherapy, and medical care outcomes observed in this study. Vulnerable populations with chronic, debilitating conditions may benefit from such policies to a greater extent compared to the overall low-income adult population. Simultaneously, reenrollment and outreach strategies may need to be more efficiently implemented to ensure individuals who are eligible for Medicaid, continue to remain enrolled in Medicaid. Such enabling strategies employed by Medicaid managed care organizations may be adopted by State Medicaid agencies to ensure greater continuity in Medicaid coverage for low-income vulnerable populations with chronic and debilitating diseases.

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**Medicaid Discontinuity among Adults with Cardiovascular Disease or High Risk
Conditions: Associations with Medication Adherence and Health Care Utilization**

A Dissertation
Presented for
The Graduate Studies Council
The University of Tennessee

In Partial Fulfillment
Of the Requirements for the Degree
Doctor of Philosophy
From The University of Tennessee

By
Kiraat Divyang Munshi
December 2015

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DEDICATION

To the Almighty for the invisible hand that has guided me throughout my life.

To my late grandparents for their blessings.

To my family for their love and unwavering support.

To my extended family and friends and the unsung heroes who have helped me in life.

To all the Average Joes and Janes out there.

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ABSTRACT

BACKGROUND: Medicaid coverage among adults is often characterized by discontinuity – loss of Medicaid coverage, and churning, or entering and exiting Medicaid – over short durations. Little is known about the impact of having discontinuous Medicaid coverage on access to care, preventive care, pharmacotherapy, primary care, and hospitalizations, among non-elderly adults with cardiovascular disease (CVD) or high-risk conditions.

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RESULTS: Overall, 31.8% of adults with CVD or high-risk conditions for CVD had < 12 months of Medicaid coverage, majority of whom (23.5%) belonged to the Discontinuous–Uninsured group. Of those who had Medicaid at the beginning of the year, only 21.9% of the Discontinuous–Uninsured, and 8% of the Discontinuous–Insured still had Medicaid by the year end. Male gender, minority race/ethnicity, receiving disability benefits or participating in a federal assistance program, Medicaid managed

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CONCLUSION: Among non-elderly adults with CVD or high-risk conditions, for CVD having discontinuous Medicaid coverage was found to be associated with poor access to care and preventive care, poor adherence to certain medication classes and lower utilization of prescription medications, higher hospitalizations for CVD or associated conditions, and lower primary care office visits. Disruptions in and loss of Medicaid coverage among adults with CVD or high-risk conditions may lead to negative health outcomes due to the disruptions in continuity of care and inability to appropriately manage these disease conditions. This research provides strong support for implementation of policies to stabilize Medicaid coverage and reduce Medicaid discontinuity among individuals with CVD or high-risk conditions for CVD, such as the 12 month continuous Medicaid eligibility provisions currently in place for low-income children. Such policies will greatly improve some of the adverse access to care, preventive care, pharmacotherapy, and medical care outcomes observed in this study. Vulnerable populations with chronic, debilitating conditions may benefit from such policies to a greater extent compared to the overall low-income adult population. Simultaneously, reenrollment and outreach strategies may need to be more efficiently implemented to ensure individuals who are eligible for Medicaid, continue to remain enrolled in Medicaid. Such enabling strategies employed by Medicaid managed care organizations may be adopted by State Medicaid agencies to ensure greater continuity in Medicaid coverage for low-income vulnerable populations with chronic and debilitating diseases.

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LIST OF ABBREVIATIONS

ACE	Angiotensin converting enzyme inhibitors
AHRQ	Agency for Healthcare Research and Quality
AIDS	Acquired immunodeficiency syndrome
ARB	Angiotensin II receptor blockers
CAD	Coronary artery disease
CCC	Clinical classification code
CCI	Charlson Comorbidity Index
CHF	Congestive heart failure
CHIP	Children's Health Insurance Program
CI	Confidence interval
CKD	Chronic kidney disease
CMS	Centers for Medicare and Medicaid
CVD	Cardiovascular disease/cardiovascular diseases
ED	Emergency department
ER	Emergency room
FPL	Federal poverty level
GED	General Educational Development
HC	Household Component
HIV	Human immunodeficiency virus infection
HTN	Hypertension
IRR	Incidence Rate Ratio
MCO	Managed care organization

MEPS	Medical Expenditure Panel Survey
MPR	Medication Possession Ratio
MSA	Metropolitan Statistical Area
NCHS	National Center for Health Statistics
OR	Odds ratio
PDC	Proportion of days covered
PPACA	Patient Protection and Affordable Care Act
PVA	Peripheral vascular disease
SE	Standard error
SSI	Supplemental Security Income
TANF	Temporary Assistance for Needy Families
WHO	World Health Organization

CHAPTER 1. INTRODUCTION

INSTABILITY IN HEALTH INSURANCE

The health care system in the United States (U.S.) has widely recognized to be “fragmented”.¹ The systemic misalignment of incentives, and lack of coordination between physician, and hospitals, leading to poor care coordination, inefficient allocation of resources, contribute to this narrative. This in turn greatly hampers the quality of care that patients receive. Likewise, the health insurance system in the U.S. too is fragmented with most individuals covered by employer-sponsored private health insurance (48%), and the poor, elderly, and disabled population groups being covered by government-sponsored public health insurance (roughly 31%), with the remaining either purchasing individual private health plans (just 5% of the total population), or remaining uninsured, which estimates peg at 15% of the total population, roughly the same as those on Medicaid and Medicare.² Rules and regulations are set separately for both private health insurance by employers, and for public health insurance by government, regarding eligibility for the insurance plans paid for, the specific covered benefits and cost sharing.

The principal requirement for employer-sponsored insurance is employment. Losing employment, changing jobs, moving to different geographic regions, and starting one’s own venture, are circumstances that can lead to losing employer-sponsored insurance coverage or gaps in insurance coverage. The number of individuals that lose insurance resulting from loss of employment increased significantly during times of economic recession.^{3,4} As a cost-cutting measure, employers may change the benefit structure of the plans by requiring their employees to pay more through higher premiums, deductibles or co-pays, and also by cutting the numbers and/or types of services covered. Such measures may make health insurance expensive for low-income employees and those who cannot afford such plan structures. When an individual employee receives employment-sponsored coverage, his/her entire family is covered under the insurance plan. Thus, employee spouses and children of may also experience insurance instability as well due to circumstances such as changes in benefit structure where spouse coverage is dropped, loss of coverage or employment, retirement, death or other circumstances.

Government-sponsored public health insurance too has its own set of eligibility criteria and enrollment procedures, and requires individuals to be aware of them to determine whether they are eligible or not, and if they are, to provide the appropriate documentation to enroll for these insurance plans, and maintain eligibility; this is especially true for Medicaid and Children’s Health Insurance Plan (CHIP) that are also referred to as means-tested public health insurance programs.⁵ The eligibility requirements and enrollment procedures for Medicaid and CHIP differ from state-to-state, and so any change in residence across states may lead to loss of insurance coverage resulting from either ineligibility or failure to enroll successfully. Traditionally, before the passage of the 2010 Patient Protection and Affordable Care Act (PPACA), not only did the income eligibility thresholds vary greatly between the states, but they were also different for different population groups, such as more generous coverage limits for

children and pregnant women, less generous coverage limits for non-elderly adults with children, and no coverage for non-elderly adults without any children.⁶ Small fluctuations in income levels for covered non-elderly adults would end up disqualifying individuals with incomes at or near eligibility thresholds. As with employer-sponsored health insurance, changes or loss in Medicaid coverage for one member of the family impacted the continuity in Medicaid for the entire family. Also, changes in family size and compositions would also impact eligibility for these public insurance programs, since yearly poverty levels that are established take into account these factors. The rules and regulations for eligibility and enrollment between public and private health insurance are different, and therefore, individuals who have to transition between one system to another also experience gaps in insurance coverage, either due to failure to follow enrollment procedures, or being unaware of such an option, or even due to the stigma of enrolling for Medicaid or other low-income public programs.⁶ It is broadly due to these diverse factors in this fragmented health insurance system, that individuals and families experience insurance instability and continue to swing into and out of various types of health insurance coverage.

INSTABILITY IN MEDICAID AMONG LOW-INCOME ADULT POPULATION

Medicaid, the nation's primary public health insurance program for the low-income population, provides much-needed health insurance coverage to an estimated 67 million individuals who meet the eligibility criteria, which include having incomes and assets below a certain threshold level determined by each State.⁷ These include low-income children, adults, elderly, and disabled who do not have access to private health insurance, and many of whom are not eligible for any other source of coverage. Medicaid however, has been described as a leaky sieve. Inefficient administrative hurdles and ponderous periodic paperwork requirements, result in discontinuity in Medicaid coverage, characterized by either loss of coverage or "churning" – frequent transitions into and out of various sources of coverage.^{8,9} Such instability often plagues Medicaid-eligible individuals, leading to both gaps in coverage and interruptions in the continuity of care. Families may have to find new providers or change their existing health treatments due to different provider networks or different services covered in their new health plans, or may end up forgoing needed care after becoming uninsured. Some research has been conducted to examine Medicaid instability and its impact on various outcomes, although the research in this area is not exhaustive.

Harman and colleagues examined Utah Medicaid claims data from 1990-1994 to study the impact of Medicaid interruptions on inpatient psychiatric services use. They found that interruptions in Medicaid were associated with an average of 0.63 more psychiatric hospitalizations per beneficiary, and 8.3 more psychiatric hospitalization days over the study period.¹⁰ Using the Ohio Medicaid claims from 1992-1999 Koroukian examined the impact of length of enrollment in Medicaid on use of screening mammography among women aged 40-64 years. They found that with each additional year of Medicaid enrollment, the proportion of women receiving screening mammography increased significantly, and [Odds Ratio (OR): 1.59; 95% Confidence

Interval (CI): 1.57–1.60], and the mean number of annual mammography exams increased from 0.08/year among women enrolled for less than a year to 0.26/year among women enrolled for more than 7 years.¹¹ Continuous Medicaid enrollment also was shown to have a positive impact on cervical cancer prevention.¹² Harman and colleagues analyzed Florida Medicaid claims data from 1999-2002 to examine the impact of Medicaid interruption on depressed beneficiaries' health care utilization and expenditures. They found that inpatient episodes, the length of inpatient stay, number of emergency department (ED) visits, and total Medicaid expenditures increased significantly ($P<0.001$) in the three-month period after an interruption in coverage of more than one month, indicating a significant negative impact of Medicaid instability among low-income elderly with mental illness.¹³ More recently, studies that have examined the impact of interruptions in Medicaid using state-specific data have found that instability in Medicaid is associated with higher risk for hospitalization for ambulatory care-sensitive conditions, and increase in overall program expenditures due to higher inpatient and ED utilization.¹⁴ In the only study using nationally representative data, Banerjee and colleagues (2010) examined the Medical Expenditure Panel Survey (MEPS) 2000-2004 to study the impact of Medicaid instability on health care utilization, and found that individuals with discontinuous Medicaid coverage had higher inpatient and ED utilization, but lower prescription drug utilization annually.¹⁵

The literature examining the impact of lack of insurance on health care outcomes in general, and disease-specific outcomes in particular is exhaustive with studies highlighting the adverse impact of lacking health insurance. The literature examining the impact of gaps in health insurance is less comprehensive, however, for the most part this literature has established that even short gaps in insurance coverage may lead to negative outcomes among individuals with chronic conditions. Although these studies have assessed the impact of insurance gaps on patients with numerous conditions including cancer, depression/anxiety and schizophrenia, and diabetes, fewer studies have focused on patient with cardiovascular disease (CVD) and CVD-associated conditions. Studies using more advanced statistical analyses such as zero-inflated linear models, difference-in-differences analyses, and instrumental variable regression analyses need to be conducted more to account for endogeneity problems that may confound the association between continuation and discontinuation in Medicaid and other insurance programs and major clinical outcomes. Studies examining Medicaid instability have mostly been conducted using data on a single state's Medicaid program, and there is a paucity of studies (could also be due to lack of national data) that examine nationally representative estimates of churning in Medicaid and its impact on health care outcomes, especially among individuals diagnosed with chronic disease conditions. Although medication adherence/compliance as an area of research has been thoroughly examined with several high-quality robust analyses of adherence, and its impact of health care utilization, research examining the impact of churning in insurance in general and Medicaid in particular is lacking. It is imperative therefore, to examine the impact of Medicaid discontinuity on pharmacotherapy outcomes, such as medication adherence and medication utilization.

CARDIOVASCULAR DISEASE OR ASSOCIATED HIGH-RISK CONDITIONS AMONG LOW-INCOME POPULATIONS

CVD is the leading cause of death in the U.S. accounting for approximately 1 in every 4 deaths being attributed to them.¹⁶ Cardiovascular complications resulting from high blood pressure, lipid disorders, and other risk factors have been known to exert a substantial burden leading to disability, morbidity, and mortality.^{17,18} The economic burden due to CVD has been estimated to exceed \$300 billion.¹⁹ Estimates peg the prevalence of cardiovascular disease (CVD), including heart failure and stroke, among low-income adult population at 28%.²⁰ Additionally, many individuals with CVD that have resulted from worsening of other chronic conditions such as hypertension, diabetes, and hyperlipidemia. High blood pressure (24%), high cholesterol (17%), and high blood sugar level (9%) are some of the prevalent risk factors or high-risk conditions for CVD.^{21,22} For instance, 82% of low-income diabetic adults, and 74% of low-income adults with CVD reported one or more additional chronic condition, which indicates the complex disease burden and health care needs in the CVD or high risk population. Additionally, spending for nonelderly adult Medicaid enrollees with CVD and many of the high-risk conditions for CVD ranged from \$9,694 - \$13,490 per capita, which was significantly higher than their counterparts without these chronic conditions, and once again reflected the complex chronic disease burden among these diseased subpopulations.²⁰

Low-income populations with CVD or conditions placing them at high risk for CVD, such as acute myocardial infarction, congestive heart failure, coronary artery disease, peripheral and visceral atherosclerosis, stroke, hypertension, diabetes, hyperlipidemia, and chronic kidney disease, would face tremendous challenges in managing these complex chronic conditions. CVD and CVD-associated conditions require appropriate and uninhibited access to medical care, preventive care, prescription medications, and periodic primary and specialty care visits for prevention and treatment. Failure to appropriately manage these chronic conditions may exacerbate the symptoms and lead to worsening of health leading, resulting in unplanned hospitalizations and emergency room admissions. Insurance coverage is a critical component and an important factor in ensuring access to and utilization of needed medical care for appropriate disease management. Medicaid, as mentioned above is a critical safety-net health insurance program for low-income populations that provides these vulnerable populations the ability to seek needed medical care. Interruptions and discontinuity in Medicaid coverage may especially have a lead to significantly poor health outcomes among these vulnerable populations, and therefore policies to stabilize Medicaid coverage among low-income populations with CVD or high-risk conditions for CVD may be even more critical to ensuring continuity of care and appropriate disease management among these individuals. While need appropriate medical care, pharmacotherapy is also an important component of disease management for these individuals. Prescribed medicines not only need to be filled in a timely manner as directed, but adhering to the appropriate prescribed medication regimen is also critical to observe positive therapeutic effects of these medications. After reviewing the literature for the definitions for CVD and conditions which are high risk factors, and in

consultation with a cardiologist to restrict the number of these conditions to those requiring appropriate and prolonged disease management in the outpatient and community settings, the present study has considered the CVD and high-risk conditions for CVD listed above as the index chronic conditions.^{17,23-26} The rationale being that continuity of Medicaid coverage may be even more critical among this subpopulation of the non-elderly adult Medicaid population, and the adverse outcomes due to Medicaid discontinuity may be more detrimental among these individuals.

RESEARCH AIMS

Using nationally representative data on non-institutionalized, non-elderly adults diagnosed with one or more CVD or high-risk conditions for CVD who reported having Medicaid at any point of time during the year, this dissertation aims to investigate the associations between instability in Medicaid coverage within a year, and access to care, preventive care, prescription drug adherence and utilization, primary care visits, and inpatient and ER hospitalizations.

Aim 1: To Examine the Individual Characteristics Associated with Medicaid Discontinuity

- 1) Are the demographic, socioeconomic, Medicaid-, and health-related characteristics significantly different between those who have continuous Medicaid coverage and their counterparts with discontinuous Medicaid coverage during the year?
- 2) Which characteristics are significantly associated with Medicaid discontinuity in this study population?

Aim 2: To Examine the Associations between Medicaid Discontinuity, Medication Adherence and Medication Utilization

- *Hypothesis 2a:* Medicaid discontinuity is associated with poor adherence to commonly prescribed therapeutic classes of medications used to treat CVD or high-risk conditions for CVD
- *Hypothesis 2b:* Medicaid discontinuity is associated with lower all-cause and disease-specific prescription drug utilization

Aim 3: To Examine the Associations between Medicaid Discontinuity, Access to Care, Preventive Services Use, and Utilization of Health Care Services

- *Hypothesis 3a:* Medicaid discontinuity is associated with poor access to care
- *Hypothesis 3b:* Medicaid discontinuity is associated with a lower likelihood of adequate preventive services use

- *Hypothesis 3c*: Medicaid discontinuity is associated with lower all-cause and disease-specific hospital outpatient visits, and office-based physician visits
- *Hypothesis 3d*: Medicaid discontinuity is associated with higher all-cause and disease-specific inpatient hospitalizations and ER hospitalizations

CHAPTER 2. CHARACTERISTICS ASSOCIATED WITH MEDICAID DISCONTINUITY IN ADULTS WITH CARDIOVASCULAR DISEASE OR ASSOCIATED HIGH-RISK CONDITIONS

BACKGROUND

Cardiovascular disease (CVD) and CVD associated conditions (e.g. hypertension, lipid disorders, and diabetes) exert substantial burden due to disability, morbidity, and mortality. CVD is the major cause of death in the U.S., accounting for 1 in 4 deaths being attributed to them.¹⁶ Economic burden due to these debilitating chronic conditions is estimated to exceed \$300 billion annually.¹⁹ Continuity of care is essential to for appropriate management of these disease conditions, and having uninterrupted access to health insurance is critical to ensure this continuity of care.^{27,28} Medicaid, the nation's primary public health insurance program for the low-income population, provides much-needed health insurance coverage to an estimated 67 million individuals at any given point in time.⁷ Estimates indicate that over 28% of the Medicaid population at or below 138 percent of the federal poverty level (FPL) are diagnosed with CVD or high-risk conditions for CVD.^{29,30}

Medicaid however, has been described as a leaky sieve.⁸ Inefficient administrative hurdles and ponderous periodic paperwork requirements, result in discontinuity in Medicaid coverage, characterized by wither loss of coverage or “churning” – frequent transitions into and out of various sources of coverage.^{8,9} Such instability often plagues Medicaid-eligible individuals, leading to both gaps in coverage and interruptions in the continuity of care. Families may have to find new providers or change their existing health treatments due to different provider networks or different services covered in their new health plans, or may end up forgoing needed care after becoming uninsured.

The consequences of lack of insurance have been well-documented. Ayanian and colleagues reported that Americans who were uninsured for one year or more were more likely to have foregone routine clinical checkup in the past 2 years, and this pattern was observed among individuals diagnosed with several chronic conditions such as hypertension, hyperlipidemia, HIV/AIDS, among others. Long-term and short-term uninsured (uninsured for less than one year), also did not receive recommended cancer screening, and diabetes care among others.³¹ Recent estimates indicate that these trends have continued into the first decade of the 21st century. Fields and colleagues analyzed the Medical Expenditure Panel Survey (MEPS) data from 2006-2010 to examine the impact of insurance stability and residence on several health care utilization measures. They found that discontinuously insured individuals had more ED visits, fewer physician office visits, and fewer prescription drug fills among others, compared to those with continuous insurance coverage.³² Several other studies examining Medicaid data have also reported adverse impacts on access to care, and health outcomes due to gaps in Medicaid coverage.^{2,10-14}

Medicaid discontinuity, due to either becoming ineligible or “drop-out” despite meeting the eligibility criteria, has been well documented in children.³³⁻³⁵ Some research has been conducted to examine Medicaid instability among adults, although this area of research is not exhaustive. For instance, estimates peg the rate of adult disenrollment from Medicaid to range from 21–33% annually.³⁶⁻³⁸ Sommers examined the predictors of loss of Medicaid and found that young age, male gender, and Hispanic ethnicity were significant risk factors for coverage discontinuity.³⁸ However, there have been no estimates of Medicaid instability among non-elderly adults with CVD or high-risk conditions for CVD among whom, the impact of such coverage discontinuity could be detrimental. Including population groups diagnosed with and managing these chronic conditions in policy discussions to improve Medicaid retention among adults is important, as these groups could benefit the most from coverage stability and the resulting continuity in disease management.

The present study aims to: 1) obtain national estimates of Medicaid coverage transitions and discontinuity among non-elderly adults with CVD or high-risk conditions for CVD, 2) characterize the population groups that have continuous and discontinuous Medicaid coverage within a year, and 3) examine the individual characteristics associated with Medicaid discontinuity in this study population.

METHODS

The present study is a retrospective cross-sectional analysis designed to compare the characteristics of non-elderly adults diagnosed with CVD or high-risk conditions for CVD with continuous full-year Medicaid coverage versus those with discontinuous Medicaid coverage. The study seeks to determine the individual characteristics associated with Medicaid discontinuity in this CVD population. Medicaid, being a means tested source of public health insurance, is not a permanent source of insurance coverage. As discussed earlier, the requirements of periodic reporting of income, assets, family status and residence, apart from changes in one or more of these circumstances are the major sources of instability in Medicaid coverage which may impact access to care and preventive services use for individuals with CVD or high-risk conditions for CVD. In order to capture this instability in the study population, we first describe the prevalence of these chronic conditions in this study population. We then examine in detail, the demographic, socioeconomic, Medicaid-, and health-related characteristics associated with Medicaid discontinuity in this study, and control for these covariates in the multivariate models to determine the individuals characteristics associated with Medicaid discontinuity in this CVD population.

Data Source

This study analyzed the Medical Expenditure Panel Survey (MEPS). MEPS is a set of large-scale surveys of families, individuals, their medical providers and employers across the United States. It is jointly sponsored by the Agency for Healthcare Research

and Quality (AHRQ) and the National Center for Health Statistics (NCHS), and has been conducted annually since 1996. It has two major components; the Household Component (HC), and the Insurance Component, with data on medical providers in the HC being supplemented by the Medical Provider Component. MEPS collects detailed information from a nationally representative sample of the civilian noninstitutionalized population of the U.S. on health services utilization and health expenditures, insurance coverage, and sources of payment.³⁹

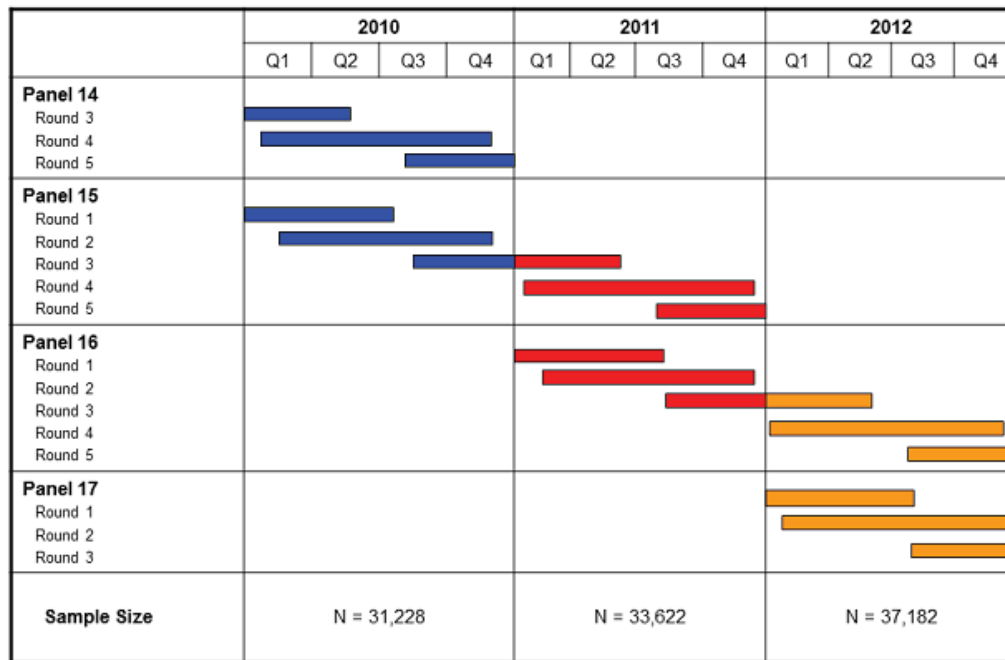
When obtaining national estimates from surveys such as MEPS, appropriate sample size is critical to obtain reliable estimates, since estimates of some population subgroups may vary from year to year. AHRQ suggests a minimum of 100 unweighted participants per cell for producing reliable national estimates. The advantage of survey such as MEPS is that they allow for several years' of data to be pooled together in order to increase the sample size and improve the precision of estimates. The pooling of the different years of data was conducted in accordance with the guidelines provided by MEPS, and data for the years 2002 – 2011 were pooled together to allow for a sufficiently large sample size to ensure reliable estimates.

MEPS Sampling Plan

MEPS has an overlapping panel design. Two separate panels of respondents are interviewed simultaneously during a calendar year. A new panel of sample households is selected each year, and data for each panel are collected for two calendar years. The two years of data for each panel are collected in five rounds of interviews that take place over a two and a half year period. This provides continuous and current estimates of health care expenditures at both the person and household level for two panels for each calendar year. **Figure 2-1** illustrates this overlapping panel design in detail with an example of Panels 15 and 16.

Study Population

The study population consisted of individuals diagnosed with CVD or high-risk conditions for CVD. CVD was defined as any diagnosis of acute myocardial infarction (Clinical Classification code (CCC): 100), coronary artery disease (CCC: 101), congestive heart failure (CCC: 108), peripheral and visceral atherosclerosis (CCCs: 114, 115, 116), stroke (CCCs: 109, 112). High-risk conditions for CVD were defined as hypertension (CCC: 098), lipid disorders (CCC: 053), diabetes (CCC: 049, 050), and chronic kidney disease (CCC: 158). CCC aggregates conditions and procedures into mutually exclusive and clinically homogeneous categories, using Clinical Classification Software.⁴⁰ The study population included respondents who were: (1) aged 18 to 64 years, (2) self-reported either being diagnosed with CVD or high-risk conditions for CVD in their first round of the MEPS survey, or reported having an event (an inpatient, emergency, outpatient, or office-based provider visits, or a prescription medication) associated with one or more of the aforementioned disease conditions at any point of time



N is equal to the number of people with a positive person weight on the file.

Figure 2-1. MEPS Overlapping Panel Design

Reprinted with permission. MEPS-HC Sample Design and Collection Process. Agency for Healthcare Research and Quality, Rockville, Md.

http://www.meps.ahrq.gov/survey_comp/hc_data_collection.jsp. Accessed on April 6, 2015.⁴¹

during the survey, (3) reported having had Medicaid coverage at any point of time during the survey year, and (4) had positive person weights, and were interviewed for all five rounds of the 2-year MEPS survey interviews. Events associated with any particular disease conditions were self-reported and coded by the MEPS data reviewers accordingly. A portion of the self-reported data were verified by contacting the health care providers associated with the events reported and coded accordingly.⁴² Selection of individuals in the study population on the basis of a medical event associated with the disease condition of interest is a valid approach to ensure that the study sample is not impacted by misreporting of chronic conditions diagnoses.⁴³ Respondents who reported having Medicare, or pregnancy were excluded from this study.

A major goal of this study was to characterize Medicaid discontinuity among individuals with CVD or high-risk conditions for CVD. Unlike previous studies, this study accounted for presence of other sources of health coverage among the discontinuous Medicaid coverage groups.¹⁵ This was done since individuals who do not report having continuous full-year Medicaid coverage and have no other health insurance during the year may be different from those with other sources of health insurance coverage, in terms of characteristics and health care utilization. Individuals in the study population who reported having Medicaid coverage during all 12 months of a survey year were classified as having continuous Medicaid coverage, which was the comparison group and designated as ‘Continuous Medicaid’ in this study. Individuals who reported having <12 months of Medicaid coverage during a survey year were divided into two groups: (1) individuals with <12 months of Medicaid coverage who did not report any other sources of health insurance during the survey year, designated in this study as ‘Discontinuous–Uninsured’, and (2) individuals with <12 months of Medicaid coverage who also reported having other source(s) of health insurance during the survey year (either private, Tricare, or any hospital- or physician group-based health insurance), designated in this study as ‘Discontinuous–Insured’.

Study Variables

Our primary outcomes of interest are an indicator for Medicaid discontinuity without other health insurance coverage (Discontinuous–Uninsured), and an indicator for Medicaid discontinuity with other sources of health insurance coverage (Discontinuous–Insured). The measures of comparison between the three groups, which were also adjusted for in the regression models included the following covariates:

- 1) Demographic covariates: age (in years), gender (female vs. male), race/ethnicity (Non-Hispanic White vs. Non-Hispanic Black, Hispanic, and other races), marital status (not married vs. married), family size (N), and geographic residence (Northeast vs. Midwest, South, and West).
- 2) Socioeconomic covariates: education (less than high school vs. high school/GED, and Bachelor’s or higher), income as a percentage of the federal poverty level (FPL) (<100% FPL vs. 100 - <125% FPL, and \geq 125% FPL), employment (always unemployed vs. intermittent employment, and always employed), binary indicator

- for Metropolitan Statistical Area (MSA) residence.
- 3) Medicaid-related covariates: binary indicator for Supplementary Security Income (SSI) for disability receipt or participation in the Temporary Assistance for Needy Families (TANF) program, and a binary indicator for Medicaid HMO/Managed Care participation.
 - 4) Health-related covariates: perceived health status (excellent/very good/good vs. fair/poor), perceived mental health status (excellent/very good/good vs. fair/poor), number of CVD or high-risk conditions for CVD (N), binary indicator for mental illness/substance abuse diagnosis, binary indicator for respiratory diseases diagnosis, binary indicator for arthritis/joint pain diagnosis, and Charlson Comorbidity Index (CCI) score.

All regression models included the year fixed effects to capture any macro-level variations during the study period.

Statistical Analysis

Survey-weighted proportions of prevalence of CVD or high-risk conditions for CVD were calculated. Chi-square tests and t-tests were conducted as part of the initial bivariate statistics to compare characteristics across the groups in the study population. Survey-weighted multivariate logistic regression analyses were then conducted to determine the predictors of Medicaid discontinuity in the study population. As denoted in **Equation 2-1** and **Equation 2-2**,

$$Pr[DU=1|x] = F^{-1}(\hat{\alpha} + \sum_{k=1}^K \hat{\beta}_k X_k) \quad (\text{Eq. 2-1})$$

$$Pr[DI=1|x] = F^{-1}(\hat{\alpha} + \sum_{k=1}^K \hat{\beta}_k X_k) \quad (\text{Eq. 2-2})$$

In the logistic regression analyses, both Discontinuous–Uninsured (denoted by DU) and Discontinuous–Insured (denoted by DI) were modeled as dichotomous outcome variables as a function of the demographic, socioeconomic, Medicaid eligibility and health-related covariates. F^{-1} denotes the inverse of the cumulative standard logistic distribution function that relates the outcome on the probability scale to the covariates. $Pr[]$ denotes the population average probability of Medicaid discontinuity (for both with and without other sources of insurance coverage models) conditional on the covariates. X denotes the vector of aforementioned individual characteristics that predict Medicaid discontinuity in the two population groups of interest. The proportion of missing data was approximately 6% for the study. As a result, the listwise deletion approach was used to account for missing values. Multicollinearity was assessed for all the regression models, and the variance inflation factor was found to be less than 5, which was lower than the widely accepted threshold of 10 for existence of multicollinearity. The Hosmer-Lemeshow tests for goodness-of-fit were highly insignificant for all the logistic regression models indicating no significant difference between the observed and predicted values of the response variables, and that the model fit the data. All data

analyses were conducted using SAS 9.3 and STATA 13 to account for the complex survey design of MEPS.

RESULTS

Figure 2-2 is a schematic representation of the process employed for deriving the unweighted, unadjusted study sample. Of the 344,933 individuals in the ten-year period between 2002 and 2011, 328,135 had positive MEPS person-level weights. After excluding individuals with Medicare (44,921), those who reported pregnancy (17), and those respondents who were not in-scope throughout the survey period (7,359), 275,838 non-institutionalized study respondents remained. Of these, 231,675 respondents who did not report being diagnosed with CVD or high-risk conditions for CVD were excluded, resulting in a total of 41,163 individuals being selected who either reported being diagnosed with, or having an event (prescribed medicine, inpatient, ER, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, which were the index chronic conditions considered in this study. Of these, a total of 38,186 individuals were further excluded as they either did not have Medicaid coverage at any point of time during the study period, or they were less than 18 or more than 64 years of age, resulting in a total overall sample of 5,977 individuals aged 18-64 years who had Medicaid at any given point in time during the study period.

The overall study sample (5,977) was divided further into groups: the continuous Medicaid coverage group consisting of 3,926 individuals (65.7%) who had uninterrupted Medicaid coverage throughout the year (hereby referred to as Continuous Medicaid), and the discontinuous Medicaid coverage group consisting of 1,900 individuals (31.8%) who had less than 12 months of Medicaid coverage. There were 151 individuals (2.5%) who reported both full-year Medicaid coverage along with other sources of health insurance, which were excluded in the main analyses and included in the sensitivity analyses. In order to capture the heterogeneity in the discontinuous Medicaid subpopulation due to presence of other sources of health insurance coverage, this group was further divided into two groups: the Discontinuous–Uninsured, which consisted of 1,407 individuals (74.1% of the discontinuous Medicaid subpopulation, 23.5% of the total study sample) who did not report any other form of health insurance during the year, and the Discontinuous–Insured, consisting of 493 individuals (25.9% of the discontinuous Medicaid subpopulation, 8.3% of the total study sample) who reported some other form of health insurance during the year (either, private, Tricare, or other forms of hospital- or physician group-based health insurance). Thus, all outcome models were analyzed separately; comparing Continuous Medicaid and the Discontinuous–Uninsured populations, and separately comparing the Continuous Medicaid vs. the Discontinuous–Insured population groups. This was done to capture the variability among the population group of interest, those with Discontinuous Medicaid. Individuals who do not report having continuous full-year Medicaid coverage and have no other health insurance during the year may have significant differences, not only in sociodemographic characteristics and health status, but may also have different health care utilization patterns and health behaviors.

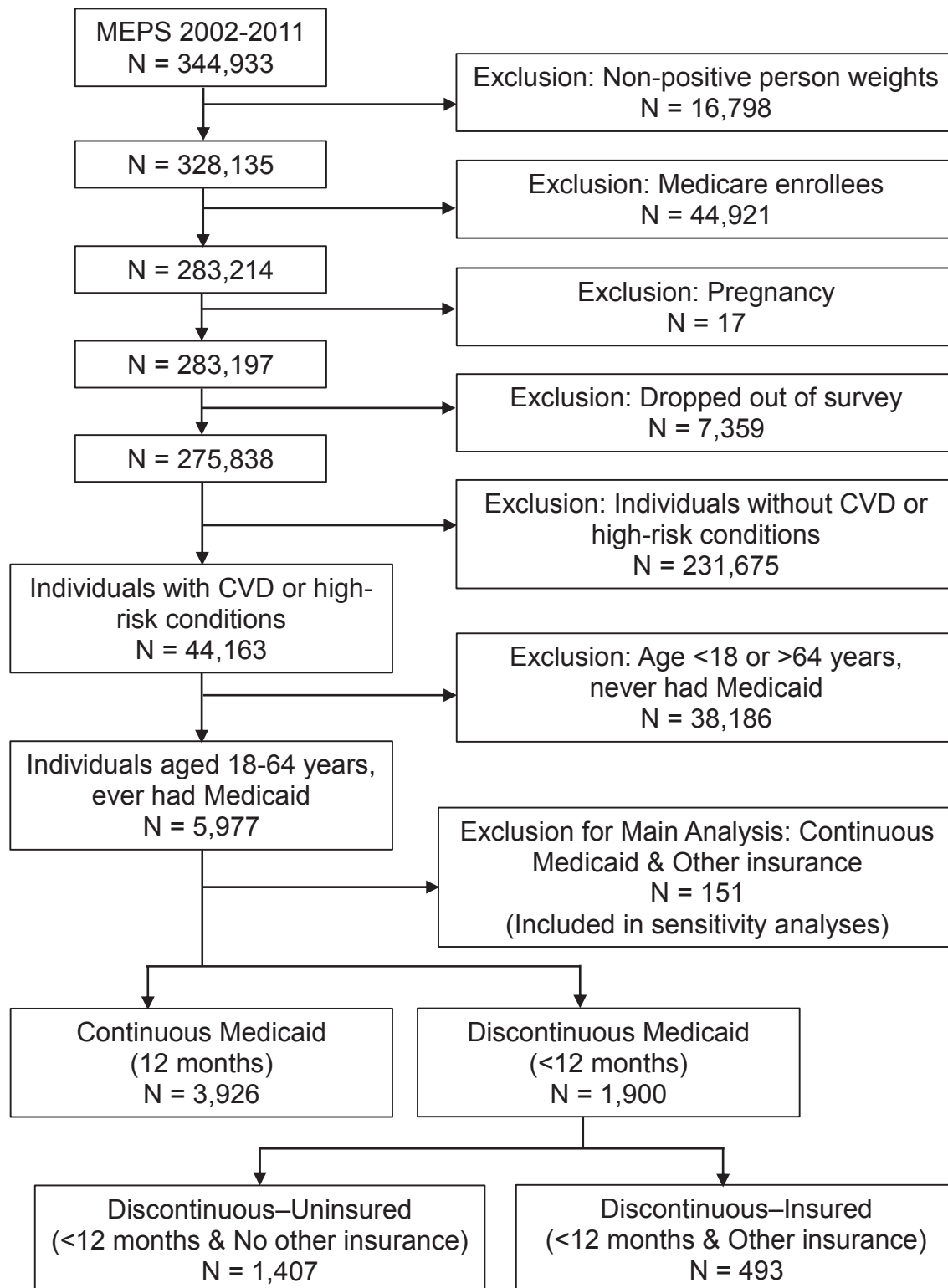


Figure 2-2. Schematic Representation of the Study Sample Derivation Process

Table 2-1 depicts the distribution of the various diagnosis and utilization components or criteria which were considered in this study for sample selection. From an overall sample of 5,997, 3,997 reported being diagnosed with one or more of the CVD or high-risk conditions for CVD examined in this study, 5,332 reported taking on or more prescription medicines associated with these index chronic conditions, and 3,713 had reported having one or more office-based provider visits for one or more of the index chronic conditions examined in this study.

Prevalence of CVD or High-Risk Conditions for CVD

Table 2-2 shows the prevalence of CVD or high-risk conditions for CVD in the study population. The prevalence of the index disease conditions was fairly even between the Continuous and Discontinuous–Uninsured groups with hypertension being the most prevalent disease (49.9% vs. 52.7%, $P=0.179$), followed by lipid disorders (34.1% vs. 31.3%, $P=0.139$), and diabetes (28.3% vs. 28.5%, $P=0.918$). The prevalence of coronary artery disease (8.1% vs. 7.3%, $P=0.407$), MI (4.3% vs. 4.0%, $P=0.739$), stroke (4.4% vs. 3.8%, $P=0.572$), CHF (3.3% vs. 3.1%, $P=0.856$), peripheral and visceral atherosclerosis (1.6% vs. 0.9%, $P=0.082$), and chronic kidney disease (0.8% vs. 0.4%, $P=0.293$) had lower prevalence in these groups. Distribution of the index conditions when comparing the Continuous with the Discontinuous–Insured groups was similar. The mean number of index chronic conditions was higher among the Continuous Medicaid group compared to the Discontinuous–Uninsured ($P=0.028$), and the Discontinuous–Insured groups ($P=0.008$). When comparing the raw numbers of cardiovascular disease and associated comorbidities in this study population, this study found that 18.7% of individuals in the Continuous group had 3 or more chronic conditions, compared to 17.7% in the Discontinuous–Uninsured ($P=0.792$), and 13.6% in the Discontinuous–Insured groups ($P=0.096$).

Among individuals with MI, the most prevalent chronic comorbid index driving diagnosis was hypertension, followed by lipid disorders, coronary artery disease, diabetes, stroke, and congestive heart failure. In this subgroup, the prevalence of stroke was significantly different between the Continuous and Discontinuous–Insured groups ($P=0.027$). When examining individuals with CHF, the most prevalent chronic comorbid index driving diagnosis was found to be hypertension, followed by lipid disorders, diabetes, coronary artery disease, myocardial infarction, and stroke. Finally, among individuals with coronary artery disease, hypertension was again the most prevalent comorbidity, followed by lipid disorders, diabetes, myocardial infarction, and congestive heart failure (results not shown).

Medicaid Enrollment among the Two Discontinuous Medicaid Groups

Table 2-3 shows the summary statistics for Medicaid enrollment among the discontinuous Medicaid insurance groups. The average duration of enrollment for the undivided discontinuous Medicaid population was 6.0 months. The Discontinuous–

Table 2-1. Diagnosis and Event Components Used to Select the Study Population

Components	Frequency (N)	Weighted Frequency (N)	% (SE)
CVD or High-Risk Conditions for CVD	3,997	2,638,024	67.8 (0.9)
Rx for CVD or High-Risk Conditions for CVD	5,332	3,474,117	89.3 (0.6)
Inpatient Visit for CVD or High-Risk Conditions for CVD	372	260,990	6.7 (0.4)
ER Visit for CVD or High-Risk Conditions for CVD	463	304,458	7.8 (0.5)
Outpatient Visit for CVD or High-Risk Conditions for CVD	480	324,800	8.3 (0.6)
Office-Based Visit for CVD or High-Risk Conditions for CVD	3,713	2,379,045	61.1 (0.9)
Study Population	5,977	3,892,363	--

CVD: Cardiovascular disease; Rx: Prescription medication; ER: Emergency room. CVD include myocardial infarction, coronary artery disease, congestive heart failure, peripheral and visceral atherosclerosis, stroke. High-risk conditions for CVD include hypertension, lipid disorders, and diabetes, chronic kidney disease. Study Population includes individuals reporting one or more of at least one of the study components listed in the table.

Table 2-2. Prevalence of CVD or High-Risk Conditions for CVD in the Study Population

Prevalence	Index Diseases	Continuous Medicaid		Discontinuous–Uninsured		Discontinuous–Insured	
		Weighted N or Mean	% (SE) or SE	Weighted N or Mean	% (SE) or SE	Weighted N or Mean	% (SE) or SE
	MI	109,746	4.3 (0.5)	36,811	4.0 (0.6)	10,903	2.6 (0.9)
	CAD	207,594	8.1 (0.6)	66,361	7.3 (0.9)	25,846	6.1 (1.2)
	CHF	83,275	3.3 (0.5)	28,341	3.1 (0.8)	5,709	1.4 (0.7)
	PVA	40,962	1.6 (0.3)	8,511	0.9 (0.3)	7,810	1.8 (0.7)
	Stroke	111,114	4.4 (0.5)	35,051	3.8 (0.8)	13,361	3.2 (1.1)
	Hypertension	1,273,888	49.9 (1.2)	481,202	52.7 (1.8)	196,635	46.4 (2.7)
	Lipid Disorders	870,818	34.1 (1.1)	285,706	31.3 (1.7)	133,370	31.4 (2.5)
	Diabetes	723,694	28.3 (1.1)	260,672	28.5 (1.8)	113,299	26.7 (2.2)
	CKD	19,144	0.8 (0.2)	4,046	0.4 (0.2)	1,323	0.3 (0.2)
Mean Index Diseases (SE)		1.73	0.01	1.70	0.01*	1.56	0.06**
No. of Index Diseases							
1		951,993	54.7 (1.4)	351,268	56.2 (2.2)	170,538	62.4 (3.2)
2		462,928	26.6 (1.1)	163,062	26.1 (1.8)	65,581	24.0 (3.2)
≥3		325,033	18.7 (1.1)	110,621	17.7 (1.8)	37,000	13.6 (2.4)

CVD: Cardiovascular disease; MI: myocardial infarction; CAD: coronary artery disease; CHF: congestive heart failure; PVA: peripheral vascular disease; CKD: chronic kidney disease; SE: standard error

*** $P < 0.001$; ** $P < 0.01$; * $P < 0.05$ for difference from the Continuous Medicaid group. Index diseases include the CVD or its associated risk factor comorbidities considered in this study. Sample consists of 3,003,997 individuals or weighted sample of 2,638,024 who reported being diagnosed with one or more index diseases.

Table 2-3. Medicaid Enrollment among the Discontinuous Medicaid Groups

Coverage Components	Discontinuous–Uninsured		Discontinuous–Insured	
	Mean or Weighted N	SE or % (SE)	Mean or Weighted N	SE or % (SE)
Mean Duration of Medicaid	6.5 months	0.1	5.0 months***	0.2
Mean Duration of No Insurance	5.5 months	0.1	1.2 months***	0.1
Insurance Status After 12 Months[†]				
Still in Medicaid	83,125	21.9 (2.2)	11,647***	8.0 (2.0)
Uninsured	297,045	78.1 (2.2)	10,058***	6.0 (2.0)
Other Insurance	--	--	147,896	86.0 (2.0)
Medicaid Coverage Transitions				
1	752,311	82.3 (1.4)	350,803	82.7 (2.0)
≥2	161,615	17.7 (1.4)	73,462	17.3 (2.0)

Sample consisted of 380,171 individuals in the Discontinuous–Uninsured group and 169,600 individuals in the Discontinuous–Insured group who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, hospital outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid at the beginning of the year.

*** $P<0.001$; ** $P<0.01$; * $P<0.05$ for difference from the Discontinuous–Uninsured group.

Uninsured group was enrolled in Medicaid for a longer period on average compared to the Discontinuous–Insured group (6.5 vs. 5.0 months, $P<0.001$). Additionally, this study examined the average duration of no insurance and found individuals in the Discontinuous–Insured group were uninsured for a little over a month (5.5 vs. 1.2 months, $P<0.001$). This study also examined Medicaid enrollment status after 12 months of initial Medicaid enrollment. Overall, a total of 4,898 individuals were enrolled in the month of January of the survey year, of which 85.3% were still enrolled in Medicaid during the month of December in the survey year, 3.5% had some other form of insurance coverage, and 10.2% were uninsured (results not shown). Of the 621 individuals in the Discontinuous–Uninsured group who had Medicaid in the beginning of the year, 78.1% were uninsured while 21.9% still had Medicaid coverage at the end of the year. Of the 200 individuals in the Discontinuous–Insured group who had Medicaid in the beginning of the survey year, only 8.0% still had Medicaid, 6.0% were uninsured, and the remaining 86% had some other source of health insurance coverage at the end of the year. Finally, when examining the number of transitions into and/or out of Medicaid coverage, this study found that overall 17.6% of individuals in this study population had more than one transitions, and the proportion of multiple transitions was similar between the two discontinuous Medicaid subgroups (17.7% vs. 17.3%, $P=0.875$).

Characteristics of the Study Population

Table 2-4 displays the demographic, socioeconomic, eligibility, and health-related characteristics of the study population. The mean age of the full sample was 46.5 years (not displayed). In comparison to the Continuous Medicaid group, those in the Discontinuous–Uninsured group were less likely to belong to other races [5.5% (Discontinuous–Insured) vs. 9.5% (Continuous), $P<0.001$], were more likely to be married (37.7% vs. 29.9%, $P<0.001$), and were more likely to belong to the South (34.5% vs. 29.8%, $P=0.035$). When examining the socioeconomic characteristics, this study found that the Discontinuous–Uninsured were more likely to have completed high school/GED (54.9% vs. 50.4%, $P=0.004$) or college (12.2% vs. 9.4%, $P=0.004$), were more likely to have incomes $\geq 125\%$ FPL (43.0% vs. 33.9%, $P<0.001$), and were more likely to have intermittent (25.0% vs. 12.4%, $P<0.001$), as well as continuous full-year employment (20.7% vs. 15.1%, $P<0.001$). Among the Medicaid eligibility variables, the Discontinuous–Uninsured were less likely to be eligible for SSI due to disability or participate in TANF (16.9% vs. 41.5%, $P<0.001$), and were less likely to report Medicaid HMO/managed care (50.6% vs. 61.3%, $P<0.001$). Finally, when examining the health-related covariates, they were less likely to report having mental /substance abuse illnesses (24.9% vs. 30.8%, $P=0.003$), respiratory illnesses (12.9% vs. 19.7%, $P<0.001$), or arthritis/joint pain diagnosis (29.3% vs. 34.1%, $P=0.015$), and were less likely to report having fair or poor perceived mental health status (25.7% vs. 33.7%, $P<0.001$) when compared to the Continuous Medicaid group. The mean CCI score was not significantly different between the two groups ($P=0.078$).

When comparing the Discontinuous–Insured group to the Continuous Medicaid group (**Table 2-4**), this study found that when examining the demographic covariates,

Table 2-4. Characteristics of the Study Population

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
DEMOGRAPHIC								
Age					0.621			0.027
18 – 34 years	388,425	15.9 (1.0)	165,340	18.1 (1.3)		98,188	23.1 (2.3)	
35 – 49 years	875,809	35.9 (1.2)	324,653	35.5 (1.8)		141,109	33.3 (2.7)	
50 – 64 years	1,173,144	48.2 (1.3)	423,933	46.4 (2.0)		184,968	43.6 (3.0)	
Gender					0.359			0.223
Female	1,632,006	67.0 (1.2)	624,410	68.3 (1.8)		267,502	63.1 (2.8)	
Race/Ethnicity					<0.001			<0.001
Non-Hispanic White	1,094,002	44.9 (1.8)	442,534	48.4 (2.1)		256,231	60.4 (3.2)	
Non-Hispanic Black	666,358	27.3 (1.5)	219,033	24.0 (1.6)		799,36	18.8 (2.0)	
Hispanic	445,601	18.3 (1.2)	202,106	22.1 (1.5)		50,704	12.0 (1.6)	
Others	231,416	9.5 (0.9)	50,253	5.5 (0.9)		37,393	8.8 (1.7)	
Marital Status					<0.001			<0.001
Married	727,457	29.9 (1.3)	344,150	37.7 (1.8)		184,414	43.5 (3.2)	
Region					0.035			0.034
Northeast	608,070	25.0 (1.6)	187,785	20.5 (1.7)		104,394	24.6 (2.9)	
Midwest	477,401	19.6 (1.4)	184,710	20.2 (1.7)		115,146	27.2 (3.2)	
South	727,520	29.8 (1.6)	315,073	34.5 (1.8)		98,072	23.1 (2.3)	
West	624,386	25.6 (1.7)	226,358	24.8 (1.9)		106,652	25.1 (3.0)	

Table 2-4. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
Family Size Mean (SE)	2.8 (0.1)		2.9 (0.1)		0.481	2.8 (0.1)		0.900
SOCIOECONOMIC								
Education					0.004			<0.001
<High School	971,041	40.2 (1.4)	298,300	32.9 (1.8)		67,896	16.0 (1.9)	
High School/GED	1,215,968	50.4 (1.4)	498,289	54.9 (2.0)		232,334	54.9 (3.2)	
College Degree	227,668	9.4 (0.7)	111,300	12.2 (1.4)		123,109	29.1 (3.2)	
Income					<0.001			<0.001
<100% FPL	1,384,551	56.8 (1.2)	428,540	46.9 (1.7)		104,719	24.7 (2.3)	
100 – <125% FPL	227,591	9.3 (0.6)	91,940	10.1 (1.1)		28,783	6.8 (1.3)	
≥125% FPL	825,236	33.9 (1.1)	393,445	43.0 (1.8)		290,763	68.5 (2.6)	
Employment					<0.001			<0.001
Always Unemployed	1,767,234	72.5 (1.2)	496,282	54.3 (1.8)		112,490	26.5 (2.2)	
Unstable Employment	301,597	12.4 (0.8)	228,246	25.0 (1.6)		129,583	30.6 (2.5)	
Always Employed	368,547	15.1 (0.9)	189,398	20.7 (1.4)		182,191	42.9 (2.9)	
MSA	1,931,797	79.3 (1.8)	721,532	78.9 (1.9)	0.854	344,287	81.1 (2.7)	0.503
MEDICAID-RELATED								
SSI/TANF	1,012,240	41.5 (1.3)	154,157	16.9 (1.3)	<0.001	41,212	9.7 (1.6)	<0.001
Medicaid HMO/MCO	1,424,726	61.3 (1.5)	459,330	50.5 (2.0)	<0.001	233,429	55.0 (3.0)	0.064

Table 2-4. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
HEALTH-RELATED								
Perceived Health Status					0.144			<0.001
Fair to Poor	1,314,794	53.9 (1.2)	467,555	51.2 (1.9)		170,598	40.2 (2.5)	
Perceived Mental Health					<0.001			<0.001
Fair to Poor	821,318	33.7 (1.2)	234,931	25.7 (1.7)		84,716	20.0 (2.2)	
CVD or High-Risk Conditions for CVD					0.775			0.037
≤1	1,678,222	68.9 (1.1)	640,242	70.1 (1.8)		321,685	75.8 (2.3)	
2	444,995	18.2 (0.8)	163,062	17.8 (1.4)		65,581	15.5 (2.2)	
≥3	314,160	12.9 (0.8)	110,621	12.1 (1.3)		37,000	8.7 (1.5)	
Mental/Substance Abuse Illnesses	751,357	30.8 (1.2)	227,885	24.9 (1.6)	0.003	115,891	27.3 (2.4)	0.176
Respiratory Illnesses	479,702	19.7 (1.1)	117,992	12.9 (1.1)	<0.001	39,380	9.3 (1.5)	<0.001
Arthritis/Joint Pain	830,408	34.1 (1.3)	267,358	29.3 (1.7)	0.015	107,575	25.4 (2.4)	0.002
CCI Score Mean (SE)	1.2 (0.0)		1.1 (0.1)		0.078	0.9 (0.1)		0.003
Population Size	2,437,378	64.6 (1.0)	913,926	24.2 (0.8)		424,265	11.2 (0.7)	

Table 2-4. (Continued)

Study population includes MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with, or reported having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with, 1 or more cardiovascular disease or associated risk factor comorbidities considered in this study.

[†]*P*-value for the difference between Discontinuous–Uninsured and Continuous Medicaid groups. ^{††}*P*-value for the difference between Discontinuous–Insured and Continuous Medicaid groups. *P*-values in **bold** indicate statistically significant difference below the threshold value of *P*<0.05.

SE: Standard error; GED: General educational development; FPL: Federal poverty level; MSA: Metropolitan Statistical Area; SSI: Supplementary Security Income due to disability; TANF: Temporary Assistance to Needy Families program participation; HMO: Health maintenance organization; MCO: Managed care organization; CCI: Charlson Comorbidity Index score (D’Hoore adapted CCI score was calculated in this study using the 3-digit ICD-9-CM codes in MEPS).

individuals in the Discontinuous–Insured group were more likely to be younger (23.1% vs. 15.9% in the 18-34 years age group, $P=0.027$), less likely to be Non-Hispanic Black (18.8% vs. 27.3%, $P<0.001$) or Hispanic (12.0% vs. 18.3%, $P<0.001$), more likely to be married (43.5% vs. 30.4%, $P<0.001$), and more likely to reside in the Midwest (27.2% vs. 19.6%, $P=0.034$). Among the socioeconomic characteristics, the Discontinuous–Insured were more likely to have a college degree (29.1% vs. 9.4%, $P<0.001$), were more likely to have family incomes $\geq 125\%$ FPL (68.5% vs. 33.9%, $P<0.001$), were more likely to have intermittent (30.5% vs. 12.4%, $P<0.001$), or continuous full-year employment (42.9% vs. 15.1%, $P<0.001$). Among the Medicaid eligibility variables it was found that the Discontinuous–Insured were less likely to be eligible for SSI due to disability or participate in TANF (9.7% vs. 41.5%, $P<0.001$), however, Medicaid managed care enrollment was not statistically significant. When examining the health-related characteristics, the Discontinuous–Insured were less likely to report having fair to poor health status (40.2% vs. 53.9%, $P<0.001$), or fair to poor mental health status (20.0% vs. 33.7%, $P<0.001$), more likely to have fewer CVD or high-risk conditions for CVD ($P=0.037$) less likely to have respiratory illnesses (9.3% vs. 19.7%, $P<0.001$), or arthritis/joint pain diagnosis (25.4% vs. 34.1%, $P=0.002$), and had a lower mean CCI score [Mean (SE): 0.9 (0.07) vs. 1.2 (0.04), $P=0.003$] in comparison to the Continuous Medicaid group.

Characteristics Associated with Medicaid Discontinuity

The final study sample with non-missing values on all variables, after undergoing the listwise deletion process, was 5,738, comprising of 3,859 individuals in the Continuous Medicaid group, 1,388 individuals in the Discontinuous–Uninsured group, and 491 individuals in the Discontinuous–Insured group. **Table 2-5** shows the results of the multivariate logistic regression analyses to determine the predictors of Medicaid discontinuity among non-elderly adults with CVD or high-risk conditions for CVD.

Discontinuous–Uninsured

When comparing individuals with discontinuous Medicaid coverage with no other sources of health coverage (Discontinuous–Uninsured) to those with continuous Medicaid coverage (Continuous) group, this study found that among the demographic predictors, male respondents had 18% lower odds than females [Odds Ratio (OR): 0.82; 95% Confidence Interval (CI): 0.68 – 0.98], Non-Hispanic Blacks had a 21% lower odds (OR: 0.79, 95% CI: 0.64 – 0.98), and individuals belonging other races had a 43% lower odds compared to Non-Hispanic Whites (OR: 0.57, 95% CI: 0.40 – 0.81), married individuals had 26% higher odds than respondents who were not married (OR: 1.26, 95% CI: 1.03 – 1.55), and individuals living in the South had about 65% higher odds compared to the Northeast residents (OR: 1.65, 95% CI: 1.29 – 2.10), of experiencing Medicaid discontinuity with no subsequent (or concurrent) sources of health coverage.

Among the socioeconomic predictors of Medicaid discontinuity, this study found that individuals with family income levels of $\geq 125\%$ FPL had about 23% higher odds

Table 2-5. Characteristics Associated with Medicaid Discontinuity in the Study Population

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
DEMOGRAPHIC		
Age		
18 – 34 years	1.00 (Reference)	1.00 (Reference)
35 – 49 years	1.10 (0.83 – 1.45)	1.12 (0.75 – 1.67)
50 – 64 years	1.25 (0.92 – 1.69)	0.85 (0.62 – 1.16)
Gender		
Female	1.00 (Reference)	1.00 (Reference)
Male	0.82 (0.68 – 0.98)*	0.94 (0.73 – 1.21)
Race/Ethnicity		
Non-Hispanic White	1.00 (Reference)	1.00 (Reference)
Non-Hispanic Black	0.79 (0.64 – 0.98)*	0.59 (0.43 – 0.82)**
Hispanic	1.09 (0.85 – 1.38)	0.65 (0.44 – 0.96)*
Others	0.57 (0.40 – 0.81)**	1.01 (0.60 – 1.71)
Marital Status		
Not Married	1.00 (Reference)	1.00 (Reference)
Married	1.26 (1.03 – 1.55)*	1.50 (1.09 – 2.06)*
Region		
Northeast	0.86 (0.66 – 1.11)	1.00 (Reference)
Midwest	1.15 (0.87 – 1.52)	1.46 (0.97 – 2.18)
South	1.40(1.09 – 1.80)**	1.13 (0.78 – 1.63)

Table 2-5. (Continued)

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
West	1.00 (Reference)	0.99 (0.67 – 1.48)
Family Size	0.94 (0.89 – 1.00)*	0.86 (0.76 – 0.96)*
SOCIOECONOMIC		
Education		
<High School	1.00 (Reference)	1.00 (Reference)
High School/GED	1.22 (1.00 – 1.49)	1.77 (1.26 – 2.49)***
College Degree	1.05 (0.76 – 1.45)	3.13 (2.16 – 4.53)***
Income		
<100% FPL	1.00 (Reference)	1.00 (Reference)
100 – <125% FPL	1.06 (0.80 – 1.41)	1.21 (0.77 – 1.92)
≥125% FPL	1.23 (1.01 – 1.50)*	2.52 (1.89 – 3.36)***
Employment		
Always Unemployed	1.00 (Reference)	1.00 (Reference)
Unstable Employment	1.92 (1.49 – 2.47)***	3.97 (2.86 – 5.51)***
Always Employed	1.23 (0.96 – 1.56)	3.52 (2.47 – 5.01)***
MSA		
Non-MSA	1.00 (Reference)	1.00 (Reference)
MSA	1.08 (0.85 – 1.37)	1.28 (0.85 – 1.92)

Table 2-5. (Continued)

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
MEDICAID-RELATED		
SSI/TANF		
No	1.00 (Reference)	1.00 (Reference)
Yes	0.33 (0.27 – 0.41)***	0.33 (0.21 – 0.51)***
Medicaid HMO/MCO		
No	1.00 (Reference)	1.00 (Reference)
Yes	0.61 (0.52 – 0.73)***	0.70 (0.53 – 0.92)*
HEALTH-RELATED		
Perceived Health Status		
Excellent/Very Good/Good	1.00 (Reference)	1.00 (Reference)
Fair to Poor	1.25 (1.04 – 1.51)*	1.09 (0.83 – 1.42)
Perceived Mental Health		
Excellent/Very Good/Good	1.00 (Reference)	1.00 (Reference)
Fair to Poor	0.83 (0.67 – 1.02)	0.84 (0.60 - 1.17)
CVD or High-Risk Conditions for CVD		
≤1	1.00 (Reference)	1.00 (Reference)
2	1.05 (0.84 – 1.32)	0.80 (0.54 – 1.17)
≥3	0.97 (0.69 – 1.36)	0.84 (0.50 – 1.41)

Table 2-5. (Continued)

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
Mental/Substance Abuse Illnesses		
No	1.00 (Reference)	1.00 (Reference)
Yes	0.91 (0.73 – 1.15)	1.23 (0.90 – 2.18)
Respiratory Illnesses		
No	1.00 (Reference)	1.00 (Reference)
Yes	0.72 (0.56 – 0.91)**	0.66 (0.44 – 1.00)*
Arthritis/Joint Pain		
No	1.00 (Reference)	1.00 (Reference)
Yes	0.97 (0.79 – 1.19)	0.92 (0.67 – 1.27)
CCI Score	0.99 (0.93 – 1.06)	1.07 (0.98 – 1.18)

Study population includes MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with, or reported having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with, 1 or more cardiovascular disease or associated risk factor comorbidities considered in this study.

*** P -value<0.001, ** P <0.01, * P <0.05 for the difference between Continuous Medicaid vs Discontinuous–Uninsured and Continuous Medicaid vs Discontinuous–Insured groups

SE: Standard error; GED: General educational development; FPL: Federal poverty level; MSA: Metropolitan Statistical Area; SSI: Supplementary Security Income due to disability; TANF: Temporary Assistance to Needy Families program participation; HMO: Health maintenance organization; MCO: Managed care organization; CCI: Charlson Comorbidity Index score (D’Hoore adapted CCI score was calculated in this study using the 3-digit ICD-9-CM codes in MEPS).

(OR: 1.23, 95% CI: 1.01 – 1.50), and individuals with intermittent employment had almost two times higher odds compared to the unemployed (OR: 1.92, 95% CI: 1.49 – 2.47) of belonging to the Discontinuous–Uninsured group. When examining the Medicaid-related variables, this study found that SSI due to disability or TANF participation was associated with about 67% lower odds (OR: 0.33, 95% CI: 0.27 – 0.41), and Medicaid HMO/managed care participation was associated with 39% lower odds (OR: 0.61, 95% CI: 0.52 – 0.73) of experiencing Medicaid discontinuity. When examining health-related variables, this study found that individuals with respiratory illnesses had about 28% lower odds (OR: 0.72, 95% CI: 0.56 – 0.91), and those reporting fair to poor health status had about 25% higher odds (OR: 1.25, 95% CI: 1.04 – 1.51) of belonging to the Discontinuous–Uninsured group.

Discontinuous–Insured

When examining the predictors of discontinuous Medicaid coverage who also reported having other sources of health insurance coverage (Discontinuous–Insured), this study found that among the demographic characteristics, Non-Hispanic Blacks had 41% lower odds (OR: 0.59, 95% CI: 0.43 – 0.82), Hispanics had 35% higher odds (OR: 0.65, 95% CI: 0.44 – 0.96), and married individuals had about 48% higher odds (OR: 1.48, 95% CI: 1.07 – 2.05) of belonging of the Discontinuous–Insured group. In addition, as the number of family members increased by one, the odds of experiencing Medicaid discontinuity decreased by about 15% (OR: 0.85, 95% CI: 0.75 – 0.96).

Among the socioeconomic variables, this study found that individuals with higher education had higher odds compared to individuals with no high school education [OR (95% CI) for high school/GED: 1.75 (1.25 – 2.46), and OR (95% CI) for College degree: 3.05 (2.10 – 4.44)], individuals with higher family income had higher odds compared to low-income group [OR (95% CI) for $\geq 125\%$ FPL: 2.55 (1.91 – 3.41)], and those with partial or full-year employment had higher odds [OR (95% CI) for intermittent employment: 3.98 (2.87 – 5.54), and OR (95% CI) for always employed: 3.46 (2.42 – 4.94)] compared to unemployed individuals of belonging to the Discontinuous–Insured group. Both the Medicaid-related variables were significant predictors with individuals who reported receiving SSI due to disability or TANF participation having 67% lower odds (OR: 0.33, 95% CI: 0.21 – 0.51), and those reporting Medicaid HMO/managed care participation having about 30% lower odds (OR: 0.70, 95% CI: 0.53 – 0.92) of experiencing Medicaid discontinuity. Finally, among the health-related variables, individuals who reported having respiratory illnesses had about 35% lower odds (OR: 0.65, 95% CI: 0.43 – 0.98) of being in the Discontinuous–Insured group.

DISCUSSION

The present study findings highlight the high prevalence of Medicaid discontinuity among non-elderly adults with CVD or high-risk conditions for CVD. National estimates indicate that of the roughly 4 million individuals who reported having CVD or an associated comorbidity in the first round of the MEPS survey and Medicaid at

any point in time during the year, a little over 1.3 million had less than full-year Medicaid coverage, over 900,000 (or 68.3%) of which did not have any other health insurance during the year. To our knowledge, no previous study has stratified the discontinuous Medicaid group based on presence (Discontinuous–Insured) or absence (Discontinuous–Uninsured) other sources of health insurance coverage. On average, the Discontinuous–Uninsured group in this study was enrolled for about 6.5 months in Medicaid during a year. The instability in Medicaid coverage is further highlighted by the finding that of the 3.1 million individuals (Unweighted frequency = 4,898) who had Medicaid at the beginning of the year, a little over 300,000 (or 10%) of these individuals were uninsured at the end of the year. When this end-of-the-year insurance status was examined among individuals with family income level of <100% of FPL, this study found that of the 1.6 million individuals who had Medicaid as the year began, over 130,000 (or 8%) of these individuals were uninsured as the year ended. Churning in Medicaid, or multiple transitions into and out of Medicaid coverage, especially to a state of having no health insurance, are a major concern among policymakers and health policy researchers.^{8,44} Churning is especially concerning when it results in loss of health insurance. This study found that about 17% of the non-elderly adult population with CVD or high-risk conditions for CVD experienced more than one transition into or out of Medicaid coverage.

Hypertension, lipid disorders and diabetes, conditions considered significant risk factors for CVD, were the most prevalent chronic conditions in this population. Among the core CVD conditions, coronary artery disease, MI, and stroke – conditions that exert significant morbidity and mortality – were the most prevalent chronic conditions in this study.^{18,21} Hypertension has been well-established as a significant high-risk condition for CVD.^{17,21} This study found hypertension to be the most prevalent chronic condition among low-income individuals with CVD or high-risk conditions for CVD. Although estimates in the literature indicate the prevalence of CVD in the Medicaid population to be over 16 million as mentioned earlier, the lower numbers in this study may indicate differences in data source and inclusion criteria of first round-disease reporting to capture the chronicity of CVD and associated risk factors. The distribution of these index diseases was similar between the continuous Medicaid coverage, and both the discontinuous Medicaid groups indicating that Medicaid discontinuity, both with or without other sources of health insurance, may be predicted by factors other than differences in CVD burden.

Of the two major public health insurance programs in the U.S., Medicaid enrollees are more likely to experience discontinuity in their health care coverage in comparison to Medicare enrollees, and the prevalence and duration of coverage discontinuity are as high as private employer-sponsored insurance among the low-income groups.⁴⁵ Being the primary, means-tested health insurance program for the low-income population, small increases in family incomes and assets often render individuals and families ineligible for Medicaid based on their state's Medicaid eligibility criteria. However, research has shown that several individuals and families lose Medicaid coverage despite being eligible, mainly due to the inability to periodically report changes in their income, assets, and residential status, among other information, and often not

being able to understand the state Medicaid reenrollment procedures.^{37,46,47} As shown in this study a significant proportion of the discontinuous Medicaid coverage group did not report having any other source of health insurance coverage. Overwhelmingly, empirical literature suggests uninsured individuals have poor access to care, utilization of health care services, unplanned inpatient admissions, and often have poor health outcomes compared to the insured populations.^{10-14,31,32}

Among the characteristics significantly associated with discontinuity, Non-Hispanic Whites had a higher likelihood of experiencing discontinuity compared to minorities, which is similar to previous findings. This may be due to the fact that minorities are more likely to be eligible for Medicaid due to differential income and assets. This is further bolstered by additional findings in this study that those with higher education, income levels, and some form of employment during the year; indicators of better socioeconomic status; have a higher likelihood of experiencing Medicaid discontinuity, which may indicate that race/ethnicity may be impacting discontinuity in conjunction with socioeconomic characteristics. Individuals residing in the southern geographic region were more likely to report having discontinuous Medicaid coverage and no other form of health insurance, which may reflect regional differences in Medicaid eligibility criteria and reenrollment procedures.⁴⁸ Respondents who reported having Medicaid through managed care organizations had significantly lower likelihood of experiencing discontinuity in their Medicaid coverage which may be due to better outreach efforts to ensure continued enrollment among their plan members compared to the traditional state fee-for-service Medicaid programs.³⁸

As of April 2015, thirty states, including the district of Columbia (D.C.) have moved forward to expand their state Medicaid programs in accordance with the regulations under the 2014 Patient Protection and Affordable Care Act (ACA), that stipulate States to provide Medicaid coverage to residents with incomes at or below 138 percent of FPL, irrespective of marital or parenthood status.⁴⁹ In addition, the health law stipulated the creation of health insurance marketplaces in each state, known as the health exchanges, as a one-stop shop for state residents and employers to purchase health insurance plans. Individuals with incomes between 138 – 400 percent of FPL, who do not meet the Medicaid eligibility criteria under the expansion provision (in states not expanding Medicaid, the range is 100 – 400 percent of FPL) would receive subsidies to purchase private plans from the health exchanges.⁵⁰ These measures intended to expand coverage to hitherto uninsured and underinsured individuals and families may decrease the uninsured rate, but research suggests that the problems of Medicaid coverage transitions may continue to persist. Research suggests that among states not expanding Medicaid, the problems of losing Medicaid coverage to become uninsured will continue to exist, since upward income mobility would need to be significant for individuals to qualify for private insurance subsidies that will kick in above the 138 percent of FPL threshold. However, among states that do expand Medicaid, the problems of transitions from Medicaid to subsidized private insurance exchanges and/or back will also exist.^{51,52}

When individuals lose Medicaid coverage to become uninsured, and especially when they are eligible for Medicaid, it is considered a policy failure, and the resulting

lack of insurance may have significant adverse health outcomes.³⁴ However, when individuals transition from one source of coverage to another, continuity of care may still be impacted, especially since different insurance plans have different network of providers, and differing benefit structures, and may cover different classes of medications and services. This study showed that over a quarter of the CVD population with discontinuous Medicaid coverage belonged to the Discontinuous–Insured group who reported having less than six months of Medicaid coverage, and were uninsured for a little over a month, on average. This may exert additional burden on individuals and families of taking out time and effort to sign up for other sources of health care coverage that includes providing the necessary documents for enrollment and may also temporarily disrupt continuity of care in this population.

Policymakers debating measures to ensure continuity in Medicaid among non-elderly adults should especially include populations with chronic and debilitating conditions such as CVD and associated risk factors in all such discussions, since the detrimental effects of coverage discontinuity and resultant disruptions in continuity of care may be amplified in such population groups. Moreover, outreach and reenrollment strategies employed by Medicaid managed care organizations can be adopted by traditional state Medicaid plans. In addition, expanding the penetration of managed care in Medicaid can also be an alternative to enable states to strike a balance between tight fiscal spending and increased and continued Medicaid enrollment. Health care providers and their staff in the meanwhile can play a major role by informing, encouraging, and if possible assisting, low-income patients with CVD, their risk factors, or disability, to sign-up and maintain their enrollment in Medicaid through periodic reporting as per the state's requirements. In a survey conducted for the Centers for Medicare and Medicaid Services (CMS), more than half of the respondents indicated that they would trust a doctor when making a decision to enroll in Medicaid or CHIP, and about 50 percent preferred getting Medicaid/CHIP-related information at the doctor's office.⁵³

As with all observational research, this study is not without its limitations. State-level Medicaid eligibility data was not available to explore the complex and dynamic patterns of discontinuity that may differ based on the state's eligibility and reporting criteria. Nonetheless, the finding of individuals residing in the south having greater likelihood of discontinuity may have captured some of the regional variations in Medicaid policies, and future research with state-level data may better account for these differences. This study only explored associations between individual characteristics and Medicaid discontinuity, and not causality, and the findings only explore predictors of within-year discontinuous Medicaid coverage and may not capture risk factors for secular trends in Medicaid discontinuity. The regression analyses may not have adjusted for unobserved covariates that may have confounded the impact of certain risk factors on the likelihood of belonging to the Discontinuous–Uninsured or the Discontinuous–Insured groups.

Limitations of small sample size due to the focus of this study on individuals with CVD or high-risk conditions for CVD, and not the entire Medicaid population, prevented us from imputing Medicaid eligibility and obtaining robust estimates of Medicaid

discontinuity predictors in this imputed subpopulation. Nonetheless, these findings have highlighted significant prevalence of overall discontinuous Medicaid coverage, and future studies with higher sample sizes of the CVD population can further explore predictors of Medicaid discontinuity separately among those losing eligibility due to upward income mobility, and those dropping out of Medicaid despite being eligible. Data in MEPS are self-reported, so the estimates may not exactly reflect the national estimates of Medicaid discontinuity. Researchers have speculated that most surveys underestimate the number of Medicaid enrollees due to a variety of reasons including stigma of public insurance. However, such misreporting is less of a problem for MEPS compared to other federal health surveys.⁵⁴ Finally, findings from this study are representative of the non-institutionalized civilian population and may not be applicable to individuals with prolonged hospitalizations or admitted to long-term care facilities.

CONCLUSION

The present study highlights the prevalence of Medicaid discontinuity among non-elderly adults with CVD or high-risk conditions for CVD and finds that a high proportion of individuals with less than full-year Medicaid coverage do not have any other source of health insurance. Women, individuals with higher education and income, those residing in the South census region, and those perceiving their health to be fair or poor may be at higher risk for Medicaid discontinuity. Medicaid managed care enrollees may have lower risk for Medicaid discontinuity. Outreach and reenrollment strategies employed by such managed care organizations can be adopted by state Medicaid agencies to improve Medicaid uptake, continuity and reenrollment, especially for enrollees with chronic conditions such as CVD or high-risk conditions for CVD. Discussions regarding policies to improve retention of and continuity in Medicaid must include individuals suffering from these chronic and debilitating disease conditions whose continuity of care may be adversely impacted due to such coverage instability. Low-income adults managing chronic conditions, especially in regions with stringent Medicaid eligibility regulations, must be encouraged to ensure continuity of Medicaid coverage.

CHAPTER 3. ASSOCIATIONS BETWEEN MEDICAID DISCONTINUITY, MEDICATION ADHERENCE AND MEDICATION UTILIZATION AMONG ADULTS WITH CARDIOVASCULAR DISEASE OR ASSOCIATED HIGH-RISK CONDITIONS

BACKGROUND

Cardiovascular disease (CVD) along with conditions that are risk factors for CVD, such as diabetes and chronic kidney disease, are some of the most common causes of death, both regionally and globally.^{16,17,55} Randomized studies and guidelines have established the importance of lowering blood pressure, cholesterol, or blood sugar levels, along with lifestyle changes such as smoking cessation, in order to reduce morbidity and mortality associated with CVD or high-risk conditions for CVD.^{56,57} Nevertheless, management of the chronic risk factor conditions has been suboptimal.⁵⁸⁻⁶⁰ Pharmacotherapy is essential to appropriately treat these chronic conditions and achieve and maintain the guidelines-recommended target levels for blood pressure, blood sugar, and cholesterol. Inability to obtain or consume medications necessary to avoid risk for chronic, debilitating conditions such as CVD may have detrimental outcomes among individuals who are at risk for these conditions.

Medicaid plays a pivotal role in providing access to care for individuals with low incomes, and disabilities. Among non-elderly adult Medicaid enrollees, the prevalence of chronic conditions is high and many have complex health care needs due to existing comorbid conditions, such as mental, and respiratory illnesses.³⁰ Prevalence of CVD or high-risk conditions for CVD among Medicaid enrollees is high; an estimated 28% of enrollees are diagnosed with CVD, 24% with high blood pressure, 17% with high cholesterol, and 9-15% with diabetes.^{21,22} Due to this high chronic disease burden, adult Medicaid enrollees have high health care needs, and need the resources to effectively manage their chronic conditions.³⁰ However, means-tested health insurance programs such as Medicaid have their own set of eligibility criteria and enrollment procedures, which vary from state to state,⁶ and it is critical for individuals and families to be cognizant of these policies and procedures to enroll in, and maintain eligibility, programs such as Medicaid. Small fluctuations in income levels for covered non-elderly adults due to increases in pay or work hours, would end up disqualifying individuals with incomes at or near eligibility thresholds. The rules and regulations for eligibility and enrollment between public and private health insurance are different, and therefore, individuals who have to transition between one system to another also experience gaps in insurance coverage, either due to failure to follow enrollment procedures, or being unaware of such an option, or even due to the stigma of enrolling for Medicaid or other low-income public programs. It is broadly due to these factors in this fragmented health insurance system, that individuals and families experience insurance instability and they continue to swing into and out of various types of health insurance coverage.

Instability in Medicaid coverage may not only impact utilization of prescription medications due to access issues, but may also affect adherence of these medications as

prescribed by health care providers. Adherence to evidence-based cardiovascular medications is essential for attaining therapeutic goals, in order to achieve favorable health outcomes.⁶¹⁻⁶⁴ Suboptimal adherence may have detrimental consequences even among non-elderly adults with multiple chronic conditions and comorbidities, who often require multiple medications to manage their disease conditions.^{65,66} Policies to ensure higher retention of Medicaid-eligible needy individuals and continuity in Medicaid coverage, along with seamless transition from Medicaid to other insurance coverage sources must be a priority for policymakers and other stakeholders, including health care providers. Pharmacotherapy may become one of the first casualties of loss of Medicaid coverage, since out-of-pocket costs have been found to be one of the most significant barriers to medication adherence and utilization.

Empirical studies examining the impact of lack of insurance coverage continuity on medication adherence among individuals with CVD or high-risk conditions for CVD is lacking. Among the few studies examining impact of discontinuous Medicaid coverage on medication utilization in general, or focusing on other disease conditions, there is a general agreement that having instability in coverage has an adverse impact on obtaining needed medications to manage chronic diseases.^{15,32,67} In intending to strengthen the literature on Medicaid coverage discontinuity by highlighting the association between discontinuous Medicaid coverage and pharmacotherapy outcomes among the CVD population, this study has the following two objectives: (1) examine whether having discontinuous Medicaid coverage is associated with poor adherence to commonly prescribed therapeutic drug classes, and (2) examine whether discontinuous Medicaid coverage is associated with low prescription drug utilization.

METHODS

Data Source

The present study is a retrospective cross-sectional comparative analysis of the Medical Expenditure Panel Survey (MEPS). MEPS is a set of large-scale surveys of families, individuals, their medical providers and employers across the United States. It is jointly sponsored by the Agency for Healthcare Research and Quality (AHRQ) and the National Center for Health Statistics (NCHS), and has been conducted annually since 1996. It has two major components; the Household Component (HC), the Insurance Component, with data on medical providers in the HC being supplemented by the Medical Provider Component. MEPS collects detailed information from a nationally representative sample of the civilian noninstitutionalized population of the U.S. on health services utilization and health expenditures, insurance coverage, and sources of payment.

MEPS has an overlapping panel design. Two separate panels of respondents are interviewed simultaneously during a calendar year. A new panel of sample households is selected each year, and data for each panel are collected for two calendar years. The two years of data for each panel are collected in five rounds of interviews that take place over

a two and a half year period. This provides continuous and current estimates of health care expenditures at both the person and household level for two panels for each calendar year. By combining data from the overlapping panels, these annual files provide nearly double the sample size of individual panels and cover the entire calendar year for each respondent. Although all MEPS data are reported by respondents during computer-assisted personal interviewing, further detailed health service use data, including on prescription drugs, are collected from a sample of providers with respondents' permission.³⁹

When obtaining national estimates from surveys such as MEPS, appropriate sample size is critical to obtain reliable estimates, since estimates of some population subgroups may vary from year to year. The advantage of survey such as MEPS is that they allow for several years' of data to be pooled together in order to increase the sample size and improve the precision of estimates. The pooling of the different years of data was conducted in accordance with the guidelines provided by MEPS, and data for the years 2002 – 2011 were pooled together to allow for a sufficiently large sample size to ensure reliable estimates.

Study Population

The broad classification of “cardiovascular disease (CVD) and associated comorbidities” in this study consisted for the following disease conditions: acute myocardial infarction (Clinical Classification code (CCC): 100), coronary artery disease (CCC: 101), congestive heart failure (CCC: 108), peripheral and visceral atherosclerosis (CCCs: 114, 115, 116), stroke (CCCs: 109, 112), hypertension (CCC: 098), lipid disorders (CCC: 053), and diabetes (CCC: 049, 050). CCC aggregates conditions and procedures into mutually exclusive and clinically homogeneous categories, using Clinical Classification Software.⁴⁰ When examining utilization of prescription medications in the form of number of medication refills, the study population included respondents who were: (1) aged 18 to 64 years, (2) reported either being diagnosed with CVD or high-risk conditions for CVD for CVD in their first round of the MEPS survey, or reported having an event (an inpatient, emergency, outpatient, or office-based provider visits, or a prescription medication) associated with one or more of the aforementioned disease conditions at any point of time during the survey, (3) reported having had Medicaid coverage at any point of time during the survey year, and (4) had positive person weights, and were interviewed for all five rounds of the 2-year MEPS survey interviews. When examining adherence to commonly prescribed medication classes for managing the CVD and associated risk factor comorbidities listed above, an additional inclusion criteria was considered: (5) respondents who had one or more fills or refills for the following therapeutic drug classes – angiotensin converting enzymes inhibitors (ACEI), angiotensin II receptor blockers (ARB), beta blockers, calcium channel blockers, diuretics, anti-hypertensive drug combinations, HMG-CoA reductase inhibitors (statins), sulfonylureas, and biguanides.

Respondents who reported having Medicare, or pregnancy were excluded from this study. Pregnant women with incomes at or below 133 percent of the federal poverty level are eligible for Medicaid coverage throughout the pregnancy; however, the coverage ceases 60 days postpartum.⁶⁸ Low-income non-elderly adults with disabilities are also eligible for Medicare if they have received monthly social security disability income for two years. Most coverage benefits for prescription drugs for dual eligible beneficiaries are covered by Medicare Part D, instead of Medicaid. As a result, these two population groups were excluded from this study.

Individuals who have less than a year of Medicaid coverage and experience within-year Medicaid discontinuity may not be a homogenous group, especially when they report having other forms of insurance coverage during the year. In order to account for these differences, individuals in the study population who reported having <12 months of Medicaid coverage during a survey year were divided into two groups: (1) individuals with <12 months of Medicaid coverage who did not report any other sources of health insurance during the survey year, designated in this study as ‘Discontinuous–Uninsured’, and (2) individuals with <12 months of Medicaid coverage who also reported having other source(s) of health insurance during the survey year (either private, Tricare, or any hospital- or physician group-based health insurance), designated in this study as ‘Discontinuous–Insured’. Individuals in the study population who reported having Medicaid coverage during all 12 months of a survey year were classified as having continuous Medicaid coverage, which was the comparison group and designated as ‘Continuous Medicaid’ in this study.

Study Variables

Our first objective was to examine whether Medicaid discontinuity was associated with adherence to commonly prescribed classes of medications among non-elderly adults with CVD or its associated risk factor comorbidities. The primary outcomes of interest was the binary indicator for medication adherence measured as the medication possession ratio (MPR), or the proportion of days during the survey year during which respondents had medications on hand. In the subsequent section, we describe how adherence was measured for this study. The second objective was to examine whether Medicaid discontinuity impacts medication utilization. The primary outcome of interest for this objective was the raw number of refills for all medications reported by the study respondents. We divided this utilization measure into all-cause, and disease-specific medication utilization, i.e. number of refills for all medications reported by the study respondents in the survey year (all-cause), and the number of refills for medications specifically prescribed for CVD or high-risk conditions for CVD for CVD that were examined in this study (disease-specific).

All regression models included the following classes of covariates:-

- 1) Demographic covariates: age (in years), gender (female vs. male), race/ethnicity (Non-Hispanic White vs. Non-Hispanic Black, Hispanic, and other races), marital

- status (not married vs. married), family size (N), and geographic residence (Northeast vs. Midwest, South, and West).
- 2) Socioeconomic covariates: education (less than high school vs. high school/GED, and Bachelor's or higher), income as a percentage of the federal poverty level (FPL) (<100% FPL vs. 100 - <125% FPL, and $\geq 125\%$ FPL), employment (always unemployed vs. intermittent employment, and always employed), binary indicator for Metropolitan Statistical Area (MSA) residence.
 - 3) Medicaid-related covariates: binary indicator for Supplementary Security Income (SSI) for disability receipt or participation in the Temporary Assistance for Needy Families program, and a binary indicator for Medicaid HMO/Managed Care participation.
 - 4) Health-related covariates: perceived health status (excellent/very good/good vs. fair/poor), perceived mental health status (excellent/very good/good vs. fair/poor) binary indicator for mental illness/substance abuse diagnosis, binary indicator for respiratory diseases diagnosis, binary indicator for arthritis/joint pain diagnosis, and Charlson Comorbidity Index (CCI) score.

When examining medication adherence, the regression models included the following additional covariates:

- 5) Medication-related covariates: number of unique prescription medications reported by the study respondents (N), and a binary indicator for whether individuals had a 90 days' supply of medications.

All regression models included the year fixed effects to capture any macro-level variations during the study period. SAS 9.3 and STATA 13 were used.

Measurement of Adherence

The measurement of adherence in this study involved several intermediate steps. From the Prescribed Medicines files in MEPS, this study used the drug refill records for each study respondent, to estimate adherence to each medication class during the year when the respondents had refills for any medication from that class of medication. Of the two most common measures of adherence, medication possession ratio (MPR) and proportion of days covered (PDC), this study used the MPR measure.^{69,70} Using an interval-based MPR measure approach, the denominator for this measure was the number of days between the index date, which was first day of the round in which the first fill or refill of a class of drug (e.g. beta blockers) was reported, and the last day of the survey year, i.e. December 31st. The numerator is the total number of days covered by drug fills during the denominator duration, i.e. the total days' supply for that particular class of medication within the follow-up (denominator) period. **Equation 3-1** displays the average MPR estimation:

$$\frac{\sum total\ Rx\ days\ of\ supply_n}{Interval_n\ (round\ date\ for\ first\ Rx\ during\ year - December\ 31)} \Bigg/ n \quad (Eq.\ 3-1)$$

As mentioned earlier, the proportion of the total days' supply for a therapeutic class of medication was estimated within the follow-up interval reported by an individual. Depending on the number of therapeutic classes reported by each individual, the MPR was estimated for each class and summed, and then divided by the number of therapeutic classes reported by individuals (n) to obtain an average MPR measure for that individual. This enables capturing non-adherence, since any discontinuation in the drug refill will be captured in the calculation as it tracks adherence till the end of the duration and not the last drug refill during the measurement period, as is common with the standard measures of MPR.⁶⁹ All MPR measures were expressed as a percentage and capped at 100% (or 1.0). Drugs that were dispensed within a therapeutic class (e.g. switching from carvedilol to propranolol among beta blockers) were considered interchangeable.⁶⁹ ACEI and ARBs were considered a single class for this study.

More than 90% of index refills had occurred in the first two rounds of the calendar year. The drug class-specific MPRs that were obtained were summed, and an overall average MPR measure was calculated to summarize adherence to all drug classes an individual was taking. In order to make policy-relevant conclusions about the impact of insurance discontinuity on medication adherence, the continuous MPR measure was transformed into a binary MPR measure, with a cutoff of 80% (0.8) or more signifying adequate adherence to prescribed medications examined in this study. The main outcome variable was the overall binary MPR measure. Class-specific MPR adherence measures were examined in secondary analyses.

Prior to 2010, MEPS did not report data on days' supply for the prescription medications reported by the respondents during the survey year. Previous studies have overcome this shortcoming, by developing algorithms for days' supply based on the quantity of drug dispensed using data from administrative claims, or other similar sources. A similar approach was applied; however, rather than using an external source of data, the 2010 and 2011 MEPS prescription medication data that had the days' supply information was used to develop an algorithm that approximates days' supply for previous years in MEPS based on their quantity prescribed/dispensed, which was available for all the 10 years in MEPS. The most frequent number of supply days furnished by each level of dispensed quantity was first identified, and found that 30, and 90 days' supply were the most common patterns observed for most quantities dispensed. After deriving a method to smoothly approximate the distribution of the days' supply, the days of supply were categorized using different sets of categories. The third categorization algorithm was found to have the highest correlation with the actual MPR measures, both continuous and binary, which were calculated using the days' supply variables from 2011 and 2010 years of MEPS data, separately as displayed in **Table 3-1** and **Table 3-2**. Finally, using this algorithm, the MPR adherence measures for all years prior to 2010 were calculated to obtain overall estimates of medication adherence.

Table 3-1. Refill Days of Supply Algorithms and MPR Validation in 2011 Sample (N= 5,475)

Categorization	Quantity Dispensed	Corresponding Days' Supply Algorithm	Continuous MPR Measure Pearson's Correlation	Binary MPR Measure	
				Observed Agreement (%)	κ- Statistic
Categorization 1	≤75	30	0.868	0.925	0.841***
	>75	90			
Categorization 2	≤50	30	0.881	0.931	0.854***
	>50 – ≤80	60			
	>80	90			
Categorization 3	≤44	30	0.889	0.934	0.864***
	45	90			
	>45 – ≤75	30			
	>75 – ≤119	90			
	120	30			
	>120	90			

MPR: Medication Possession Ratio

*** $P < 0.001$. The null hypothesis for estimating kappa statistic is that there is no agreement between the two discrete measures of MPR.

Table 3-2. Refill Days of Supply Algorithms and MPR Validation in 2010 Sample (N= 5,058)

Categorization	Quantity Dispensed	Corresponding Days' Supply Algorithm	Continuous MPR Measure Pearson's Correlation	Binary MPR Measure	
				Observed Agreement (%)	κ- Statistic
Categorization 1	≤75	30	0.835	0.918	0.827***
	>75	90			
Categorization 2	≤50	30	0.856	0.876	0.747***
	>50 – ≤80	60			
	>80	90			
Categorization 3	≤44	30	0.868	0.924	0.841***
	45	90			
	>45 – ≤75	30			
	>75 – ≤119	90			
	120	30			
	>120	90			

MPR: Medication Possession Ratio

*** $P < 0.001$. The null hypothesis for estimating kappa statistic is that there is no agreement between the two discrete measures of MPR.

Statistical Analysis

Bivariate statistics were conducted when examining the characteristics of the study population. Survey-weighted proportions of the use of different classes of medications examined in this study were calculated. When examining medication adherence, Survey-weighted multivariate logistic regression analyses were conducted. **Equation 3-2** and **Equation 3-3** depict the logistic regression models to examine adherence among both the Discontinuous–Uninsured, and the Discontinuous–Insured groups relative to the Continuous Medicaid insurance group.

$$Pr[MPR \geq 0.8 | x] = F^{-1}(\hat{\alpha} + \hat{\beta}_1 \cdot Discontinuous - Uninsured + \sum_{k=1}^K \hat{\beta}_k X_k) \quad (\text{Eq. 3-2})$$

$$Pr[MPR \geq 0.8 | x] = F^{-1}(\hat{\alpha} + \hat{\beta}_1 \cdot Discontinuous - Insured + \sum_{k=1}^K \hat{\beta}_k X_k) \quad (\text{Eq. 3-3})$$

Medication adherence is modeled as a dichotomous outcome variable as a function of Medicaid discontinuity as well as the demographic, socioeconomic, eligibility, health-related, and medication-related covariates mentioned above. F^{-1} denotes the inverse of the cumulative standard logistic distribution function that relates the outcome on the probability scale to the covariates. $Pr[]$ denotes the population average probability of adherence to medications conditional on the covariates denoted by x . *Discontinuous–Uninsured* and *Discontinuous–Insured* are binary indicator variables for the main independent variable of interest, that equals one for individuals having discontinuous annual Medicaid coverage without any other sources of insurance and with other sources of insurance during the year, respectively, and equals zero for the comparison group of individuals with continuous Medicaid coverage. X denotes the vector of aforementioned individual characteristics that predict Medicaid discontinuity in the two population groups of interest.

When examining all-cause and disease-specific medication utilization using the number of prescription drug fills as the outcome, the negative binomial regression model was utilized due to the skewness in the distribution exhibited by the outcomes measures. **Appendix A, Figures A-1 to A-4**, display the distribution in the all-cause and disease-specific prescription drug utilization outcome measures. It is a type of generalized linear model where the dependent variable Y is a count of the number of times an event occurs (in this case, the number of drug fills). A convenient parameterization of the negative binomial distribution is denoted by **Equation 3-4**.⁷¹

$$p(y) = P(Y = y) = \frac{\Gamma(y + \frac{1}{\alpha})}{\Gamma(y+1)\Gamma(\frac{1}{\alpha})} \left(\frac{1}{1+\alpha\mu}\right)^{1/\alpha} \left(\frac{\alpha\mu}{1+\alpha\mu}\right)^y \quad (\text{Eq. 3-4})$$

In **Equation 3-4**, $\mu > 0$ is the mean of Y , which is the count variable indicating the number of prescription drug fills (either all-cause, or disease-specific) and $\alpha > 0$ is the heterogeneity parameter. It can be derived as a Poisson-gamma mixture, or as the number of failures before the $(1/\alpha)^{\text{th}}$ success. A simpler depiction of the traditional negative binomial regression model, the NB2 model is depicted in **Equation 3-5** and **Equation 3-6** as follows:

$$\ln \mu_{all-cause} = \beta_0 + \beta_1 Discontinuous - Uninsured + \beta_2 \cdot x_2 + \dots + \beta_k \cdot x_k \quad (\text{Eq. 3-5})$$

$$\ln \mu_{disease-specific} = \beta_0 + \beta_1 Discontinuous - Uninsured + \beta_2 \cdot x_2 + \dots + \beta_k \cdot x_k \quad (\text{Eq. 3-6})$$

Apart from the main indicator variables for the Discontinuous–Uninsured group, the covariates x_2 to x_k are known, and the population regression coefficients β_0 to β_k are to be estimated. The two equations above depict the negative binomial regression models to estimate the differences in all-cause prescription drug utilization (**Equation 3-5**), and disease-specific prescription drug utilization (**Equation 3-6**) between the Continuous Medicaid, and Discontinuous–Uninsured groups. Similar regression models (not shown here) were run to estimate differences in prescription drug utilization between the Continuous Medicaid and Discontinuous–Insured groups. The coefficient of interest for the Discontinuous–Uninsured is the measure of differences in the utilization of overall, and disease-specific drug utilization. Because of the functional form of the negative binomial regression, the effect size of a factor is the antilog of its coefficient. This estimated coefficients are transformed to obtain the incidence rate ratio (IRR). An IRR of less than one for the discontinuous Medicaid groups would suggest lower medication utilization among them when compared to their counterparts with continuous Medicaid coverage.

In addition to obtaining IRRs, this study further estimated the average marginal effect to determine the predicted difference in the number of all-cause and disease-specific prescription drug fills between the continuous and discontinuous Medicaid groups. The marginal effects measure the discrete change in the outcome variable, i.e. for a binary independent variable, how do predicted probabilities change as the binary independent variable changes from 0 (Continuous Medicaid) to 1 (Discontinuous–Uninsured, or Discontinuous–Insured).⁷² In order to obtain these marginal effects, after developing and running the multivariate negative binomial regression model, the estimated coefficients of the model were used to calculate two predicted prescription drug fills for all individuals in the study population. The first predicted utilization assumed all individuals had discontinuous Medicaid coverage (by setting the Medicaid discontinuity indicator variable to 1), and the second predicted utilization assumed all individual to have continuous full-year Medicaid coverage (by setting the Medicaid discontinuity indicator variable to 0). The average per-person difference in prescription drug fills attributable to Medicaid coverage was estimated by obtaining the difference in predicted drug utilization for each person and computing the weighted average of the difference across the entire study sample. All data analyses were conducted using SAS 9.3 and STATA 13 to account for the complex survey design of MEPS.

RESULTS

Characteristics of the Study Populations

The characteristics of the study population when examining adherence to commonly prescribed medications to manage CVD or high-risk conditions for CVD, and when examining utilization of prescription medications, are displayed in **Table 3-3** and **Table 3-4**, respectively. The study sample when examining overall medication adherence included 3,210 respondents in the Continuous Medicaid coverage group (nationally representative of about 2 million individuals), 1,088 individuals in the Discontinuous–Uninsured group (nationally representative of over 700,000 individuals), and 388 individuals in the Discontinuous–Insured group (nationally representative of over 300,000 individuals). When examining utilization of prescription medications in the study population, the study sample included 3,926 respondents in the Continuous Medicaid coverage group (nationally representative of about 2.4 million individuals), 1,407 individuals in the Discontinuous–Uninsured group (nationally representative of over 900,000 individuals), and 493 individuals in the Discontinuous–Insured group (nationally representative of over 400,000 individuals). The difference being, when examining medication adherence there was an additional inclusion criteria; the study population consisted of individuals who reported taking one or more of the eight classes of medications for which adherence was examined.

When examining medication adherence (**Table 3-3**), in comparison to the Continuous Medicaid group, individuals in the Discontinuous–Uninsured group were less likely to belong to other race groups ($P=0.016$), and more likely to be married ($P<0.001$), and were more likely to belong to the South geographic region ($P=0.010$). Among the socioeconomic covariates, this study found that respondents were more likely to have completed high school/GED or college ($P=0.006$), were more likely to have family incomes $\geq 125\%$ FPL ($P<0.001$), and more likely to have intermittent, as well as continuous full-year employment ($P<0.001$). Respondents were less likely to be eligible for SSI due to disability or participate in TANF ($P<0.001$), and less likely to report Medicaid HMO/managed care ($P<0.001$). Among the health-related variables, respondents were less likely to report having fair or poor perceived mental health status ($P=0.001$), less likely to report having mental /substance abuse illnesses ($P=0.014$), and less likely to report having respiratory illnesses ($P<0.001$). The mean CCI score was not significantly different between the two groups ($P=0.181$). When examining the medication-related variables, this study found that although the Discontinuous–Uninsured group on average were prescribed fewer medications than the Continuous Medicaid group (7.6 vs. 9.1 medications, $P<0.001$), and both groups had similar proportion of 90 days' supply for their medications ($P=0.242$).

When comparing the Discontinuous–Insured group to the Continuous Medicaid group, this study found that among the demographic covariates, individuals in the Discontinuous–Insured group were more likely to be younger ($P=0.034$), were less likely to belong to minority race/ethnicity groups ($P<0.001$), more likely to be married

Table 3-3. Study Population Characteristics when Examining Medication Adherence

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
DEMOGRAPHIC								
Age					0.876			0.014
18 – 34 years	248,650	12.5 (0.9)	93,295	13.1 (1.3)		65,519	19.6 (2.4)	
35 – 49 years	669,245	33.8 (1.3)	242,313	34.2 (2.1)		100,718	30.1 (3.0)	
50 – 64 years	1,065,905	53.7 (1.4)	373,450	52.7 (2.3)		168,190	50.3 (3.3)	
Gender					0.939			0.190
Female	1,334,759	67.3 (1.3)	475,870	67.1 (2.0)		210,679	63.0 (3.2)	
Race/Ethnicity					0.016			<0.001
Non-Hispanic White	894,528	45.1 (1.9)	336,390	47.4 (2.3)		211,937	63.4 (3.3)	
Non-Hispanic Black	562,325	28.3 (1.5)	183,702	25.9 (1.9)		64,835	19.4 (2.2)	
Hispanic	341,418	17.2 (1.2)	145,309	20.5 (1.6)		34,802	10.4 (1.6)	
Others	185,529	9.4 (0.9)	43,657	6.2 (1.0)		22,852	6.8 (1.5)	
Marital Status					<0.001			<0.001
Married	602,503	30.4 (1.4)	271,354	38.3 (2.1)		154,101	46.1 (3.6)	
Region					0.064			0.040
Northeast	481,061	24.3 (1.7)	141,392	19.9 (2.0)		82,125	24.5 (3.1)	
Midwest	393,671	19.8 (1.4)	139,468	19.7 (1.8)		92,872	27.8 (4.0)	
South	607,916	30.6 (1.7)	255,996	36.1 (2.0)		76,828	23.0 (2.7)	
West	501,152	25.3 (1.7)	172,201	24.3 (2.0)		82,602	24.7 (3.4)	
Family Size Mean (SE)	2.7 (0.1)		2.8 (0.1)		0.532	2.8 (0.1)		0.449

Table 3-3. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	P-value [†]	Weighted N	% (SE)	P-value ^{††}
SOCIOECONOMIC								
Education					0.006			<0.001
<High School	776,408	39.5 (1.5)	226,415	32.2 (1.9)		51,716	15.5 (2.1)	
High School/GED	997,819	50.7 (1.4)	392,411	55.7 (2.2)		176,153	52.8 (3.6)	
College Degree	193,484	9.8 (0.8)	85,248	12.1 (1.6)		105,631	31.7 (3.6)	
Income					<0.001			<0.001
<100% FPL	1,113,782	56.1 (1.3)	324,544	45.8 (2.0)		74,686	22.3 (2.3)	
100 – <125% FPL	185,942	9.4 (0.6)	69,730	9.8 (1.2)		24,692	7.4 (1.5)	
≥125% FPL	684,076	34.5 (1.2)	314,784	44.4 (2.1)		235,048	70.3 (2.8)	
Employment					<0.001			<0.001
Always Unemployed	1,468,299	74.0 (1.3)	399,549	56.3 (2.0)		91,161	27.3 (2.5)	
Unstable Employment	231,945	11.7 (0.9)	165,706	23.4 (1.9)		102,142	30.5 (2.9)	
Always Employed	283,555	14.3 (1.0)	143,803	20.3 (1.5)		141,123	42.2 (3.2)	
MSA	1,564,442	78.9 (1.8)	552,171	77.9 (2.0)	0.651	267,784	80.1 (2.7)	0.696
MEDICAID-RELATED								
SSI/TANF	852,158	43.0 (1.5)	127,179	17.9 (1.5)	<0.001	30,671	9.2 (1.8)	<0.001
Medicaid HMO/MCO	1,139,857	60.4 (1.6)	354,416	50.3 (2.1)	<0.001	180,662	54.0 (3.4)	0.077

Table 3-3. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
HEALTH-RELATED								
Perceived Health Status					0.050			<0.001
Fair to Poor	1,112,169	56.1 (1.4)	366,568	51.7 (2.1)		140,242	41.9 (2.8)	
Perceived Mental Health					0.001			<0.001
Fair to Poor	670,915	33.9 (1.4)	189,140	26.7 (1.9)		62,634	18.7 (2.4)	
CVD or high-risk conditions for CVD					0.998			0.066
≤1	1,281,387	64.6 (1.3)	457,590	64.5 (2.0)		240,536	71.9 (2.7)	
2	400,899	20.2 (0.9)	144,019	20.3 (1.6)		58,036	17.4 (2.5)	
≥3	301,513	15.2 (1.0)	107,449	15.2 (1.5)		35,854	10.7 (1.5)	
Mental/Substance Abuse Illnesses	606,104	30.6 (1.4)	177,733	25.1 (1.8)	0.014	94,737	28.3 (2.8)	0.463
Respiratory Illnesses	394,992	19.9 (1.2)	87,482	12.3 (1.2)	<0.001	32,380	9.7 (1.8)	<0.001
Arthritis/Joint Pain	693,646	35.0 (1.4)	221,951	31.3 (1.9)	0.073	97,019	29.0 (3.0)	0.073
CCI Score Mean (SE)	1.2 (0.0)		1.1 (0.1)		0.181	1.1 (0.1)		0.058
Population Size	1,983,799	65.5 (1.1)	709,058	23.4 (0.9)		334,426	11.1 (0.8)	

Table 3-3. (Continued)

Study population includes MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, had ≤ 1 month of Medicaid coverage during the survey year, and who consumed one or more medications from the following therapeutic drug classes: ACEI/ARB, beta blockers, calcium channel blockers, diuretics, anti-HTN combinations, statins, sulfonylureas, and biguanides.

[†]*P*-value for the difference between Discontinuous–Uninsured and Continuous Medicaid groups. ^{††}*P*-value for the difference between Discontinuous–Insured and Continuous Medicaid groups. *P*-values in **bold** indicate statistically significant difference below the threshold value of *P*<0.05.

SE: Standard error; GED: General educational development; FPL: Federal poverty level; MSA: Metropolitan Statistical Area; SSI: Supplementary Security Income due to disability; TANF: Temporary Assistance to Needy Families program participation; HMO: Health maintenance organization; MCO: Managed care organization; CCI: Charlson Comorbidity Index score (D’Hoore adapted CCI score was calculated in this study using the 3-digit ICD-9-CM codes in MEPS).

Table 3-4. Study Population Characteristics when Examining Medication Utilization

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
DEMOGRAPHIC								
Age					0.385			0.013
18 – 34 years	388,425	15.9 (1.0)	165,340	18.1 (1.3)		98,188	23.1 (2.3)	
35 – 49 years	875,809	35.9 (1.2)	324,653	35.5 (1.8)		141,109	33.3 (2.7)	
50 – 64 years	1,173,144	48.2 (1.3)	423,933	46.4 (2.0)		184,968	43.6 (3.0)	
Gender					0.491			0.169
Female	1,632,006	67.0 (1.2)	624,410	68.3 (1.8)		267,502	63.1 (2.8)	
Race/Ethnicity					<0.001			<0.001
Non-Hispanic White	1,094,002	44.9 (1.8)	442,534	48.4 (2.1)		256,231	60.4 (3.2)	
Non-Hispanic Black	666,358	27.3 (1.5)	219,033	24.0 (1.6)		79,936	18.8 (2.0)	
Hispanic	445,601	18.3 (1.2)	202,106	22.1 (1.5)		50,704	12.0 (1.6)	
Others	231,416	9.5 (0.9)	50,253	5.5 (0.9)		37,393	8.8 (1.7)	
Marital Status					<0.001			<0.001
Married	727,457	29.9 (1.3)	344,150	37.7 (1.8)		184,414	43.5 (3.2)	
Region					0.049			0.023
Northeast	608,070	25.0 (1.6)	187,785	20.5 (1.7)		104,394	24.6 (2.9)	
Midwest	477,401	19.6 (1.4)	184,710	20.2 (1.7)		115,146	27.2 (3.2)	
South	727,520	29.8 (1.6)	315,073	34.5 (1.8)		98,072	23.1 (2.3)	
West	624,386	25.6 (1.7)	226,358	24.8 (1.9)		106,652	25.1 (3.0)	
Family Size Mean (SE)	2.8 (0.1)		2.9 (0.1)		0.409	2.8 (0.1)		0.976

Table 3-4. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
SOCIOECONOMIC								
Education					0.001			<0.001
<High School	971,041	40.2 (1.4)	298,300	32.9 (1.8)		67,896	16.0 (1.9)	
High School/GED	1,215,968	50.4 (1.4)	498,289	54.9 (2.0)		232,334	54.9 (3.2)	
College Degree	227,668	9.4 (0.7)	111,300	12.2 (1.4)		123,109	29.1 (3.2)	
Income					<0.001			<0.001
<100% FPL	1,384,551	56.8 (1.2)	428,540	46.9 (1.7)		104,719	24.7 (2.3)	
100 – <125% FPL	227,591	9.3 (0.6)	91,940	10.1 (1.1)		28,783	6.8 (1.3)	
≥125% FPL	825,236	33.9 (1.1)	393,445	43.0 (1.8)		290,763	68.5 (2.6)	
Employment					<0.001			<0.001
Always Unemployed	1,767,234	72.5 (1.2)	496,282	54.3 (1.8)		112,490	26.5 (2.2)	
Unstable Employment	301,597	12.4 (0.8)	228,246	25.0 (1.6)		129,583	30.6 (2.5)	
Always Employed	368,547	15.1 (0.9)	189,398	20.7 (1.4)		182,191	42.9 (2.9)	
MSA	1,931,797	79.3 (1.8)	721,532	78.9 (1.9)	0.880	344,287	81.1 (2.7)	0.494
MEDICAID-RELATED								
SSI/TANF	1,012,240	41.5 (1.3)	154,157	16.9 (1.3)	<0.001	41,212	9.7 (1.6)	<0.001
Medicaid HMO/MCO	1,424,726	61.3 (1.5)	459,330	50.5 (2.0)	<0.001	233,429	55.0 (3.0)	0.048

Table 3-4. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
HEALTH-RELATED								
Perceived Health Status					0.162			<0.001
Fair to Poor	1,314,794	53.9 (1.2)	467,555	51.2 (1.9)		170,598	40.2 (2.5)	
Perceived Mental Health					<0.001			<0.001
Fair to Poor	821,318	33.7 (1.2)	234,931	25.7 (1.7)		84,716	20.0 (2.2)	
CVD or high-risk conditions for CVD					0.775			0.037
≤1	1,678,222	68.9 (1.1)	640,242	70.1 (1.8)		321,685	75.8 (2.3)	
2	444,995	18.2 (0.8)	163,062	17.8 (1.4)		65,581	15.5 (2.2)	
≥3	314,160	12.9 (0.8)	110,621	12.1 (1.3)		37,000	8.7 (1.5)	
Mental/Substance Abuse Illnesses	751,357	30.8 (1.2)	227,885	24.9 (1.6)	0.003	115,891	27.3 (2.4)	0.176
Respiratory Illnesses	479,702	19.7 (1.1)	117,992	12.9 (1.1)	<0.001	39,380	9.3 (1.5)	<0.001
Arthritis/Joint Pain	830,408	34.1 (1.3)	267,358	29.3 (1.7)	0.009	107,575	25.4 (2.4)	0.002
CCI Score Mean (SE)	1.2 (0.0)		1.1 (0.1)		0.079	0.9 (0.1)		0.003
Population Size	2,437,378	64.6 (1.0)	913,926	24.2 (0.8)		424,265	11.2 (0.7)	

Table 3-4. (Continued)

Study population includes MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, had ≤ 1 month of Medicaid coverage during the survey year.

[†]*P*-value for the difference between Discontinuous–Uninsured and Continuous Medicaid groups. ^{††}*P*-value for the difference between Discontinuous–Insured and Continuous Medicaid groups. *P*-values in **bold** indicate statistically significant difference below the threshold value of $P < 0.05$. SE: Standard error; GED: General educational development; FPL: Federal poverty level; MSA: Metropolitan Statistical Area; SSI: Supplementary Security Income due to disability; TANF: Temporary Assistance to Needy Families program participation; HMO: Health maintenance organization; MCO: Managed care organization; CCI: Charlson Comorbidity Index score (D’Hoore adapted CCI score was calculated in this study using the 3-digit ICD-9-CM codes in MEPS).

($P<0.001$). Among the socioeconomic covariates, this study found that individuals in the Discontinuous–Insured group were more likely to complete high school/GED or college ($P<0.001$), more likely to have higher family income levels ($P<0.001$), and were more likely to have some form of employment during the year ($P<0.001$). The Discontinuous–Insured group was also to be eligible for SSI due to disability or participate in TANF ($P<0.001$), however, there was no significant difference between having Medicaid through an HMO/MCO ($P=0.087$). Among the health-related covariates, this study found that the Discontinuous–Insured group was less likely to have respiratory illnesses ($P<0.001$), less likely to report having fair to poor health status ($P<0.001$), or fair to poor mental health status ($P<0.001$), and was similar to the Continuous Medicaid coverage group in terms of having mental/substance abuse illnesses ($P=0.428$), arthritis/joint pain ($P=0.086$), and mean CCI score or burden of disease ($P=0.078$). Finally, when examining the medication-related variables, the Discontinuous–Insured group was found to be prescribed lower number of unique medications on average ($P<0.001$), and more likely to have a 90-day supply of medications ($P<0.001$) than the Continuous Medicaid group.

For the second aim of examining medication utilization, the patterns in the demographic, socioeconomic, Medicaid -, and health-related characteristics of the study population were found to be similar to the patterns described above, with some exceptions. When examining geographic census region, individuals in the Discontinuous – Uninsured group were found to be more likely to reside in the South, compared to their counterparts in the Continuous Medicaid group. However, the magnitude the statistical significance in difference was small ($P=0.049$). They were also less likely to have mental/substance abuse illnesses, and arthritis/joint pain. Among the Discontinuous – Uninsured group, individuals were less likely to participate in Medicaid administered by an HMO/MCO ($P=0.048$), less likely to have 2 or more CVD or high-risk conditions for CVD ($P=0.037$), less likely to have arthritis/joint pain ($P=0.002$), and had a lower disease burden or mean Charlson score ($P=0.003$), when compared to those with continuous Medicaid coverage.

Prescription Medication Use in the Study Population

Table 3-5 displays the summary statistics for prescription medication use in the study population. The distribution of the classes of medications prescribed was similar among all three groups; the Continuous Medicaid, Discontinuous–Uninsured, and the Discontinuous–Insured. Statins and thiazolidinediones were the most commonly prescribed classes of medications, followed by diuretics and anti-hypertensive combination medications. On average, the Continuous Medicaid group reported higher number of all-cause prescription fills annually than the Discontinuous–Uninsured (46.0 vs. 30.0 fills, $P<0.001$), and the Discontinuous–Insured groups (46.0 vs. 30.0 fills, $P<0.001$). Similarly, the Continuous Medicaid group reported higher mean disease-specific (CVD or high-risk conditions for CVD) prescription fills annually than the Discontinuous–Uninsured (16.6 vs. 13.5 fills, $P<0.001$), and the Discontinuous–Insured groups (16.6 vs. 12.8 fills, $P<0.001$). The proportion of days during which

Table 3-5. Summary Statistics for Prescription Medication Use in the Study Population

Measures	Therapeutic Drug Classes	Continuous Medicaid		Discontinuous–Uninsured		Discontinuous–Insured	
		Weighted N or Mean	% (SE) or SE	Weighted N or Mean	% (SE) or SE	Weighted N or Mean	% (SE) or SE
	ACEI/ARB	136,818	6.9 (0.7)	47,265	6.7 (1.0)	28,419	8.5 (1.8)
	Beta blockers	147,001	7.4 (0.7)	57,759	8.1 (1.2)	37,979	11.4 (2.3)
	CCB	125,709	6.3 (0.6)	57,759	6.3 (0.9)	13,328	4.0 (1.1)
	Diuretics	288,166	14.5 (1.0)	105,673	14.9 (1.5)	39,934	11.9 (2.2)
	Anti-HTN combinations	197,107	9.9 (0.7)	72,220	10.2 (1.0)	42,047	12.6 (2.0)
	Statins	538,014	27.1 (1.1)	167,148	23.6 (1.8)	85,635	25.6 (2.9)
	Sulfonylureas	114,938	5.9 (0.6)	43,755	6.2 (1.0)	16,097	4.8 (1.6)
	Biguanides	436,045	22.0 (1.1)	170,483	24.0 (2.0)	70,987	21.2 (2.7)
Mean No. of Drug Classes		2.4	0.02	2.0	0.02	1.6*	0.1
Mean No. of All-Cause Drug Fills		46.0	1.1	30.0***	1.3	30.0***	1.5
Mean No. of Disease-Specific Drug Fills		16.6	0.5	13.5***	0.6	12.8***	0.8

Study population included MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid anytime during the year.

Estimates in **bold** indicate statistical significance at $P < 0.05$.

ACEI: Angiotensin converting enzyme inhibitors, ARB: Angiotensin II converting enzyme receptor blockers, CCB: Calcium channel blockers, HTN: Hypertension; SE: Standard Error. * $P < 0.05$, ** $P < 0.01$, *** $P < 0.001$ for comparison between the continuous and discontinuous Medicaid groups.

respondents had possession of medications was significantly higher among the Continuous Medicaid group when compared to the Discontinuous–Uninsured (67.8% vs. 64.9%, $P<0.001$), but was not significantly different from the Discontinuous–Insured group (67.8% vs. 68.4%, $P=0.710$) (results not shown).

The results for medication adherence outcomes, both overall and therapeutic drugclass-specific, are displayed in **Table 3-6**. When examining overall adherence using the average MPR measure, this study found that although the Discontinuous–Uninsured group had about 17% lower odds of being adherent to their medications than the Continuous Medicaid group, this difference was not statistically significant [Odds Ratio (OR): 0.83; 95% Confidence Interval (CI): 0.69 – 1.00]. Difference in overall adherence was found to be highly insignificant when comparing Discontinuous–Insured and Continuous Medicaid groups ($P=0.775$). **Appendix Table A-1** displays the full model coefficients exponentiated as odds ratios for the comparisons between Continuous Medicaid, Discontinuous–Uninsured, and the Discontinuous–Insured. Among the covariates, having fair or poor perceived health status, arthritis/joint pain, and a 90 days' supply of prescription drugs were associated with higher odds of being adherent to commonly prescribed medications that are essential for managing CVD or high-risk conditions for CVD. When examining class-specific medication adherence, this study found that the Discontinuous–Uninsured group were 41% less likely to be adherent to sulfonylureas when compared to the Continuous Medicaid group (OR: 0.59; 95% CI: 0.38 – 0.92). Adherence to the remaining classes of medications examined was not significantly different between these two groups. When examining class-specific adherence among the Discontinuous–Insured, this study found that this group was 43% less likely to be adherent to diuretics (OR: 0.57; 95% CI: 0.33 – 0.98), and had almost 3 times higher odds of being adherent to calcium channel blockers (OR: 2.86; 95% CI: 1.44 – 5.69).

Table 3-7 displays the results for the multivariate negative binomial regression analyses to examine medication utilization. This study found that the incidence of overall prescription fills (all-cause) was 27% lower among the Discontinuous–Uninsured group compared to the Continuous Medicaid group, or in other words, the expected number of all-cause prescription fills among the Discontinuous–Uninsured group were 27% less when compared to the Continuous Medicaid group [Incidence Rate Ratio (IRR): 0.73; 95% CI: 0.68 – 0.79]. The average incremental difference in overall prescription medication fills between the Discontinuous–Uninsured and Continuous Medicaid group was found to be -12.0 fills (Standard Error (SE): 1.4 fills; $P<0.001$). In other words, individuals in the Discontinuous–Uninsured group were predicted to have 12 fewer prescription drug fills per person annually, compared to the Continuous Medicaid group. Further, it was found that the Discontinuous–Uninsured group had about 12% lower prescription drug fills annually specific to the CVD or high-risk conditions for CVD for CVD considered in this study (disease-specific), when compared to the Continuous Medicaid group (IRR: 0.88; 95% CI: 0.80 – 0.97). When estimating the average marginal effect of Medicaid discontinuity, the Discontinuous–Uninsured group was predicted to have about 2.0 fewer disease-specific prescription drug fills per person annually (SE: 0.76 fills; $P=0.009$), compared to the Continuous Medicaid group.

Table 3-6. Association between Medicaid Discontinuity and Prescription Medication Adherence in the Study Population

Outcome	Medicaid Coverage	Odds Ratio (95% Confidence Interval)
Average Adherence (MPR: <0.8 v ≥0.8)	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.83 (0.69 – 1.00)
	Discontinuous–Insured	1.04 (0.79 – 1.36)
Therapeutic Class-Specific Adherence (MPR: <0.8 v ≥0.8)		
ACEI/ARB	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.97 (0.74 – 1.28)
	Discontinuous–Insured	1.56 (1.02 – 2.39)*
Beta Blockers	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.74 (0.53 – 1.04)
	Discontinuous–Insured	0.84 (0.51 – 1.36)
Calcium Channel Blockers	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	1.20 (0.86 – 1.66)
	Discontinuous–Insured	3.06 (1.56 – 6.01)**
Diuretics	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.80 (0.60 – 1.08)
	Discontinuous–Insured	0.50 (0.29 – 0.87)*
Anti-HTN Combinations	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.98 (0.65 – 1.50)
	Discontinuous–Insured	1.60 (0.82 – 3.15)
Statins	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.85 (0.63 – 1.16)
	Discontinuous–Insured	1.02 (0.63 – 1.66)
Sulfonylureas	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.60 (0.38 – 0.94)*
	Discontinuous–Insured	0.63 (0.28 – 1.43)
Biguanides	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	1.02 (0.69 – 1.50)
	Discontinuous–Insured	1.30 (0.63 – 2.67)

Table 3-6. (Continued)

Study population included MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid anytime during the year. Estimates in bold indicate statistical significance at $P < 0.05$. * $P < 0.05$, ** $P < 0.01$, *** $P < 0.001$. MPR: Medication Possession Ratio. All models were adjusted for age, gender race/ethnicity, marital status, family size, census region, education, income, employment, MSA, SSI/TANF, Medicaid HMO/MCO, mental/substance abuse illnesses, arthritis/joint pain, no. of CVD conditions, perceived health status, perceived mental health status, Charlson Comorbidity Index score, and survey year.

Table 3-7. Association between Medicaid Discontinuity and Prescription Medication Utilization in the Study Population

Outcome	Medicaid Coverage	Incidence Rate Ratio (95% CI)	Predicted Difference in Events
Prescription Drug Fills			
All-Cause	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	0.73 (0.68 – 0.79)***	– 12.1 fills***
	Discontinuous–Insured	0.86 (0.76 – 0.97)*	– 6.5 fills**
Disease-Specific	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	0.86 (0.79 – 0.95)**	– 2.3 fills**
	Discontinuous–Insured	0.84 (0.73 – 0.96)*	– 2.8 fills**

Study population included MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid anytime during the year.

Two separate covariates-adjusted regression models for each outcome comparing Continuous to Discontinuous–Uninsured, and to Discontinuous–Insured, respectively. Estimates in bold indicate statistical significance at $P < 0.05$.

MPR: Medication Possession Ratio

All models were adjusted for age, gender race/ethnicity, marital status, family size, census region, education, income, employment, MSA, SSI/TANF, Medicaid HMO/MCO, mental/substance abuse illnesses, arthritis/joint pain, no. of CVD conditions, perceived health status, perceived mental health status, Charlson Comorbidity Index score, and survey year.

* $P < 0.05$, ** $P < 0.01$, *** $P < 0.001$ for comparison between the continuous and discontinuous Medicaid groups.

Medication utilization was similarly lower among the Discontinuous–Insured group. When examining all-cause medication utilization, this study found that the Discontinuous–Insured had approximately 14% lower prescription drug fills annually (IRR: 0.86; 95% CI: 0.76 – 0.97), with a predicted marginal difference of 6.5 fewer medication fills per person (SE: 2.4; $P=0.007$), compared to the Continuous Medicaid group. Similarly, the Discontinuous–Insured had an approximately 14% lower fills of overall prescription medications annually (IRR: 0.86; 95% CI: 0.74 – 0.99), and a predicted marginal difference of 2.4 fewer drug fills on average per person annually (SE: 1.1; $P=0.023$), when compared to the Continuous Medicaid coverage population.

DISCUSSION

Discontinuous Medicaid coverage among adults with CVD or high-risk conditions for CVD was found to have a negative association with prescription drug utilization. The effects were more pronounced among individuals with less than full-year Medicaid coverage and having no other insurance during the year. This Discontinuous–Uninsured group was predicted to have 12 fewer fills per person on average for all prescription medications, and about 2 fewer fills per person on average for prescription drugs specific to CVD or high-risk conditions for CVD, compared to the Continuous Medicaid group. Similarly, the Discontinuous–Insured group, or individuals with less than full-year coverage who reported having other sources of insurance coverage during the year, were found to have an average of about 6 fewer drug fills per person overall, and about 3 fewer fills per person for disease-specific medications. Adherence to commonly prescribed medications to manage CVD or high-risk conditions for CVD, measured as average MPR to determine overall concurrent medication adherence, was not found to be significantly different between the continuous and discontinuous Medicaid coverage groups in this study. However, when examining adherence specific to individual therapeutic drug classes, this study found lower adherence among the discontinuous Medicaid groups to sulfonyleureas and diuretics. Adherence to ACEI/ARB and calcium channel blockers was found to be higher among the Discontinuous–Insured groups, and was not significantly different between the Discontinuous–Uninsured and the Continuous Medicaid groups as well. Being among the first line therapies to treat high blood pressure that are commonly prescribed, access to these medication classes may not have been impacted by changes in Medicaid coverage during the year in these population groups.

Medication adherence is a highly complex phenomena and barriers to adherence are multifactorial. The World Health Organization (WHO) has done a phenomenal job in classifying the barriers of adherence that capture a wide range of factors that contribute to non-adherence.⁷³ As per the WHO classification, barriers to adherence result from: 1) factors associated with the health care team and system in place, 2) disease-related factors, 3) patient-related factors, 4) therapy-related factors, and 5) socio-economic factors. Health system- and health care team-related factors can include lack of access to health care and prescribed medications, and lack of continuity of care due to changes in insurance status, among others. Complex chronic debilitating conditions, such as heart

failure or stroke, or asymptomatic disorders such as hypertension and diabetes may prevent taking therapies as prescribed or seeking needed medical care. Patient factors such as younger age, minority race or ethnicity, presence of physical or cognitive impairments can also be barriers to medication adherence. Complexity of medication therapy regimen, and side effects of therapies previously undetected in clinical trials can also hinder appropriate management of conditions through pharmacotherapy.⁷⁴ Finally, socioeconomic factors such as low incomes, high medication costs, low literacy, and poor social support may hinder effective use of medications to manage disease conditions.⁶²

The present study examined the impact of a health system-related barrier of not having continuous full-year Medicaid coverage on medication adherence after accounting for several other factors that may hinder or confound medication adherence. Given that this study found Medicaid discontinuity to be significantly associated with lower prescription fills, we have strong reasons to believe that having discontinuous Medicaid coverage has a negative impact on prescription medication utilization, and by estimating adherence within a limited time frame of one year using an imputed MPR measure due to absence of complete days' supply information, this study may have underestimated the impact of Medicaid discontinuity on adherence. Adherence to certain classes of antidiabetic and antihypertensive medications was found to be lower among the discontinuous Medicaid groups which further highlights the negative impact that discontinuity and disruptions in Medicaid coverage may have on managing high-risk conditions for CVD.

To our knowledge, there have been no previous studies which have examined the association between discontinuous Medicaid coverage and prescription medication adherence, and the paucity of empirical evidence establishing an association between Medicaid discontinuity and prescription drug utilization, especially among subpopulations with chronic diseases, warranted an in-depth examination of these outcomes. Our findings on prescription drug fills are in line with the findings of Banerjee and colleagues, who examined the impact transitions into and/or out of Medicaid coverage on prescription drug fills, among other outcome measures.¹⁵ They found that individuals experiencing discontinuity in their Medicaid coverage had lower incidence of prescription drug refills compared to those without any disruptions in their Medicaid coverage. Among studies that investigated disease-specific prescription drug utilization, Smith and Kirking found that individuals with Medicaid coverage disruptions were associated with significantly lower acquisition of medications for HIV/AIDS (human immunodeficiency virus infection and acquired immune deficiency syndrome), such as antiretrovirals.⁶⁷ Fields and colleagues examined the impact of insurance stability and residence on several health care utilization measures and found that discontinuously insured individuals had fewer prescription drug fills, physician office visits, among other outcome measures, compared to those with continuous insurance coverage.³² Baicker and colleagues utilized the random assignment of individuals embedded in the Oregon Medicaid lottery to examine the impact of Medicaid coverage on several outcome measures, dubbed the Oregon Experiment. They found that although the use of diabetes medications increased 2 years after the Medicaid lottery to expand Medicaid coverage, the utilization of medications for hypertension or high cholesterol were not significantly

different.⁷⁵ Differences in sampling, study design (observational vs. random assignment), and duration, besides differences in the study sample may have contributed to the discrepancies between some of these findings and findings from the present study.

Medicaid, the nation's primary public health insurance program for the low-income population, provides much-needed health insurance coverage to an estimated 67 million individuals who meet the eligibility criteria, which include having incomes and assets below a certain threshold level determined by each State.⁷ Medicaid enrollees have very little copays, with most states requiring about \$1 – 3 per drug as copay, although most of them have limits of up to 5 concurrent prescription medications that can be filled per month.⁷⁶ Although seemingly restrictive, studies have found Medicaid enrollees to be more likely to be prescribed medications for certain chronic conditions. Rice and Colleagues in their analysis of the 2001 California Health Interview Survey found that California Medicaid enrollees had higher odds of reporting taking medications for heart disease, high blood pressure, and asthma.⁷⁷ Nonetheless, continuity of insurance coverage, especially for low-income populations, is as critical as gaining insurance coverage, and the findings from this study reinforce the conclusions made by previous studies that discontinuous Medicaid coverage may have an adverse impact on enrollees' prescription drug utilization.^{15,32,67}

The literature examining the impact of lack of insurance on health care outcomes in general, and disease-specific outcomes in particular is exhaustive with studies highlighting the adverse impact of lacking health insurance. However, there is a lack of studies examining the impact of discontinuous health insurance on medication outcomes, especially for certain disease conditions such as cardiovascular diseases that require prolonged disease management. And although medication adherence as an area of research has been thoroughly examined with several high-quality robust analyses of adherence and its impact on health care utilization, research examining the impact of churning in insurance in general and Medicaid on medication adherence in particular is limited.

Adherence to medications for CVD or high-risk conditions for CVD varies significantly depending on the study population, and therapeutic classes examined. Studies have shown adherence to several therapeutic classes of medications such as statins, beta blockers, and antiplatelet agents declined following hospitalizations for AMI, CAD, and stroke, and that the trend in decreased adherence continued over periods of more than a year after discharge.^{66,78-80} Using medication event monitoring data, Vrijens and colleagues found medication annual nonadherence rate of 50% among patients prescribed antihypertensive drugs, whereas Bramley and colleagues in their examination of association between adherence and blood pressure control found 75% of patients to be adherent to antihypertensives.^{81,82} Bailey and colleagues analyzed Tennessee's Medicaid claims data to examine association between antihypertensive medication adherence and ambulatory visits to hazards of stroke and mortality, and found the mean adherence rate to antihypertensive medications to be 67%.⁶⁵ Examinations of medication adherence among health failure patients have also found differing estimates of adherence.^{83,84} Adherence to sulfonylureas was found to be lower among the Discontinuous–Uninsured

compared to the Continuous Medicaid population. Estimates have found adherence to oral hypoglycemic agents to be low among Medicaid enrollees, and range from 35 to about 50%.^{85,86} Discontinuous Medicaid coverage with no other coverage during the year may exacerbate the already low levels of adherence to this therapeutic class, which may have been reflected in these findings. Although nonadherence to cardiovascular medications is prevalent, the varying methodologies to estimate adherence, data sources, and study designs, make it difficult to compare results across studies focused on cardiovascular conditions.

This study is not without limitations. Adherence was calculated using an imputed MPR measure that approximated days' supply for medications based on an algorithm created using quantity of medications dispensed. However, unlike previous empirical analyses, this study utilized MEPS data itself to impute MPR, using a plausible assumption that over the 10-year period from 2002, there have not been major changes in the strengths and dosage frequencies of cardiovascular medications. The quality of the prescription drug data in MEPS has been found to be comparable to Medicare claims data, and combined with the high degree of correlation between the imputed and actual days' supply variables, we believe benchmarking our analysis solely to MEPS data is a credible approach.⁸⁷ Additionally, each state establishes their own Medicaid eligibility and procedures for reenrollment, which were not available in MEPS to explore the complex and dynamic patterns of discontinuity that may differ based on the state's eligibility and reporting criteria. Further, this study only explored associations between individual characteristics and Medicaid discontinuity, and not causality, and the findings only explore predictors of within-year discontinuous Medicaid coverage and may not capture risk factors for secular trends in Medicaid discontinuity. Regression models may not have adjusted for certain covariates that may confound the relationship between Medicaid insurance and medication adherence and utilization, or account for unobserved differences between the study population groups. Finally, MEPS data are self-reported, so the estimates may not exactly reflect the national estimates of Medicaid discontinuity. Researchers have speculated that most surveys underestimate the number of Medicaid enrollees due to a variety of reasons including stigma of public insurance. However, such misreporting is less of a problem for MEPS compared to other federal health surveys.⁵⁴ Finally, findings from this study are representative of the non-institutionalized civilian population and may not be applicable to individuals with prolonged hospitalizations or admitted to long-term care facilities.

CONCLUSION

The findings from this study suggest that continuous Medicaid coverage is critical for non-elderly adults with CVD or high-risk conditions for CVD, and having discontinuous Medicaid coverage was associated with lower prescription drug utilization, and lower adherence to certain therapeutic classes of medications, which are essential components of appropriate disease management. Losing Medicaid coverage may increase the cost sharing burden for former enrollees who have a majority of their maintenance medications covered with nominal copays, resulting in forgoing needed pharmacotherapy, and exacerbating already low levels of adherence to medication

regimens. Further research with more complete prescription drug data to analyze the impact of Medicaid discontinuity on adherence to important therapeutic classes may better inform policymakers on the determinants of nonadherence, and optimize the delivery of interventions to improve adherence to these vulnerable populations. Discussions on policies to improve retention of and continuity in Medicaid must include individuals suffering from these chronic and debilitating conditions whose pharmacotherapy may be adversely impacted due to such coverage instability.

CHAPTER 4. ASSOCIATIONS BETWEEN MEDICAID DISCONTINUITY, ACCESS TO CARE, PREVENTIVE CARE, AND HEALTH SERVICES UTILIZATION AMONG ADULTS WITH CARDIOVASCULAR DISEASE OR ASSOCIATED HIGH-RISK CONDITIONS

BACKGROUND

Medicaid provides much-needed health insurance coverage to an estimated 67-70 million low-income and disabled children, adults, and elderly at any given point in time, who meet the eligibility criteria, which include having incomes and assets below a certain threshold level determined by each State.^{7,88} Recent estimates indicate that over 28% of non-elderly adult Medicaid enrollees with family incomes at or below 138 percent of the federal poverty level (FPL) are diagnosed with cardiovascular disease (CVD), including heart disease and stroke.²⁰ Conditions such as hypertension, hypercholesterolemia, and diabetes, among others, are significant risk factors for CVD; however, these conditions are preventable and manageable. For low-income vulnerable populations having CVD or high-risk conditions for CVD, accessing needed medical care, prescription drugs, routine preventive care, and periodic primary care visits are extremely critical to manage their complex chronic conditions. Failure to do so may exacerbate their cardiovascular symptoms leading to prolonged morbidity and result in avoidable hospitalizations that may exert a substantial financial stress on the health care system.

Overwhelming evidence suggests that access to medical care, routine preventive services, and having a usual source of care for primary care visits are critical in managing disease conditions, and maintaining good health.⁸⁹⁻⁹² Having stable, continuous health insurance coverage has been shown to be critical in ensuring that individuals do not have to forego medical care or have high financial burden, ensure appropriate preventive care services are accessed on a timely basis, and have regular access to primary care providers to obtain the medical care necessary for preventing and treating their chronic diseases.^{2,93-95} Poor disease management, resulting from inadequate access to care, lack of adherence to pharmacotherapy, or lack of sufficient access to primary care physicians may result in increased risk of hospitalizations that could have been avoided.^{91,96-98}

Low family incomes, assessed as a percentage of the federal poverty level, family status, residence, and assets test, are some of the criteria used to determine eligibility for Medicaid coverage, which are specific to each states. Although most states have moved towards a 12 month coverage renewal for non-elderly adults in recent years, they may still require certain individuals to periodically report their income and/or other criteria in order to maintain eligibility. Upward income mobility would render individuals ineligible for their state's Medicaid coverage if their incomes cross the eligibility threshold; however, research has shown that individuals may passively end up not reenrolling into Medicaid due to a variety of reasons, ranging from lack of awareness to various barriers faced in reenrollment.⁹⁹ This results in individuals and families losing Medicaid coverage within 12 months of enrollment. Instability in Medicaid, characterized by dropouts and moving back and forth into various sources of coverage or a state of being uninsured, has

been a long-standing issue that has impacted non-elderly adult Medicaid enrollees due to stringent administrative regulations governing coverage for this population.^{9,100} Previous estimates have shown that about 43% of individuals who are newly enrolled in Medicaid, end up losing their Medicaid coverage within twelve months.³⁸ Transitions into and out of Medicaid coverage may adversely impact the ability for low-income populations with CVD or high-risk conditions for CVD to seek, and maintain continuity in, medical care. Losing Medicaid coverage to become uninsured would be detrimental, as there is strong consensus regarding the negative outcomes associated with lack of insurance.^{10-12,31,32} However, transitioning between Medicaid and other sources of insurance coverage may also prove to be problematic, as this may lead to moving between different provider networks, and coverage with different benefit structures that may limit use of previously covered health care resources.

Estimates of the impact of discontinuous Medicaid coverage on access to care and utilization of health care services among individuals with CVD or high-risk conditions for CVD are very few and dated. Findings from state-specific Medicaid data have found Medicaid coverage discontinuity to be associated with higher inpatient and ER hospitalizations associated, especially for individuals with heart disease, diabetes, among other ambulatory care sensitive conditions.^{14,101} Newer, more comprehensive national estimates of the impact of discontinuous Medicaid coverage on access to care, preventive services use, and health care services utilization among individuals with CVD or its risk factor chronic conditions are critical to determine the outcomes among vulnerable diseased populations for whom insurance coverage continuity may be even more critical. The present study aims to obtain national estimates among adults with CVD or high-risk conditions for CVD of associations between Medicaid discontinuity and (1) access to care, (2) preventive services use, and (3) utilization of overall (all-cause) and disease-specific health care services, such as hospital inpatient visits, emergency room (ER) visits, hospital outpatient visits, and physician office visits. We hypothesized that having discontinuous Medicaid coverage, relative to full-year continuous Medicaid coverage, would be associated with significantly poor access to care, lower preventive services use, lower primary care utilization, and higher hospitalizations in this study population.

METHODS

The present study is a retrospective cross-sectional analysis comparing adult populations with CVD or high-risk conditions for CVD, with and without continuous full-year Medicaid coverage. Instability and disruptions in Medicaid coverage may especially have a detrimental impact on the continuity of care among low-income adults with CVD or any of its risk factor conditions. We hypothesized that Medicaid discontinuity would be associated with poor access to care, decreased primary care – characterized by lower preventive services use and physician visits – and increased hospitalizations, both overall (all-cause), and those specific to CVD or high-risk conditions for CVD (disease-specific).

Data Source

The present study analyzed data from the Medical Expenditure Panel Survey (MEPS). MEPS is a set of large-scale surveys of families, individuals, their medical providers and employers across the United States. It is jointly sponsored by the Agency for Healthcare Research and Quality (AHRQ) and the National Center for Health Statistics (NCHS), and has been conducted annually since 1996. It has two major components; the Household Component (HC), the Insurance Component, with data on medical providers in the HC being supplemented by the Medical Provider Component. MEPS collects detailed information from a nationally representative sample of the civilian noninstitutionalized population of the U.S. on health services utilization and health expenditures, insurance coverage, and sources of payment.

MEPS has an overlapping panel design. Two separate panels of respondents are interviewed simultaneously during a calendar year. A new panel of sample households is selected each year, and data for each panel are collected for two calendar years. The two years of data for each panel are collected in five rounds of interviews that take place over a two and a half year period. This provides continuous and current estimates of health care expenditures at both the person and household level for two panels for each calendar year. By combining data from the overlapping panels, these annual files provide nearly double the sample size of individual panels and cover the entire calendar year for each respondent. Although all MEPS data are reported by respondents during computer-assisted personal interviewing, further detailed health service use data, including on prescription drugs, are collected from a sample of providers with respondents' permission.³⁹

When obtaining national estimates from surveys such as MEPS, appropriate sample size is critical to obtain reliable estimates, since estimates of some population subgroups may vary from year to year. The advantage of survey such as MEPS is that they allow for several years' of data to be pooled together in order to increase the sample size and improve the precision of estimates. The pooling of the different years of data was conducted in accordance with the guidelines provided by MEPS, and data for the years 2002 – 2011 were pooled together to allow for a sufficiently large sample size to ensure reliable estimates.

Study Population

The broad classification of “cardiovascular disease (CVD) and associated comorbidities” in this study consisted for the following disease conditions: acute myocardial infarction (Clinical Classification code (CCC): 100), coronary artery disease (CCC: 101), congestive heart failure (CCC: 108), peripheral and visceral atherosclerosis (CCCs: 114, 115, 116), stroke (CCCs: 109, 112), hypertension (CCC: 098), lipid disorders (CCC: 053), and diabetes (CCC: 049, 050). CCC aggregates conditions and procedures into mutually exclusive and clinically homogeneous categories, using Clinical Classification Software.⁴⁰ The study population included respondents who were: (1) aged

18 to 64 years, (2) reported either being diagnosed with CVD or high-risk conditions for CVD for CVD in their first round of the MEPS survey, or reported having an event (an inpatient, emergency, outpatient, or office-based provider visits, or a prescription medication) associated with one or more of the aforementioned disease conditions at any point of time during the survey, (3) reported having had Medicaid coverage at any point of time during the survey year, and (4) had positive person weights, and were interviewed for all five rounds of the 2-year MEPS survey interviews. Respondents who reported having Medicare, or pregnancy were excluded from this study.

Respondents who reported having Medicare, or pregnancy were excluded from this study. Pregnant women with incomes at or below 133 percent of the federal poverty level are eligible for Medicaid coverage throughout the pregnancy; however, the coverage ceases 60 days postpartum.⁶⁸ Low-income non-elderly adults with disabilities are also eligible for Medicare if they have received monthly social security disability income for two years. Most coverage benefits for prescription drugs for dual eligible beneficiaries are covered by Medicare Part D, instead of Medicaid. As a result, these two population groups were excluded from this study.

Individuals with less than a year of Medicaid coverage who experience within-year Medicaid discontinuity may not be a homogenous group, especially when they report having other forms of insurance coverage during the year. In order to account for these differences, individuals in the study population who reported having <12 months of Medicaid coverage during a survey year were divided into two groups: (1) individuals with <12 months of Medicaid coverage who did not report any other sources of health insurance during the survey year, designated in this study as ‘Discontinuous–Uninsured’, and (2) individuals with <12 months of Medicaid coverage who also reported having other source(s) of health insurance during the survey year (either private, Tricare, or any hospital- or physician group-based health insurance), designated in this study as ‘Discontinuous–Insured’. Individuals in the study population who reported having Medicaid coverage during all 12 months of a survey year were classified as having continuous Medicaid coverage, which was the comparison group and designated as ‘Continuous Medicaid’ in this study.

Study Variables

Our first objective was to examine whether Medicaid discontinuity was associated with poor access to care among non-elderly adults with CVD or high-risk conditions for CVD. The outcomes of interest for this objective were the following three binary indicators for: (1) having a usual source of care, (2) having difficulty in obtaining medical care, and (3) having difficulty obtaining prescription medications. When examining association between Medicaid discontinuity and preventive services use, the following three binary outcomes were examined: (1) having blood pressure checkup within the past year, (2) having cholesterol checkup within the past year, and (3) having routine medical checkup within the past year. For the final objective of examining associations between Medicaid discontinuity, primary care visits, and hospitalizations, we

divided the utilization outcome measures into all-cause, i.e. overall utilization during the survey year irrespective of diagnosis, and disease-specific utilization, i.e. utilization specific to CVD or high-risk conditions for CVD considered in this study. Specifically, the total number of inpatient visits, ER visits, hospital outpatient visits, and physician office visits, were examined, resulting in eight outcome measures (one all-cause, and one disease-specific for each) that were characterized as discrete counts of the events

All regression models included the following classes of covariates:-

- 1) Demographic covariates: age (in years), gender (female vs. male), race/ethnicity (Non-Hispanic White vs. Non-Hispanic Black, Hispanic, and other races), marital status (not married vs. married), family size (N), and geographic residence (Northeast vs. Midwest, South, and West).
- 2) Socioeconomic covariates: education (less than high school vs. high school/GED, and Bachelor's or higher), income as a percentage of the federal poverty level (FPL) (<100% FPL vs. 100 - <125% FPL, and $\geq 125\%$ FPL), employment (always unemployed vs. intermittent employment, and always employed), binary indicator for Metropolitan Statistical Area (MSA) residence.
- 3) Medicaid-related covariates: binary indicator for Supplementary Security Income (SSI) for disability receipt or participation in the Temporary Assistance for Needy Families program, and a binary indicator for Medicaid HMO/Managed Care participation.
- 4) Health-related covariates: perceived health status (excellent/very good/good vs. fair/poor), perceived mental health status (excellent/very good/good vs. fair/poor) binary indicator for mental illness/substance abuse diagnosis, binary indicator for respiratory diseases diagnosis, binary indicator for arthritis/joint pain diagnosis, and Charlson Comorbidity Index (CCI) score.

Additionally, all of the multivariate models included the year fixed effects to capture any macro-level variations during the study period.

Statistical Analysis

Bivariate statistics were conducted to examine the characteristics of the study population. When examining the binary outcomes for access to care and preventive services use, adjusted multivariate logistic regression analyses were conducted. As an example, **Equation 4-1** and **Equation 4-2** depict the logistic regression models usual source of care, as part of the access to care measures, among both the Discontinuous–Uninsured, and the Discontinuous–Insured groups relative to the Continuous Medicaid insurance group.

$$Pr[USC=1|x] = F^{-1}(\hat{\alpha} + \hat{\beta}_1 \cdot Discontinuous - Uninsured + \sum_{k=1}^K \hat{\beta}_k X_k) \quad (\text{Eq. 4-1})$$

$$Pr[USC=1|x] = F^{-1}(\hat{\alpha} + \hat{\beta}_1 \cdot Discontinuous - Insured + \sum_{k=1}^K \hat{\beta}_k X_k) \quad (\text{Eq. 4-2})$$

Usual source of care is modeled as a dichotomous outcome variable as a function of Medicaid discontinuity as well as the demographic, socioeconomic, Medicaid-related, and health-related covariates mentioned above. F^{-1} denotes the inverse of the cumulative standard logistic distribution function that relates the outcome on the probability scale to the covariates. $Pr[\]$ denotes the population average probability of having a usual source of care (denoted by 1) that is conditional on the covariates denoted by x . *Discontinuous–Uninsured* and *Discontinuous–Insured* are binary indicator variables for the main independent variable of interest, that equals one for individuals having discontinuous annual Medicaid coverage without any other sources of insurance and with other sources of insurance during the year, respectively, and equals zero for the comparison group of individuals with continuous Medicaid coverage. X denotes the vector of the aforementioned individual characteristics that predict Medicaid discontinuity in the two population groups of interest.

The number of physician office visits data (both all-cause and disease-specific) displayed overdispersion, however, this count outcome measure did not have excess zeros. Thus, an adjusted negative binomial regression was utilized. It is a type of generalized linear model where the dependent variable Y is a count of the number of times an event occurs (in this case, the number of physician office visits). A convenient parameterization of the negative binomial distribution is denoted by **Equation 4-3**.⁷¹

$$p(y) = P(Y = y) = \frac{\Gamma(y + \frac{1}{\alpha})}{\Gamma(y+1)\Gamma(\frac{1}{\alpha})} \left(\frac{1}{1+\alpha\mu}\right)^{1/\alpha} \left(\frac{\alpha\mu}{1+\alpha\mu}\right)^y \quad (\text{Eq. 4-3})$$

In **Equation 4-3**, $\mu > 0$ is the mean of Y , which is the count variable indicating the number of prescription drug fills (either all-cause, or disease-specific) and $\alpha > 0$ is the heterogeneity parameter. It can be derived as a Poisson-gamma mixture, or as the number of failures before the $(1/\alpha)^{\text{th}}$ success. A simpler depiction of the traditional negative binomial regression model, the NB2 model is depicted is depicted in **Equation 4-4** and **Equation 4-5** as follows,

$$\ln \mu_{\text{all-cause}} = \beta_0 + \beta_1 \text{Discontinuous} - \text{Uninsured} + \beta_2 \cdot x_2 + \cdots + \beta_k \cdot x_k \quad (\text{Eq. 4-4})$$

$$\ln \mu_{\text{disease-specific}} = \beta_0 + \beta_1 \text{Discontinuous} - \text{Uninsured} + \beta_2 \cdot x_2 + \cdots + \beta_k \cdot x_k \quad (\text{Eq. 4-5})$$

where, apart from the main indicator variables for the Discontinuous–Uninsured group, the covariates x_2 to x_k are known, and the population regression coefficients β_0 to β_k are to be estimated. The two equations above depict the negative binomial regression models to estimate the differences in overall (**Equation 4-4**), and disease-specific physician office visits (**Equation 4-5**) between the Continuous Medicaid, and Discontinuous–Uninsured groups. Similar regression models (not shown here) were run to estimate differences in the number of physician office visits between the Continuous Medicaid and Discontinuous–Insured groups.

When analyzing the discrete count data for the number of all-cause and disease-specific inpatient visits, ER visits, and hospital outpatient visits, it was important to account for the overdispersion in the data characterized by the presence of an excessive number of zeros, a problem that is commonly known as zero inflation (ZI). **Appendix B, Figures B-1 to B-16**, display the distribution of the all-cause and disease-specific outcome variables examined in this study. The distribution of the ZI data is often right-skewed, with a huge peak at zero, representing the individuals with zero health care consumption, and a skewed tail representing the rest of the study population with varying amounts of non-zero consumption. ZI data are often analyzed using a two-part mixture models combining a point mass at zero with a proper count distribution. These models assume an initial process to determine membership into one of two latent groups, generally referred to the ‘susceptible’ and ‘non-susceptible’ population groups.¹⁰² The approach for ZI models uses two regression models, one is usually a logit regression (or probit) modeling the susceptible probability, and the other is either a Poisson or negative binomial model, modeling the mean for the susceptible population (**Equation 4-6**). Explanatory variables are allowed to have a different impact for the two processes, and the two models are fit simultaneously using maximum likelihood estimation.¹⁰² The outcomes measures with excess zeros were therefore analyzed using zero inflated negative binomial regression models, mathematically represented as follows:

$$P(Y = y) \begin{cases} p + (1 - p) \left(\frac{1}{1 + \alpha\mu} \right)^{\frac{1}{\alpha}}, & y = 0 \\ (1 - p) \frac{\Gamma(y + \frac{1}{\alpha})}{\Gamma(y + 1) \Gamma(\frac{1}{\alpha})} \left(\frac{1}{1 + \alpha\mu} \right)^{\frac{1}{\alpha}} \left(\frac{\alpha\mu}{1 + \alpha\mu} \right)^y, & y = 1, 2, \dots \end{cases} \quad (\text{Eq. 4-6})$$

Having a mean $E(Y) = (1 - p)\mu$, and a variance $Var(Y) = (1 - p)\mu(1 + p\mu + \alpha\mu)$. Research has shown zero-inflated negative binomial models to provide superior fit when compared to zero-inflated Poisson regression.¹⁰³ For the binary model estimating the susceptible probability, the following variables were included

The coefficient of interest for the Discontinuous–Uninsured is the measure of differences in the utilization of overall, and disease-specific drug utilization. Because of the functional form of the negative binomial regression, the effect size of a factor is the antilog of its coefficient. This estimated coefficients are transformed to obtain the incidence rate ratio (IRR). An IRR of less than one for the discontinuous Medicaid groups would suggest lower hospital outpatient utilization for instance, among them when compared to their counterparts with continuous Medicaid coverage.

In addition to obtaining IRRs, this study further estimated the average marginal effect to determine the predicted difference in the number of all-cause and disease-specific prescription drug fills between the continuous and discontinuous Medicaid groups. The marginal effects measure the discrete change in the outcome variable, i.e. for a binary independent variable, how do predicted probabilities change as the binary independent variable changes from 0 (Continuous Medicaid) to 1 (Discontinuous–Uninsured, or Discontinuous–Insured).⁷² In order to obtain these marginal effects, after developing and running the multivariate negative binomial regression model, the

estimated coefficients of the model were used to calculate two predicted event visits for all individuals in the study population. The first predicted utilization assumed all individuals had discontinuous Medicaid coverage (by setting the Medicaid discontinuity indicator variable to 1), and the second predicted utilization assumed all individual to have continuous full-year Medicaid coverage (by setting the Medicaid discontinuity indicator variable to 0). The average per-person difference in the event visits attributable to Medicaid coverage was estimated by obtaining the difference in predicted drug utilization for each person and computing the weighted average of the difference across the entire study sample.

The proportion of missing data was around 6% in this study, as a result, the listwise deletion approach was used to account for missing values. Multicollinearity was assessed for all the regression models, and the variance inflation factor was found to be less than 5, which was lower than the widely accepted threshold of 10 for existence of multicollinearity. The Hosmer-Lemeshow tests for goodness-of-fit were highly insignificant for all the logistic regression models indicating that the model fit the data. All data analyses were conducted using SAS 9.3 and STATA 13 to account for the complex survey design of MEPS.

RESULTS

Characteristics of the Study Populations

Table 4-1 displays the characteristics of the study population. The study sample included 3,926 respondents in the Continuous Medicaid coverage group (nationally representative of about 2.4 million individuals), 1,407 individuals in the Discontinuous–Uninsured group (nationally representative of over 900,000 individuals), and 493 individuals in the Discontinuous–Insured group (nationally representative of over 400,000 individuals).

When compared to the Continuous Medicaid group, individuals in the Discontinuous–Uninsured group were less likely to belong to other race groups ($P<0.001$), more likely to be married ($P<0.001$), and were more likely to reside in the South region ($P=0.049$). Among the socioeconomic covariates, this study found that respondents were more likely to have completed high school/GED or college ($P=0.001$), were more likely to have family incomes $\geq 125\%$ FPL ($P<0.001$), and more likely to have intermittent, as well as continuous full-year employment ($P<0.001$). Respondents were less likely to be eligible for SSI due to disability or participate in TANF ($P<0.001$), and less likely to report being Medicaid managed care enrollees ($P<0.001$). Among the health-related variables, respondents were less likely to report having fair or poor perceived mental health status ($P<0.001$), less likely to report having mental/substance abuse illnesses ($P=0.003$), less likely to report having respiratory illnesses ($P<0.001$), and less likely to report having arthritis/joint pain ($P=0.009$). The mean CCI score was not significantly different between the two groups ($P=0.181$).

Table 4-1. Characteristics of the Study Population (2002 – 2011)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
DEMOGRAPHIC								
Age					0.385			0.013
18 – 34 years	388,425	15.9 (1.0)	165,340	18.1 (1.3)		98,188	23.1 (2.3)	
35 – 49 years	875,809	35.9 (1.2)	324,653	35.5 (1.8)		141,109	33.3 (2.7)	
50 – 64 years	1,173,144	48.2 (1.3)	423,933	46.4 (2.0)		184,968	43.6 (3.0)	
Gender					0.491			0.169
Female	1,632,006	67.0 (1.2)	624,410	68.3 (1.8)		267,502	63.1 (2.8)	
Race/Ethnicity					<0.001			<0.001
Non-Hispanic White	1,094,002	44.9 (1.8)	442,534	48.4 (2.1)		256,231	60.4 (3.2)	
Non-Hispanic Black	666,358	27.3 (1.5)	219,033	24.0 (1.6)		79,936	18.8 (2.0)	
Hispanic	445,601	18.3 (1.2)	202,106	22.1 (1.5)		50,704	12.0 (1.6)	
Others	231,416	9.5 (0.9)	50,253	5.5 (0.9)		37,393	8.8 (1.7)	
Marital Status					<0.001			<0.001
Married	727,457	29.9 (1.3)	344,150	37.7 (1.8)		184,414	43.5 (3.2)	
Region					0.049			0.023
Northeast	608,070	25.0 (1.6)	187,785	20.5 (1.7)		104,394	24.6 (2.9)	
Midwest	477,401	19.6 (1.4)	184,710	20.2 (1.7)		115,146	27.2 (3.2)	
South	727,520	29.8 (1.6)	315,073	34.5 (1.8)		98,072	23.1 (2.3)	
West	624,386	25.6 (1.7)	226,358	24.8 (1.9)		106,652	25.1 (3.0)	
Family Size Mean (SE)	2.8 (0.1)		2.9 (0.1)		0.409	2.8 (0.1)		0.976

Table 4-1. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
SOCIOECONOMIC								
Education					0.001			<0.001
<High School	971,041	40.2 (1.4)	298,300	32.9 (1.8)		67,896	16.0 (1.9)	
High School/GED	1,215,968	50.4 (1.4)	498,289	54.9 (2.0)		232,334	54.9 (3.2)	
College Degree	227,668	9.4 (0.7)	111,300	12.2 (1.4)		123,109	29.1 (3.2)	
Income					<0.001			<0.001
<100% FPL	1,384,551	56.8 (1.2)	428,540	46.9 (1.7)		104,719	24.7 (2.3)	
100 – <125% FPL	227,591	9.3 (0.6)	91,940	10.1 (1.1)		28,783	6.8 (1.3)	
≥125% FPL	825,236	33.9 (1.1)	393,445	43.0 (1.8)		290,763	68.5 (2.6)	
Employment					<0.001			<0.001
Always Unemployed	1,767,234	72.5 (1.2)	496,282	54.3 (1.8)		112,490	26.5 (2.2)	
Unstable Employment	301,597	12.4 (0.8)	228,246	25.0 (1.6)		129,583	30.6 (2.5)	
Always Employed	368,547	15.1 (0.9)	189,398	20.7 (1.4)		182,191	42.9 (2.9)	
MSA	1,931,797	79.3 (1.8)	721,532	78.9 (1.9)	0.880	344,287	81.1 (2.7)	0.494
MEDICAID-RELATED								
SSI/TANF	1,012,240	41.5 (1.3)	154,157	16.9 (1.3)	<0.001	41,212	9.7 (1.6)	<0.001
Medicaid HMO/MCO	1,424,726	61.3 (1.5)	459,330	50.5 (2.0)	<0.001	233,429	55.0 (3.0)	0.048

Table 4-1. (Continued)

Characteristics	Continuous Medicaid		Discontinuous–Uninsured			Discontinuous–Insured		
	Weighted N	% (SE)	Weighted N	% (SE)	<i>P</i> -value [†]	Weighted N	% (SE)	<i>P</i> -value ^{††}
HEALTH-RELATED								
Perceived Health Status					0.162			<0.001
Fair to Poor	1,314,794	53.9 (1.2)	467,555	51.2 (1.9)		170,598	40.2 (2.5)	
Perceived Mental Health					<0.001			<0.001
Fair to Poor	821,318	33.7 (1.2)	234,931	25.7 (1.7)		84,716	20.0 (2.2)	
CVD or high-risk conditions for CVD					0.775			0.037
≤1	1,678,222	68.9 (1.1)	640,242	70.1 (1.8)		321,685	75.8 (2.3)	
2	444,995	18.2 (0.8)	163,062	17.8 (1.4)		65,581	15.5 (2.2)	
≥3	314,160	12.9 (0.8)	110,621	12.1 (1.3)		37,000	8.7 (1.5)	
Mental/Substance Abuse Illnesses	751,357	30.8 (1.2)	227,885	24.9 (1.6)	0.003	115,891	27.3 (2.4)	0.176
Respiratory Illnesses	479,702	19.7 (1.1)	117,992	12.9 (1.1)	<0.001	39,380	9.3 (1.5)	<0.001
Arthritis/Joint Pain	830,408	34.1 (1.3)	267,358	29.3 (1.7)	0.009	107,575	25.4 (2.4)	0.002
CCI Score Mean (SE)	1.2 (0.0)		1.1 (0.1)		0.079	0.9 (0.1)		0.003
Population Size	2,437,378	64.6 (1.0)	913,926	24.2 (0.8)		424,265	11.2 (0.7)	

Study population includes MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, had ≤1 month of Medicaid coverage during the survey year. [†]*P*-value for the difference between Discontinuous–Uninsured and Continuous Medicaid groups. ^{††}*P*-value for the difference between Discontinuous–Insured and Continuous Medicaid groups. *P*-values in **bold** indicate statistically significant difference at *P*<0.05. SE: Standard error; GED: General educational development; FPL: Federal poverty level; MSA: Metropolitan Statistical Area; SSI: Supplementary Security Income.

When comparing the Discontinuous–Insured group to the Continuous Medicaid group in the above table, this study found that among the demographic covariates, individuals in the Discontinuous–Insured group were more likely to be younger ($P=0.013$), were less likely to belong to minority race/ethnicity groups ($P<0.001$), more likely to be married ($P<0.001$), and were more likely to reside in the Midwest ($P=0.023$). Among the socioeconomic covariates, this study found that individuals in the Discontinuous–Insured group were more likely to complete high school/GED or college ($P<0.001$), more likely to have higher family income levels ($P<0.001$), and were more likely to have some form of employment during the year ($P<0.001$). The Discontinuous–Insured group was also less likely to be eligible for SSI due to disability or participate in TANF ($P<0.001$), and less likely to report having Medicaid through an HMO/MCO ($P=0.048$). Among the health-related covariates, this study found that the Discontinuous–Insured group was less likely to report having self-perceived health status ($P<0.001$), and self-perceived mental health status ($P<0.001$), more likely to report fewer CVD or high-risk conditions for CVD ($P=0.037$), were less likely to report having respiratory illnesses ($P<0.001$), less likely to report having arthritis/joint pain ($P=0.002$), and lower chronic disease burden ($P=0.003$).

Medicaid Enrollment among the Discontinuous Medicaid Coverage Groups

Table 4-2 shows the summary statistics for Medicaid enrollment among the discontinuous Medicaid insurance groups. The average duration of enrollment for the undivided discontinuous Medicaid population was 6.0 months. The Discontinuous–Uninsured group was enrolled in Medicaid for a longer period on average compared to the Discontinuous–Insured group (6.5 vs. 5.0 months, $P<0.001$). Additionally, this study examined the average duration of no insurance and found individuals in the Discontinuous–Insured group were uninsured for a little over a month (5.5 vs. 1.2 months, $P<0.001$). This study also examined Medicaid enrollment status after 12 months of initial Medicaid enrollment. Overall, a total of 4,898 individuals were enrolled in the month of January of the survey year, of which 85.3% were still enrolled in Medicaid during the month of December in the survey year, 3.5% had some other form of insurance coverage, and 10.2% were uninsured (results not shown). Of the 621 individuals in the Discontinuous–Uninsured group who had Medicaid in the beginning of the year, 78.1% were uninsured while 21.9% still had Medicaid coverage at the end of the year. Of the 200 individuals in the Discontinuous–Insured group who had Medicaid in the beginning of the survey year, only 8.0% still had Medicaid, 6.0% were uninsured, and the remaining 86% had some other source of health insurance coverage at the end of the year. Finally, when examining the number of transitions into and/or out of Medicaid coverage, this study found that overall 17.6% of individuals in this study population had more than one transitions, and the proportion of multiple transitions was similar between the two discontinuous Medicaid subgroups (17.7% vs. 17.3%, $P=0.875$).

Table 4-2. Patterns of Medicaid Coverage among Discontinuous Medicaid Groups

Coverage Components	Discontinuous–Uninsured		Discontinuous–Insured	
	Mean or Weighted N	SE or % (SE)	Mean or Weighted N	SE or % (SE)
Mean Duration of Medicaid	6.5 months	0.1	5.0 months***	0.2
Mean Duration of No Insurance	5.5 months	0.1	1.2 months***	0.1
Insurance Status After 12 Months[†]				
Still in Medicaid	83,125	21.9 (2.2)	11,647***	8.0 (2.0)
Uninsured	297,045	78.1 (2.2)	10,058***	6.0 (2.0)
Other Insurance	--	--	147,896	86.0 (2.0)
Medicaid Coverage Transitions				
1	752,311	82.3 (1.4)	350,803	82.7 (2.0)
≥2	161,615	17.7 (1.4)	73,462	17.3 (2.0)

Sample consisted of 380,171 individuals in the Discontinuous–Uninsured group and 169,600 individuals in the Discontinuous–Insured group who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, hospital outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid at the beginning of the year.

*** $P < 0.001$; ** $P < 0.01$; * $P < 0.05$ for difference from the Discontinuous–Uninsured group.

Access to Care

In comparison to the Continuous Medicaid group, individuals in the Discontinuous–Uninsured group were less likely to report having a usual source of care (93.6% vs. 84.5%; $P<0.001$), and more likely to report having difficulties in obtaining medical care (5.6% vs. 19.1%; $P<0.001$), and prescription medications (5.8% vs. 15.8%; $P<0.001$). When compared to the Continuous Medicaid group, those in the Discontinuous–Insured group were less likely to report having a usual source of care (93.6% vs. 88.6%; $P=0.002$), and more likely to report having difficulty in obtaining prescription medications (5.8% vs. 11.3%; $P<0.001$); however, there were no differences between the two groups in facing difficulty in obtaining medical care ($P=0.094$). Results of these bivariate analyses are not displayed. **Table 4-3** displays the results of the multivariate logistic regression examining association between Medicaid discontinuity and access to care, among adults with CVD or high-risk conditions for CVD. In comparison to the Continuous Medicaid group, the Discontinuous–Uninsured group were found to have almost 60% lower odds of having a usual source of care [Odds Ratio (OR): 0.41; 95% Confidence Interval (CI): 0.32 – 0.52]. Additionally, the odds were a staggering 4 times higher of reporting having difficulty in obtaining medical care (OR: 4.20; 95% CI: 3.14 – 5.62), and were more than 3 times higher of reporting having difficulty in obtaining prescription medications (OR: 3.32; 95% CI: 2.50 – 4.40) when compared to those with continuous Medicaid coverage. Among the Discontinuous–Insured group, the odds of having a usual source of care were 43% lower when compared to the Continuous Medicaid group (0.57; 95% CI: 0.37 – 0.88). The odds of reporting having difficulty in obtaining medical care were 74% higher (OR: 1.74; 95% CI: 1.02 – 2.98), and the odds of reporting having difficulty in obtaining prescription medications were more than twice, when compared to the Continuous Medicaid group (OR: 2.16; 95% CI: 1.42 – 3.27).

Preventive Services Use

When examining the bivariate statistics between the Continuous Medicaid and Discontinuous–Uninsured groups, this study found no significant differences between the two when reporting about the routine BP checkup ($P=0.073$) and routine cholesterol checkup ($P=0.140$); however, the Discontinuous–Uninsured group was less likely to report having had a medical checkup within the past year ($P=0.004$). When comparing the Continuous Medicaid and Discontinuous–Insured groups, this study found that those in the Discontinuous–Insured group were less likely to report having routine BP checkup (96.9% vs. 94.5%; $P=0.023$), whereas differences in routine cholesterol checkup ($P=0.897$) and medical checkup ($P=0.146$) were insignificant (results not displayed). **Table 4-4** displays the results of the multivariate logistic regression examining association between Medicaid discontinuity and preventive services use, among adults with CVD or high-risk conditions for CVD. When examining routine blood pressure checkup and routine blood cholesterol checkup, there were no significant differences between the continuous and discontinuous Medicaid populations. However, the Discontinuous–Uninsured were found to have 22% lower odds of having had a medical checkup within

Table 4-3. Association between Medicaid Discontinuity and Access to Care in the Study Population

Outcome	Medicaid Coverage	Odds Ratio (95% Confidence Interval)
Have Usual Source of Care	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.41 (0.32 – 0.52)***
	Discontinuous–Insured	0.57 (0.37 – 0.88)*
Difficulty in Obtaining Necessary Medical Care	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	4.20 (3.14 – 5.62)***
	Discontinuous–Insured	1.74 (1.02 – 2.98)*
Difficulty in Obtaining Prescription Medications	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	3.32 (2.50 – 4.40)***
	Discontinuous–Insured	2.16 (1.42 – 3.27)***

Study population included MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid anytime during the year.

Two separate covariates-adjusted regression models for each outcome comparing Continuous to Discontinuous–Uninsured, and to Discontinuous–Insured, respectively. Estimates in bold indicate statistical significance at $P < 0.05$. * $P < 0.05$, ** $P < 0.01$, *** $P < 0.001$ for comparison between the continuous and discontinuous Medicaid groups.

All models were adjusted for age, gender race/ethnicity, marital status, family size, census region, education, income, employment, MSA, SSI/TANF, Medicaid HMO/MCO, mental/substance abuse illnesses, arthritis/joint pain, no. of CVD conditions, perceived health status, perceived mental health status, Charlson Comorbidity Index score, and survey year.

Table 4-4. Associations between Medicaid Discontinuity and Preventive Services Use in the Study Population

Outcome	Medicaid Coverage	Odds Ratio (95% Confidence Interval)
Routine BP Checkup	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.90 (0.61 – 1.31)
	Discontinuous–Insured	0.75 (0.42 – 1.33)
Routine Cholesterol Checkup	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.85 (0.64 – 1.13)
	Discontinuous–Insured	1.06 (0.68 – 1.64)
Routine Medical Checkup	Continuous	1.00 (Reference)
	Discontinuous–Uninsured	0.78 (0.63 – 0.96)*
	Discontinuous–Insured	0.80 (0.56 – 1.13)

Study population included MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid anytime during the year.

Two separate covariates-adjusted regression models for each outcome comparing Continuous to Discontinuous–Uninsured, and to Discontinuous–Insured, respectively. Estimates in bold indicate statistical significance at $P < 0.05$. * $P < 0.05$, ** $P < 0.01$, *** $P < 0.001$ for comparison between the continuous and discontinuous Medicaid groups.

All models were adjusted for age, gender race/ethnicity, marital status, family size, census region, education, income, employment, MSA, SSI/TANF, Medicaid HMO/MCO, mental/substance abuse illnesses, arthritis/joint pain, no. of CVD conditions, perceived health status, perceived mental health status, Charlson Comorbidity Index score, and survey year.

the past year when compared to the Continuous Medicaid population (OR: 0.78; 95% CI: 0.63 – 0.96).

Utilization of Health Care Services

The results of the final objective of examining associations between Medicaid discontinuity and health care services utilization are displayed in **Table 4-5**. Overall inpatient visits were not significantly different between the continuous and discontinuous Medicaid groups. However, when examining disease-specific inpatient visits, this study found that the Discontinuous–Uninsured group had about 62% higher inpatient visits compared to the Continuous Medicaid group [Incidence Rate Ratio (IRR): 1.62; 95% CI: 1.19 – 2.22]. The marginal predicted difference in the number of inpatient visits was found to be 0.05 visits ($P=0.007$); in other words, individuals in the Discontinuous–Uninsured group were predicted to have 0.05 higher inpatient visits per person annually, compared to the Continuous Medicaid group. Similarly, the incidence of disease-specific inpatient visits was found to be 95% higher among the Discontinuous–Insured population (IRR: 1.95; 95% CI: 1.23 – 3.10), which at the margins was predicted to be 0.08 higher inpatient visits per person ($P=0.029$) when compared to the Continuous Medicaid group. When examining ER visits, this study did not find significant associations between Medicaid discontinuity and all-cause ER visits. When examining disease-specific ER visits, this study found that the incidence of these visits were 52% higher among the Discontinuous–Uninsured population, and the predicted marginal difference from the Continuous Medicaid group was found to be 0.05 additional visits per person ($P=0.008$). The Discontinuous–Insured too had a 79% higher incidence of disease-specific ER visits compared to the Continuous Medicaid population, with a predicted difference of 0.07 greater ER visits per person ($P=0.034$).

Among the primary care utilization outcomes, this study found the Discontinuous–Uninsured group to have an 86% higher incidence of disease-specific hospital outpatient visits compared to the Continuous Medicaid insurance group (IRR: 1.86; 95% CI: 1.32 – 2.63), and a predicted marginal difference of 0.21 additional visits ($P=0.007$). Finally, the Discontinuous–Uninsured group was found to have approximately 29% lower incidence of all-cause physician office visits (IRR: 0.71; 95% CI: 0.64 – 0.80), and were predicted to have 3.4 fewer overall physician office visits per person annually, when compared to their counterparts with continuous Medicaid coverage ($P<0.001$).

DISCUSSION

Adults with CVD or high-risk conditions for CVD who did not have continuous full-year Medicaid coverage, due to either loss of coverage, or disruptions/gaps, reported significantly worse health care and medication access, and routine healthcare checkup visits. Additionally, those in with discontinuous Medicaid coverage reported lower

Table 4-5. Associations between Medicaid Discontinuity and Health Care Resources Utilization in the Study Population

Outcome	Medicaid Coverage	Incidence Rate Ratio (95% Confidence Interval)	Predicted Difference in Visits
Inpatient Visits			
All-Cause	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	1.14 (0.95 – 1.37)	0.05 visits
	Discontinuous–Insured	1.24 (0.97 – 1.58)	0.08 visits
Disease-Specific	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	1.62 (1.19 – 2.22)**	0.05 visits**
	Discontinuous–Insured	1.95 (1.23 – 3.10)**	0.08 visits*
Emergency Room Visits			
All-Cause	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	0.96 (0.79 – 1.17)	-0.02 visits
	Discontinuous–Insured	1.00 (0.73 – 1.38)	0.00 visits
Disease-Specific	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	1.52 (1.14 – 2.03)**	0.05 visits**
	Discontinuous–Insured	1.79 (1.16 – 2.77)**	0.07 visits*
Hospital Outpatient Visits			
All-Cause	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	0.77 (0.55 – 1.06)	-0.33 visits
	Discontinuous–Insured	1.20 (0.71 – 2.04)	0.28 visits
Disease-Specific	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	1.86 (1.32 – 2.63)***	0.21 visits**
	Discontinuous–Insured	0.86 (0.47 – 1.60)	-0.04 visits

Table 4-5. (Continued)

Outcome	Medicaid Coverage	Incidence Rate Ratio (95% Confidence Interval)	Predicted Difference in Visits
Physician Office Visits			
All-Cause	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	0.71 (0.62 – 0.81)***	-3.37 visits***
	Discontinuous–Insured	0.89 (0.77 – 1.11)	-1.27 visits
Disease-Specific	Continuous	1.00 (Reference)	
	Discontinuous–Uninsured	0.90 (0.80 – 1.02)*	-0.30 visits
	Discontinuous–Insured	0.86 (0.71 – 1.06)	-0.43 visits

Study population included MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with or having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with CVD or high-risk conditions for CVD, and who reported having Medicaid anytime during the year. Estimates in **bold** indicate statistical significance at $P < 0.05$.

All models were adjusted for age, gender race/ethnicity, marital status, family size, census region, education, income, employment, MSA, SSI/TANF, Medicaid HMO/MCO, mental/substance abuse illnesses, arthritis/joint pain, no. of CVD conditions, perceived health status, perceived mental health status, Charlson Comorbidity Index score, and survey year.

* $P < 0.05$, ** $P < 0.01$, *** $P < 0.001$ for comparison between the continuous and discontinuous Medicaid groups

physician office visits, and higher inpatient, and ER hospitalizations. In examining these associations, this study further classified individuals with discontinuous Medicaid coverage into those who did not report any other coverage (the Discontinuous–Uninsured), and those who reported having other sources of insurance coverage during the year (the Discontinuous–Insured), in order to capture the heterogeneity of the discontinuous Medicaid group. In doing so, we found that the outcomes of Medicaid discontinuity were not always similar, either in magnitude, directionality, or statistical significance. For instance, in comparison to the Continuous Medicaid group, among both the Discontinuous–Uninsured and Discontinuous–Insured groups, the access to care outcomes were significantly poor; however, the odds of reporting difficulty in obtaining medical care were over 4 times, and almost 2 times higher, respectively. Similarly, the Discontinuous–Uninsured population in this study had significantly lower odds of reporting routine medical checkup within the past year, whereas no significant difference in this outcome was observed between the Discontinuous–Insured and the Continuous Medicaid populations.

Nearly three-quarters (74.1%) of the individuals who did not have continuous full-year Medicaid coverage in this CVD population belonged to the Discontinuous–Uninsured group. On average, this group was uninsured for over 5 months during the year, and around 18% of individuals experienced multiple transitions in their Medicaid coverage, indicating significant churning, and relapses into the state of being uninsured. Of the roughly 400,000 individuals who had Medicaid at the beginning of the year, a whopping 78% were found to be uninsured by the end of the year. These individuals collectively had worse outcomes associated with Medicaid discontinuity when compared to those who reported other sources of insurance coverage during the year, the Discontinuous–Insured, with significantly fewer physician office visits and higher incidences of inpatient and ER hospitalizations, when compared to those with continuous Medicaid coverage. Although preventive services use, and primary care visit outcomes were not significantly poor among the Discontinuous–Insured, these individuals did collectively have poor access to medical care and pharmacotherapy, and higher incidences of hospitalizations when compared to those with continuous Medicaid coverage. Our study findings indicate that for individuals with CVD or some of its risk factor conditions that require continuous, prolonged disease management, lacking continuous Medicaid coverage may adversely impact access to care, and primary care, and have a strong association with increased incidences of hospitalizations, irrespective of whether or not they have other sources of coverage during the year. And these outcomes may worsen among individuals and families for whom Medicaid is the only source of health insurance coverage.

The increased hospitalizations, both inpatient and ER, were The results of this study, are in line with previous estimates, however, very few have examined Medicaid discontinuity outcomes among adults with CVD or high-risk conditions for CVD. Among the earliest studies, Lurie and colleagues examined access to care among a cohort of California Medicaid enrollees within a 6-month period after loss of coverage. A significantly higher proportion of individuals in the lost coverage group (62%) reported poor access to needed medical care, compared to those still on Medicaid (7%), with the

discontinuous Medicaid group being more likely to have uncontrolled hypertension.¹⁰⁴ In a follow-up study conducted at one year found no improvements in access to care between the discontinuous and continuous Medicaid groups.¹⁰⁵ Carlson and colleagues examined the impact of disrupted or lost Medicaid coverage within a year of benefit structure changes among the Oregon Medicaid program, on unmet medical care or pharmacotherapy needs, utilization, and medical debt.¹⁰⁶ They found that enrollees with discontinuous Medicaid coverage were more likely to report unmet medical care, prescription medication needs (lost Medicaid coverage group), and were less likely to report having primary care visits, when compared to those enrolled continuously in the program.

Harman and colleagues examined associations between interruptions in a state Medicaid program coverage and inpatient psychiatric services among persons with schizophrenia.¹⁰ They found an 86% increase in the incidence of hospitalizations, representing 0.63 more psychiatric hospitalizations per beneficiary, and longer hospital stays to be associated with discontinuous Medicaid coverage. Hall and colleagues examined the impact of gaps in Florida's Medicaid coverage on health care utilization and expenditures among individuals with diabetes.¹⁰¹ That study found that gaps in coverage were associated with higher inpatient and ER utilization, and higher expenditures, especially in the period after lapse in Medicaid coverage. Banerjee and colleagues analyzed the MEPS data for the years 2000 – 2004 to examine association between Medicaid discontinuity and health resource utilization among all Medicaid enrollees.¹⁵ They defined discontinuity as having one or more than one transitions into or out of Medicaid coverage, which was different from the present study. The aforementioned study found Medicaid discontinuity to be associated with higher inpatient, ER, and higher outpatient physician visits. It is unclear whether they examined hospital outpatient visits and physician office visits separately, or combined the two outcomes together. Nonetheless, the results from this study are broadly in agreement with the study of Banerjee and colleagues, indicating that Medicaid discontinuity, either due to coverage loss or disruptions in coverage continuity have detrimental outcomes, both overall, and among individuals with chronic conditions. Alternatively, it may be possible that the results obtained in the analyses of Medicaid enrollees in general, may in particular be contributed by individuals with chronic, debilitating conditions, such as CVD or conditions that are high risk factors for CVD, as this study has shown.

Medicaid provides much-needed health care coverage to the poor and disabled vulnerable populations who often are unable to afford private health insurance. Medicaid is often referred to as a “Cadillac” health insurance program, which provides a rich benefits package at very little or even no cost to the enrollees.¹⁰⁷ A substantial body of research has shown that low-income Medicaid enrollees have comparable access to physician care, preventive services, and unmet medical care needs when compared to their counterparts with private health insurance.¹⁰⁷ Even low-income mothers on Medicaid report similar preventive services use and access to care compared to those with private insurance.¹⁰⁸ Together, these and other estimates shown the importance of Medicaid coverage for low-income populations, and dispel the notion that Medicaid is a costly, low-quality health coverage program.^{109,110} Loss of Medicaid coverage may

deprive these vulnerable populations suffering from chronic conditions of the ability to afford and seek much-needed medical care and pharmacotherapy. Such disruptions in continuity of care may be responsible for respondents reporting lack of usual source of care and difficulties in obtaining medical care and prescription drugs. The Discontinuous–Uninsured group also reported lower incidence of physician office visits overall, which may be precipitated by the lack of usual source of care, and inability of afford physician visit reimbursements following loss of coverage. This group reported higher rates of hospital outpatient visits compared to those with continuous Medicaid coverage, a finding that was puzzling. It may be that these individuals may be more inclined to seek care in the outpatient setting following loss of Medicaid coverage, and subsequent inability to afford physician office visits as highlighted earlier. The higher incidences of disease-specific hospitalizations observed in this study could have been precipitated by poor access to care, lower likelihood of getting routine medical checkups, lower primary care office visits, and the disruptions in continuity of care resulting from loss of Medicaid coverage. A follow-up evaluation of the Oregon Health Experiment, found higher ER utilization among individuals who gained Medicaid coverage, relative to those who remained uninsured.¹¹¹ The present study found increases in CVD or high-risk conditions for CVD-specific ER visits among the discontinuous Medicaid groups compared to the Continuous Medicaid group, which may have been precipitated by poor access to care as observed in our findings, and also due to disruptions in disease management caused by loss or gaps in Medicaid coverage. Differences in the study population (diseased vs. general population, and national vs. state-specific) and estimation techniques may have contributed to the differences in our findings from the Oregon analysis of ER utilization.

This study has several limitations. State-level Medicaid eligibility data was not available to explore the complex and dynamic patterns of discontinuity that may differ based on the state's eligibility and reporting criteria. This study found strong associations between Medicaid discontinuity and poor access to care, lower odds of routine medical checkup, higher disease-specific hospitalizations, and lower all-cause physician office visits. Nonetheless, temporality between discontinuity and the various outcomes, and the resulting causality could not be established due to data limitations. The regression analyses may not have adjusted for unobserved covariates that may have confounded the impact of certain risk factors on the likelihood of belonging to the Discontinuous–Uninsured or the Discontinuous–Insured groups. Data in MEPS are self-reported, and the reasons for discontinuity of Medicaid coverage were not captured, resulting in estimates that may not exactly reflect the national estimates of the impact of Medicaid discontinuity on access to care, preventive services use, and health care services utilization. Researchers have speculated that most surveys underestimate the number of Medicaid enrollees due to a variety of reasons including stigma of public insurance. However, such misreporting is less of a problem for MEPS compared to other federal health surveys.⁵⁴ Finally, findings from this study are representative of the non-institutionalized civilian population and may not be applicable to individuals with prolonged hospitalizations or admitted to long-term care facilities.

Loss of Medicaid coverage has an adverse impact on individuals' ability to seek medical care which may in turn lead to poor health outcomes and increased burden on the health care system resulting from providing care to uninsured individuals. Individuals eligible for Medicaid may also lose coverage due to the inefficient administrative hurdles and burdensome paperwork requirements, and may be considered a policy failure.³⁴ However, when individuals transition from one source of coverage to another, continuity of care may still be impacted, especially since different insurance plans have different network of providers, and differing benefit structures, and may cover different classes of medications and services. The present study highlighted the adverse impact that Medicaid discontinuity may have on ability to access needed medical care, and the increased hospitalizations for CVD or high-risk conditions for CVD associated with coverage instability. Enabling continuity in Medicaid coverage for low-income populations with chronic and debilitating conditions would ensure increased effectiveness of the care they receive to manage these disease, and may prove cost-effective in the long run.

Discontinuity in Medicaid and the Children's Health Insurance Program (CHIP) coverage among low-income children is an issue that has gained considerable scrutiny previously, and the detrimental effects of such a policy failure have been well established.^{35,112} Apart from simplification of administrative regulations governing policy renewals, a policy initiative adopted by most states is the 12 months continuous eligibility provision. Under this policy, children continue to be enrolled under Medicaid/CHIP irrespective for 12 months after gaining or renewing coverage, irrespective of any income or other eligibility midway. Researchers and advocacy groups, such as the Medicaid and CHIP Payment Access Commission, have advocated for extending such a measure for non-elderly adults to ensure continuity of Medicaid coverage in this population.^{8,44} Our study findings indicate that individuals with chronic conditions requiring prolonged and continuous disease management, such as CVD or its high-risk conditions, may be especially adversely impacted by instability in their Medicaid coverage. States can choose to have a straight-forward option of establishing a minimum guaranteed eligibility period, such as the 12 month continuous eligibility provision, for subpopulations of non-elderly adults with chronic conditions, such as CVD, or other such as cancers, respiratory illnesses, mental illnesses, HIV/AIDS, among others, which require prolonged management of disease symptoms and risk factors. By using annual eligibility redetermination periods, and guaranteeing enrollment for the period in between, states can ensure that loss of coverage or churning midway during enrollment can be reduced in these vulnerable populations.

CONCLUSION

The present study highlighted the significantly worse access to care, preventive medical checkups, lower primary care office visits, and higher hospitalizations for CVD or high-risk conditions for CVD, associated with having discontinuous Medicaid coverage. Policy initiatives to stabilize Medicaid coverage among non-elderly adults may provide the most benefit for individuals with chronic conditions requiring continuous long-term medical care interventions, and appropriate disease management. Low-income

adults managing chronic conditions, especially in regions with stringent Medicaid eligibility regulations, must be encouraged to ensure continuity of Medicaid coverage. States may adopt annual eligibility reassessment policies, with continuous eligibility for the intermittent period, for non-elderly adults with certain chronic conditions among whom continuous Medicaid coverage would ensure greater effectiveness of disease management.

CHAPTER 5. CONCLUSION

SUMMARY

The present study was aimed at conducting a detailed examination of the impact of having discontinuous Medicaid coverage on various health care utilization outcomes and medication consumption behaviors, which are extremely crucial in effectively preventing and treating chronic diseases of the cardiovascular system along with many of its risk factor disease conditions. Although previous estimates have characterized general populations that experience discontinuous Medicaid coverage, and compared them with their counterparts with continuous Medicaid insurance, none have focused on national estimates of subpopulations diagnosed with complex chronic conditions requiring prolonged disease management, such as CVD or chronic conditions which are high risk factors for CVD. The first aim of this dissertation intended to characterize the populations of non-elderly adults with CVD or high-risk conditions for CVD who have discontinuous Medicaid coverage, and compare them to their counterparts with continuous Medicaid coverage. Another difference between this analyses and previous ones is the classification of the subpopulations with less than full-year Medicaid coverage. Individuals with discontinuous Medicaid coverage are not a homogenous population, since although many individuals may not have any other source of insurance coverage during the year, the Discontinuous–Uninsured group in this study, this population would also comprise of individuals with other sources of coverage, ranging from private insurance, Tricare, or other hospital-, or physician group-based health insurance. Interestingly, this study found about three-fourths of this population to belong to the Discontinuous–Uninsured group, and a little over a fourth to belong to the Discontinuous–Insured group.

Hypertension, lipid disorders, and diabetes, significant risk factors for CVD, were found to be the most prevalent chronic conditions in this study population. These are diseases that require prolonged, and sometimes lifelong treatment, which includes adequate preventive care to monitor health or disease progression, getting regular medical checkups and making regular primary care visits to diagnose and treat these conditions early, and filling prescription medications regularly and consuming them as prescribed by health care providers. This finding, coupled with the fact that insurance is a major source that enables individuals, especially low-income sick populations, to seek needed medical care and pharmacotherapy, further strengthens our justification for examining the impact of Medicaid discontinuity among low-income populations with CVD or high-risk conditions for CVD.

When examining Medicaid coverage patterns among the two discontinuous Medicaid coverage groups, this study found that on average, the Discontinuous–Uninsured were uninsured for over 5 months during the year, and the Discontinuous–Insured too were uninsured on average for a little over a month, despite reporting other insurance sources. Further, a subsample of individuals who had coverage at the beginning of the year were examined to determine their coverage patterns by the year end. It was observed that among the Discontinuous–Uninsured who had Medicaid coverage at the

start of the year, 78% were uninsured by the year end. Among the Discontinuous-Insured, only 8% had Medicaid, 86% had other insurance, and around 6% were uninsured by the end of the year. Finally, this study found that almost 18% of both the discontinuous Medicaid groups had more than one transitions, indicating a decent amount of churning taking place in these groups, i.e. going back-and-forth from Medicaid to either a state of being uninsured, or enrolling in another source of insurance coverage, or vice versa (uninsured – Medicaid – uninsured, or other insurance – Medicaid – other insurance). These findings highlight the instability in Medicaid coverage encountered by low-income, chronically diseased, non-elderly adult population, and the disruptions in continuity of care associated with transitioning from Medicaid to other sources of coverage, or vice versa, not to mention the difficulties involved with adjusting to differences in provider networks, services covered, services and resources covered, and benefit structures.

When examining the population characteristics associated with Medicaid discontinuity among the Discontinuous–Uninsured relative to the Continuous Medicaid group, this study found that those who were married, resided in the South, had intermittent unstable employment, and had fair to poor perceived health status, had higher odds of having discontinuous Medicaid coverage, whereas men, minorities, those with higher incomes, SSI/TANF beneficiaries, Medicaid managed care enrollees, and those with respiratory illnesses had lower odds of experiencing discontinuity in their Medicaid coverage during the year. Among the Discontinuous–Insured group, married individuals, those with higher education, higher incomes, and having some form of employment during the year were the characteristics associated with higher odds of Medicaid discontinuity, whereas minorities, those who were married, SSI/TANF beneficiaries, Medicaid managed care enrollees, and those with respiratory illnesses were associated with lower odds of Medicaid discontinuity. These findings may have highlighted racial and ethnic differences in Medicaid continuity; for instance, low-income minorities have been known to have lower incomes and education, and may be more likely to be unemployed, resulting in greater continuity in their Medicaid coverage compared to their Caucasian counterparts. An interesting finding was the lower odds of Medicaid discontinuity among Medicaid managed care enrollees, which may be the result of better outreach efforts by managed care organizations, although they may be incentivized to do so given they are paid on a per-member basis by Medicaid.³⁴

When examining the associations between Medicaid discontinuity, medication adherence, and medication utilization, as part of the second aim of this dissertation, it was found that although the point estimate for overall average adherence to frequently prescribed medications was lower for the Discontinuous–Uninsured, the estimate was not statistically significantly different from the Continuous Medicaid group. This study found lower adherence to sulfonylureas among the Discontinuous–Uninsured compared to the Continuous Medicaid, which may result in higher health care costs in this subpopulation.¹¹³ The Discontinuous–Insured were found to have higher odds of adherence to ACEI/ARB and calcium channel blockers, which could have been contributed by better access to these medications through other insurance programs. The odds of clinically adequate adherence to diuretics, however, was found to be significantly

lower among the Discontinuous–Insured, which is traditionally used as a first-line antihypertensive therapy. Medication utilization on the other hand, both all-cause and disease-specific, was found to be significantly lower among both the discontinuous Medicaid groups compared to the Continuous Medicaid. On average, the Discontinuous–Uninsured had 12 fewer drug fills per person for all medications, and 2.3 fewer fills per person for CVD-specific medications. Even the Discontinuous–Insured had 6.5 fewer fills, and 2.8 fewer fills per person, for all-cause and disease-specific medications, respectively. Given the significantly lower prescription drug fills among the discontinuous Medicaid groups, this study may have overestimated adherence to medications, which may also be lower among populations lacking continuous Medicaid coverage.

Finally, as part of the third aim, this study examined associations between Medicaid discontinuity, access to care, preventive services use, and health care utilization. Access to care was poor among the discontinuous Medicaid populations, with these groups having lower odds of having usual source of care, and higher odds for experiencing difficulties in obtaining needed medical care and medications. The odds of routine medical checkups were also found to be lower among the Discontinuous–Uninsured, indicating inadequate medical checkups which are critical for early detection and management of chronic conditions. Incidences of disease-specific hospitalizations, both inpatient and ER visits, were found to be significantly higher among the discontinuous Medicaid populations compared to those with continuous Medicaid. Although ER use has been found to be higher among individuals who gain Medicaid coverage, as ER may be a usual source of care for certain vulnerable populations even for conditions that can be managed in outpatient setting, the increased hospitalizations are more likely due to poor management of CVD or high-risk conditions for CVD resulting from poor access to care and inadequate pharmacotherapy. The reason being, when individuals are asked about whether or not they have a usual source of care, the options include ER as well, apart from provider offices, outpatient and other clinic settings. Moreover, all-cause office-based provider visits were lower, whereas all-cause hospital outpatient visits were higher among the Discontinuous–Uninsured, which may indicate inability to seek care and build patient-provider relationships with a usual source of primary care physician, and overreliance on the hospital clinical care setting. It may also be a result of lower acceptance of Medicaid patients by private practice physicians, however, the comparison group is those with continuous Medicaid coverage, which would have had worse primary care office visits if lack of acceptance were predominant and prevalent. These findings indicate that Medicaid discontinuity among low-income CVD or high risk chronic condition patients is associated with poor access to care, inadequate medical checkups, lower prescription drug utilization and poor adherence to certain commonly prescribed therapeutic drug classes, higher CVD-related hospitalizations and ER visits, lower overall physician office visits to seek primary care, and a greater reliance on hospital outpatient setting.

IMPLICATIONS FOR POLICY

Gaps in Medicaid coverage may have deleterious health consequences for low-income enrollees, and this study found that the health care access, and utilization outcomes, and medication consumption behaviors among adults with CVD or high-risk conditions for CVD may be adversely impacted due to instability in Medicaid coverage. Continuity of Medicaid, especially for those who continue to be eligible for coverage, represents a cost-effective policy solution, and lead to greater effectiveness of the prevention and treatment services and resources, especially among low-income diseased populations. State Medicaid agencies have a policy option of guaranteeing enrollment for adults in Medicaid until the next reenrollment and eligibility determination cycle, and have the duration of these cycles to be six months or annual. This would significantly reduce within enrollment period churning and ensure better continuity of Medicaid coverage, and in turn the continuity of care afforded by it. On a long term basis, cost reductions may be realized due to providing better continuity in Medicaid coverage to non-elderly adults, especially if the administrative costs associated with frequent reenrollment are higher. The immediate impact of prolonged continuity of Medicaid may however, result in a higher financial burden on the states, which may be a reason for the stringent administrative regulations governing Medicaid coverage for none-elderly adults. If that may be the case, then rather than a blanket policy that provides guaranteed Medicaid enrollment to all non-elderly adults, state Medicaid agencies may focus on specific subpopulations suffering from chronic, debilitating disease conditions, who often require prolonged medical care to manage their conditions and comorbidities, and for whom having continuous Medicaid coverage may prove to be more effective, and cost-effective. Indeed, the present study showed that adults with CVD or high-risk conditions for CVD had greater access to care problems, substandard medication adherence, inadequate preventive care and primary care, and higher incidences of hospitalizations associated Medicaid discontinuity, after adjusting for a host of demographic, socioeconomic, health-related and Medicaid eligibility-related covariates in order to obtain robust estimates. And although this study did not examine individuals with other disease conditions such as respiratory illnesses, HIV/AIDS, mental illnesses, these subpopulations too may experience similar negative outcomes of poor access to and use of health care resources. Therefore, policies to stabilize Medicaid coverage among the low-income population with chronic, debilitating disease conditions, may provide more bang for the buck.

Instability in insurance, both public and private, have negative impact on health and well-being of individuals as has been established by studies above. Although there has not been any positive impact of this loss of and gaps in coverage investigated on patients and society, there are some aspects or reasons of that contribute to churning in insurance or Medicaid that are needed for the benefit of the society. One is the need for deductibles and co-payments for insurance programs. Studies have shown that inability to afford paying for deductibles and co-payments is one of the reasons that people opt out of insurance plans, or do not enroll for public health plans. Although high deductibles and co-pays have a negative effect on individuals' utilization of needed care, complete elimination of these payment structures might also be detrimental. This is because with

“no skin in the game”, there will be the likelihood of what economists call “moral hazard”, that is individuals will consume more health care than what they need resulting in inefficient allocation and spending of resources and excessive health care burden on the society. In order to promote rational utilization of health care resources such contribution from health plan enrollees is needed; the challenge to determining the “sweet spot” or the right balance that would ensure that it does not exert a tremendous financial burden. Another aspect is the requirement for sufficient documentation to ensure that means-tested insurance plans enroll individuals who truly qualify for the health plans, and prevent any fraud which may divert needed health care resources to individuals who may not need them. Once again, rather than getting rid of documentation requirements, the process should be made less burdensome, with greater choices or avenues given to patients, such as online enrollment, and longer time periods, and note having short periodic reporting requirements to report eligibility. Also, provisions that enable smooth transitions from public to private health insurance, or vice versa, especially for low-income population, can also minimize churning and gaps, while still ensuring that minimum checks and balances of determining eligibility are in place to prevent wastage of much needed health care resources.

As part of the 2010 PPACA provisions of Medicaid expansion, states have the option for expanding Medicaid coverage to their non-elderly adult residents with incomes at or below 138% FPL. As of May 2015, however, 18 states have chosen status quo over adopting these provisions and expanding Medicaid coverage, which would result in continuation of these issues of instability of Medicaid in non-elderly adults and the resulting adverse impact on continuity and quality of care and disease management.⁴⁹ Even under PPACA provisions, states will be required to have 12 month recertification periods, meaning individuals will have to reenroll in Medicaid annually; however, even within this enrolled period, individuals and families may lose coverage midway if their incomes and other criteria change. For these states, ensuring continuous Medicaid coverage, at least for the diseased vulnerable populations could be a policy initiative that can be adopted to ensure better health outcomes among these populations, and possibly lower financial burden on the health care system due to providing high cost care resulting from avoidable hospitalizations. Due to competing priorities, and even among the health care sector, competing health priorities, resources available for management of these conditions may be limited and policy initiatives to ensure continuity of insurance coverage among low-income adult populations may be lacking. Failure to address issues of Medicaid coverage continuity among adults with complex chronic conditions such as CVD, may lead to large increases in prevalence of avoidable chronic conditions, which may exert a tremendous burden on the health care system and the economy.

IMPLICATIONS FOR FUTURE RESEARCH

The present study found strong associations between Medicaid discontinuity and the negative access to care and health care resource utilization outcomes among the low-income CVD adult population. However, due to data limitations, we were unable to establish temporality between Medicaid discontinuity and various outcomes under

investigation. Moreover, the reasons for losing or gaps in Medicaid coverage were not known. Future research may incorporate these shortcomings to obtain a holistic view of the outcomes associated, and possibly caused by disruptions in Medicaid coverage. Using state-specific eligibility would also be able to better control for state-level differences in observed and unobserved characteristics that may not have been adequately controlled by geographic region.

Costs associated with Medicaid discontinuity are an important outcome, and may be of more interest to policymakers. The obvious follow-up to this study would examine the costs, including total costs and costs specific to each health care services, to provide a comprehensive picture of the impact of Medicaid discontinuity on health care resources and costs, from the health care system perspective or payer perspective. Disruptions in care continuity associated with poor access to care, lack of preventive care and primary care, and the resulting increased hospitalizations and ER visits may lead to significant increases in health care costs. And with a significant increase ER use by the uninsured to obtain their usual care, or due hospitalizations resulting from chronic disease exacerbations, the financial burden may increase tremendously on the health care providers and hospitals due to provision of uncompensated care. The insured may end up subsidizing a significant portion of the uncompensated care costs through their insurance plans, and related out-of-pocket expenditures due to cost shifting. This vicious cycle may not only exert a financial strain on providers and payers, but on patients as well. Examinations of cost implications due to Medicaid discontinuity from not just payer or provider perspective, but patient perspective may also provide meaningful information of the impact of Medicaid discontinuity, and produce a complete and comprehensive picture of the implications of Medicaid discontinuity among the non-elderly adult population, or subpopulations of these with chronic conditions such as CVD or several of its high-risk conditions.

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APPENDIX A. SUPPLEMENTAL MATERIALS FOR CHAPTER 3

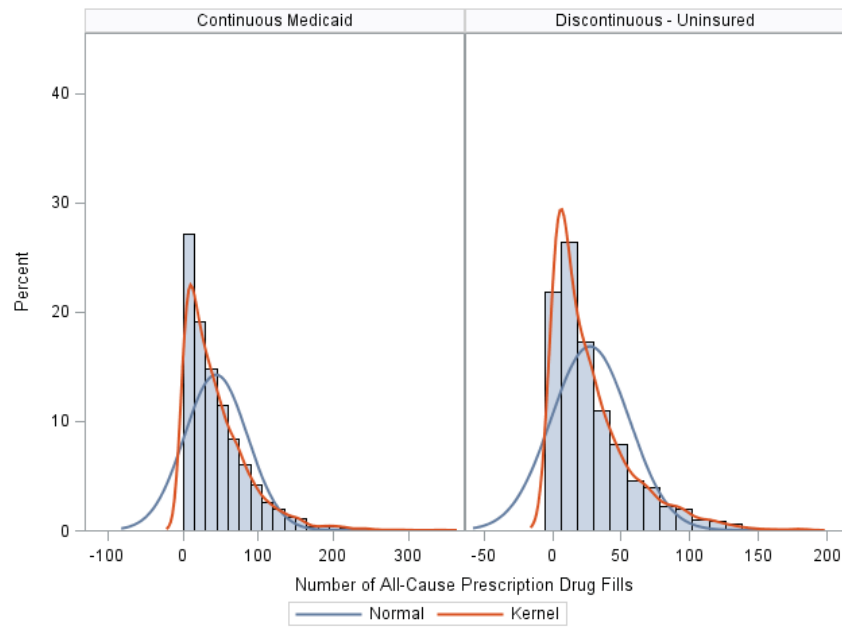


Figure A-1. Distribution of All-Cause Prescription Drug Fills among Continuous Medicaid and Discontinuous-Uninsured

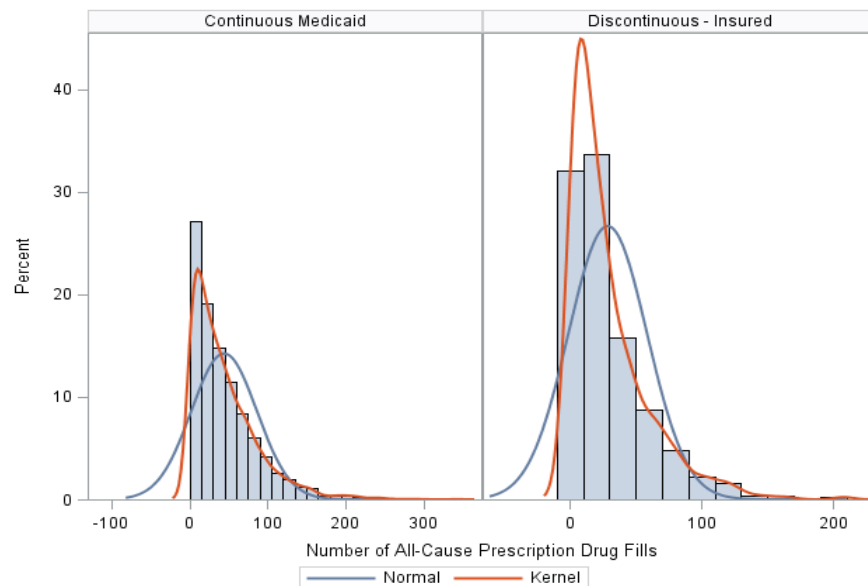


Figure A-2. Distribution of All-Cause Prescription Drug Fills among Continuous Medicaid and Discontinuous-Insured

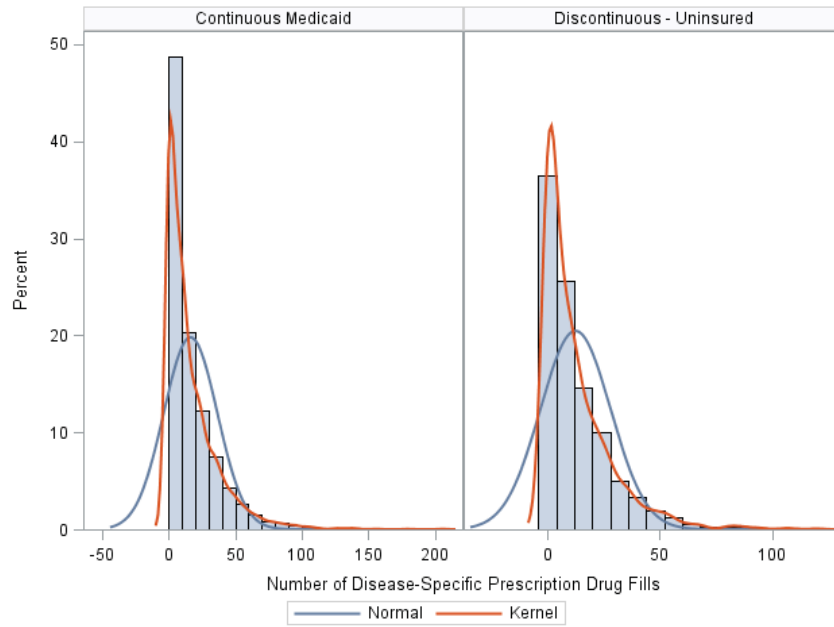


Figure A-3. Distribution of Disease-Specific Prescription Drug Fills among Continuous Medicaid and Discontinuous–Uninsured

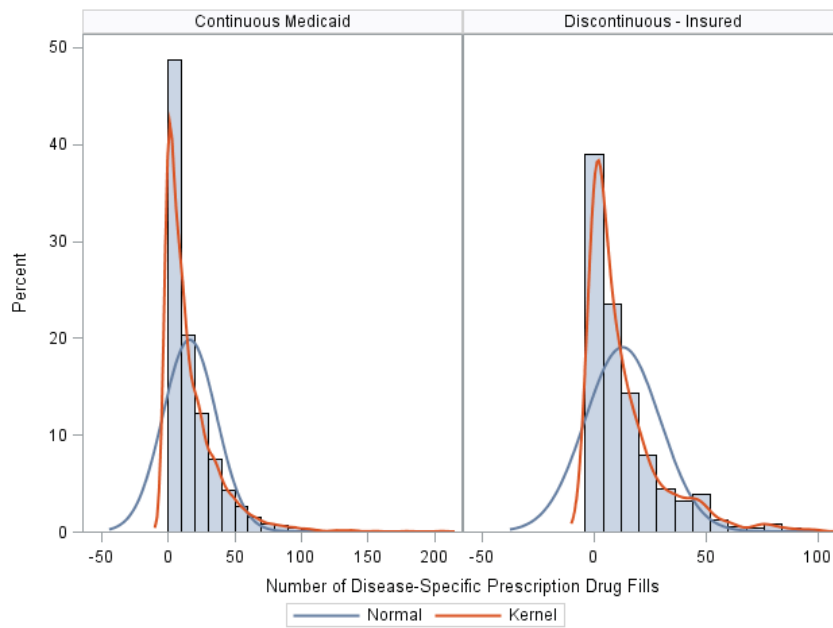


Figure A-4. Distribution of Disease-Specific Prescription Drug Fills among Continuous Medicaid and Discontinuous–Insured

Table A-1. Characteristics Associated with Medication Adherence

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
DEMOGRAPHIC		
Age		
18 – 34 years	1.00 (Reference)	1.00 (Reference)
35 – 49 years	1.25 (0.97 – 1.61)	1.16 (0.90 – 1.49)
50 – 64 years	1.26 (0.97 – 1.63)	1.22 (0.93 – 1.59)
Gender		
Female	1.00 (Reference)	1.00 (Reference)
Male	1.11 (0.91 – 1.35)	1.09 (0.88 – 1.35)
Race/Ethnicity		
Non-Hispanic White	1.00 (Reference)	1.00 (Reference)
Non-Hispanic Black	0.83 (0.67 – 1.04)	0.81 (0.65 – 1.01)
Hispanic	0.86 (0.68 – 1.09)	0.79 (0.61 – 1.02)
Others	0.74 (0.53 – 1.02)	0.74 (0.54 – 1.01)
Marital Status		
Not Married	1.00 (Reference)	1.00 (Reference)
Married	0.96 (0.79 – 1.17)	1.05 (0.86 – 1.30)
Region		
Northeast	1.00 (Reference)	1.00 (Reference)
Midwest	1.12 (0.84 – 1.49)	1.01 (0.76 – 1.37)
South	1.05 (0.83 – 1.34)	1.00 (0.78 – 1.29)

Table A-1. (Continued)

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
West	1.02 (0.79 – 1.31)	1.02 (0.79 – 1.32)
Family Size	0.97 (0.93 – 1.02)	0.95 (0.90 – 1.00)
SOCIOECONOMIC		
Education		
<High School	1.00 (Reference)	1.00 (Reference)
High School/GED	0.99 (0.84 – 1.18)	0.97 (0.81 – 1.71)
College Degree	0.85 (0.61 – 1.19)	0.85 (0.62 – 1.17)
Income		
<100% FPL	1.00 (Reference)	1.00 (Reference)
100 – <125% FPL	0.87 (0.66 – 1.19)	0.84 (0.61 – 1.15)
≥125% FPL	1.11 (0.93 – 1.33)	1.04 (0.85 – 1.27)
Employment		
Always Unemployed	1.00 (Reference)	1.00 (Reference)
Unstable Employment	1.01 (0.72 – 1.40)	1.00 (0.73 – 1.38)
Always Employed	0.97 (0.75 – 1.25)	1.01 (0.79 – 1.30)
MSA		
Non-MSA	1.00 (Reference)	1.00 (Reference)
MSA	1.00 (0.80 – 1.25)	0.97 (0.77 – 1.22)

Table A-1. (Continued)

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
MEDICAID-RELATED		
SSI/TANF		
No	1.00 (Reference)	1.00 (Reference)
Yes	1.02 (0.87 – 1.20)	1.01 (0.84 – 1.21)
Medicaid HMO/MCO		
No	1.00 (Reference)	1.00 (Reference)
Yes	1.15 (0.97 – 1.36)	1.10 (0.92 – 1.31)
HEALTH-RELATED		
Perceived Health Status		
Excellent/Very Good/Good	1.00 (Reference)	1.00 (Reference)
Fair to Poor	1.22 (1.01 – 1.48)	1.10 (0.90 – 1.34)
Perceived Mental Health		
Excellent/Very Good/Good	1.00 (Reference)	1.00 (Reference)
Fair to Poor	1.07 (0.89 – 1.30)	1.11 (0.92 - 1.35)
CVD or High-Risk Conditions for CVD		
≤1	1.00 (Reference)	1.00 (Reference)
2	0.82 (0.63 – 1.06)	0.93 (0.70 – 1.22)
≥3	1.01 (0.76 – 1.34)	1.00 (0.72 – 1.38)

Table A-1. (Continued)

Characteristics	Discontinuous–Uninsured Odds Ratio (95% Confidence Interval)	Discontinuous–Insured Odds Ratio (95% Confidence Interval)
Mental/Substance Abuse Illnesses		
No	1.00 (Reference)	1.00 (Reference)
Yes	1.12 (0.91 – 1.38)	0.91 (0.73 – 1.13)
Respiratory Illnesses		
No	1.00 (Reference)	1.00 (Reference)
Yes	0.88 (0.70 – 1.10)	0.81 (0.63 – 1.04)
Arthritis/Joint Pain		
No	1.00 (Reference)	1.00 (Reference)
Yes	1.13 (0.93 – 1.37)	1.30 (1.04 – 1.62)*
CCI Score	1.00 (0.94 – 1.07)	1.03 (0.97 – 1.11)

Study population includes MEPS respondents from the years 2002-2011, aged 18-64 years, who either reported being diagnosed with, or reported having an event (prescription medication, or inpatient, emergency, outpatient, or office-based provider visits) associated with, 1 or more cardiovascular disease or associated risk factor comorbidities considered in this study.

*** P -value<0.001, ** P <0.01, * P <0.05 for the difference between Continuous Medicaid vs Discontinuous–Uninsured and Continuous Medicaid vs Discontinuous–Insured groups

SE: Standard error; GED: General educational development; FPL: Federal poverty level; MSA: Metropolitan Statistical Area; SSI: Supplementary Security Income due to disability; TANF: Temporary Assistance to Needy Families program participation; HMO: Health maintenance organization; MCO: Managed care organization; CCI: Charlson Comorbidity Index score (D’Hoore adapted CCI score was calculated in this study using the 3-digit ICD-9-CM codes in MEPS).

APPENDIX B. SUPPLEMENTAL MATERIALS FOR CHAPTER 4

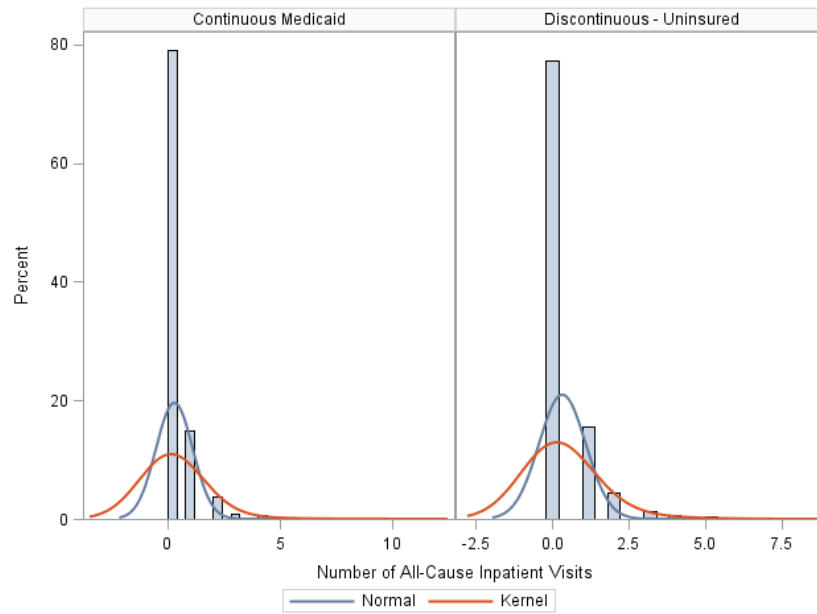


Figure B-1. Distribution of All-Cause Inpatient Visits among Continuous Medicaid and Discontinuous–Uninsured

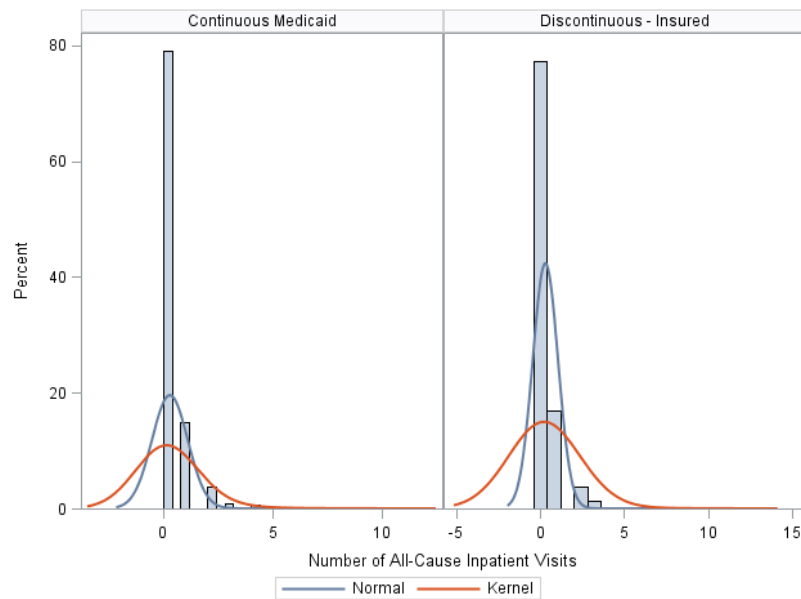


Figure B-2. Distribution of All-Cause Inpatient Visits among Continuous Medicaid and Discontinuous–Insured

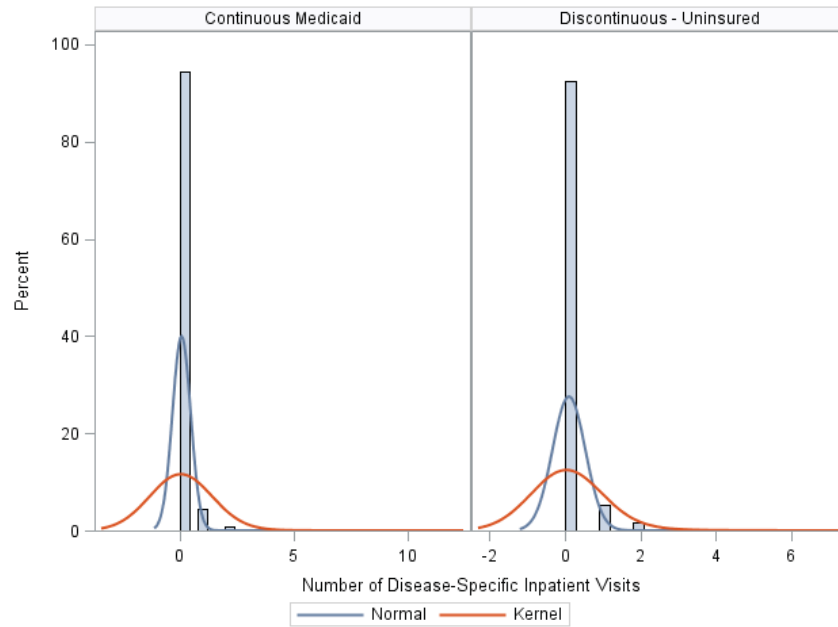


Figure B-3. Distribution of Disease-Specific Inpatient Visits among Continuous Medicaid and Discontinuous-Uninsured

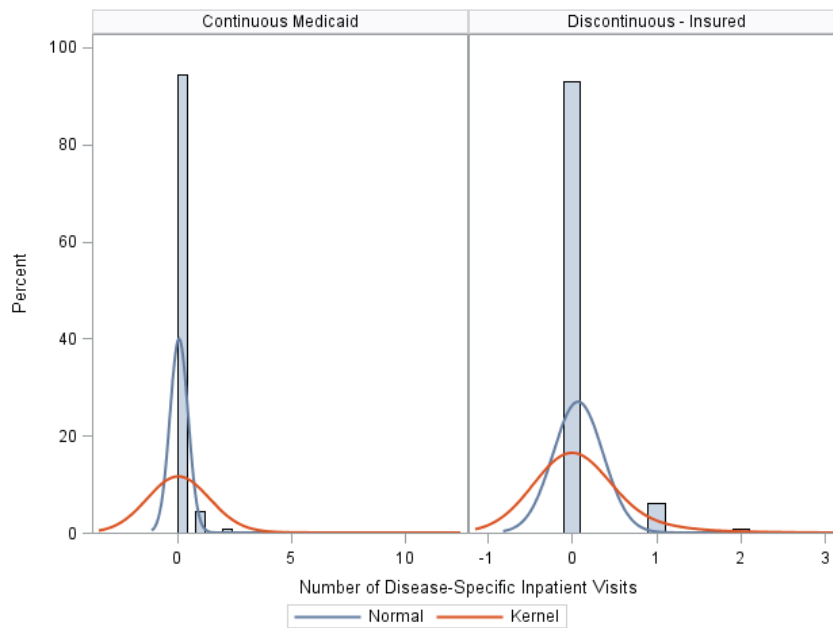


Figure B-4. Distribution of Disease-Specific Inpatient Visits among Continuous Medicaid and Discontinuous-Insured

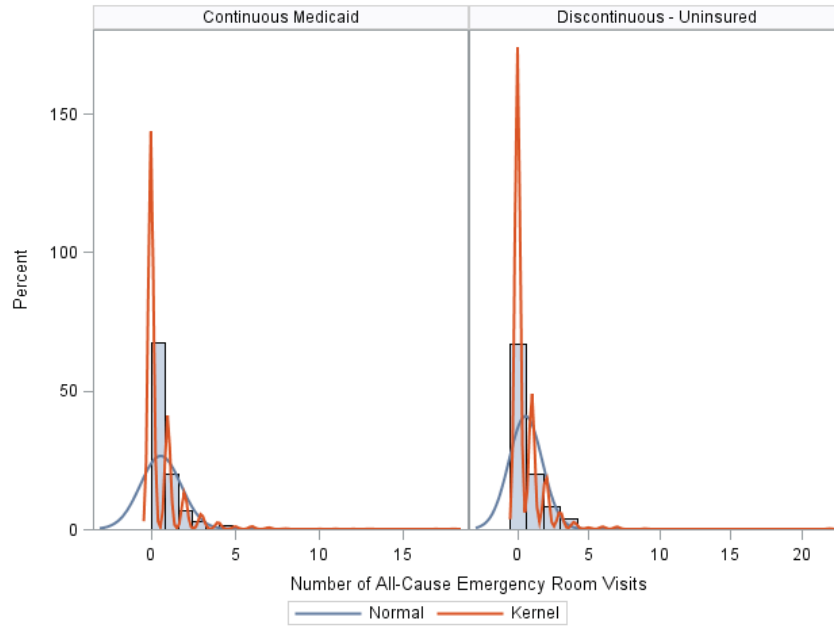


Figure B-5. Distribution of All-Cause Emergency Room Visits among Continuous Medicaid and Discontinuous–Uninsured

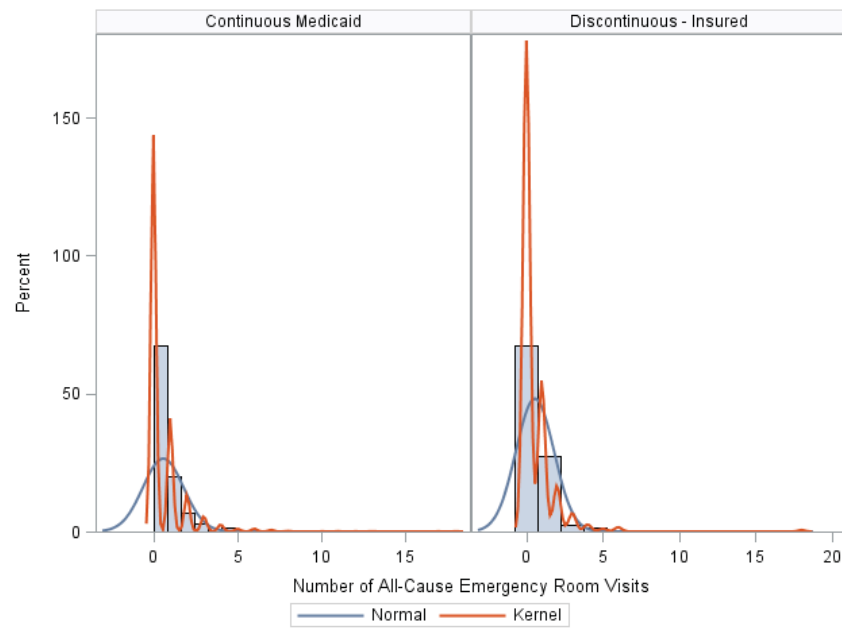


Figure B-6. Distribution of All-Cause Emergency Room Visits among Continuous Medicaid and Discontinuous–Insured

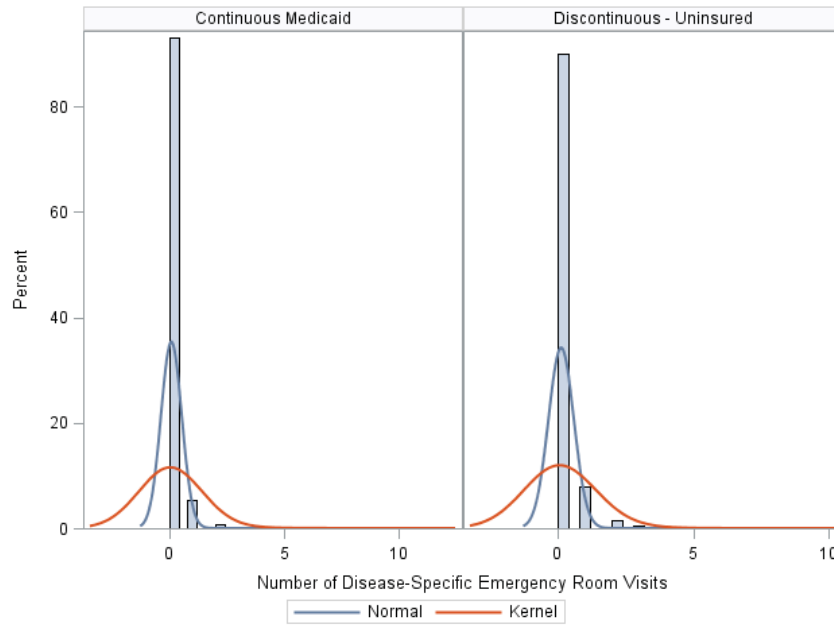


Figure B-7. Distribution of Disease-Specific Emergency Room Visits among Continuous Medicaid and Discontinuous-Uninsured

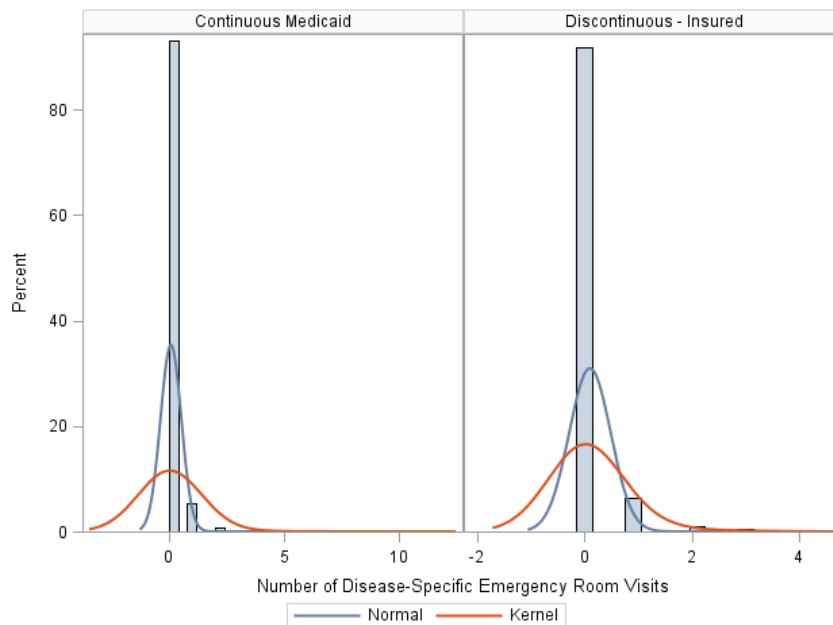


Figure B-8. Distribution of Disease-Specific Emergency Room Visits among Continuous Medicaid and Discontinuous-Insured

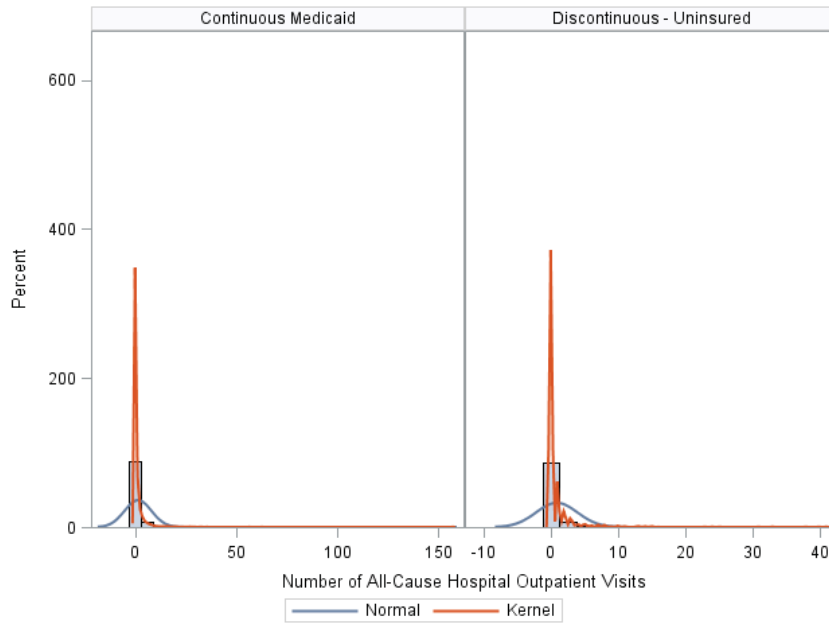


Figure B-9. Distribution of All-Cause Hospital Outpatient Visits among Continuous Medicaid and Discontinuous–Uninsured

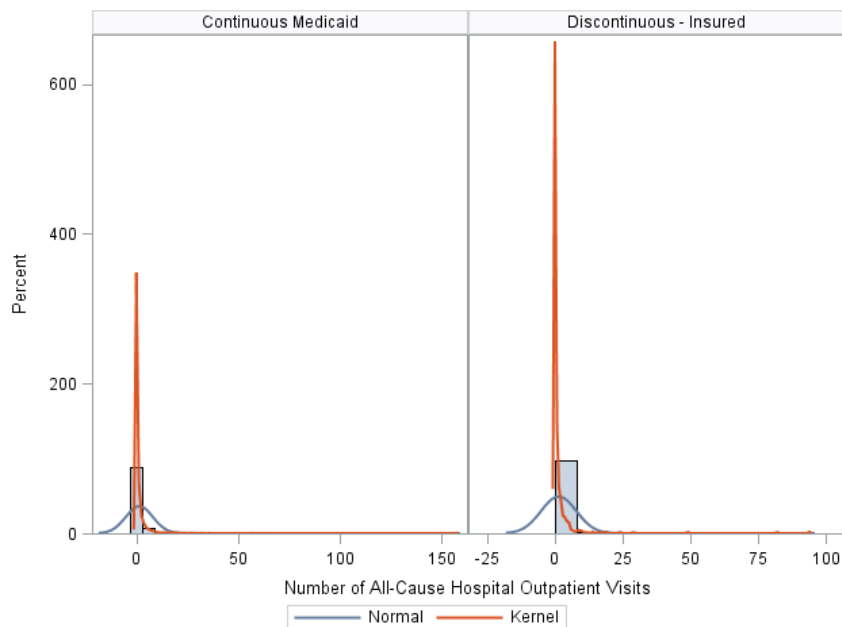


Figure B-10. Distribution of All-Cause Hospital Outpatient Visits among Continuous Medicaid and Discontinuous–Insured

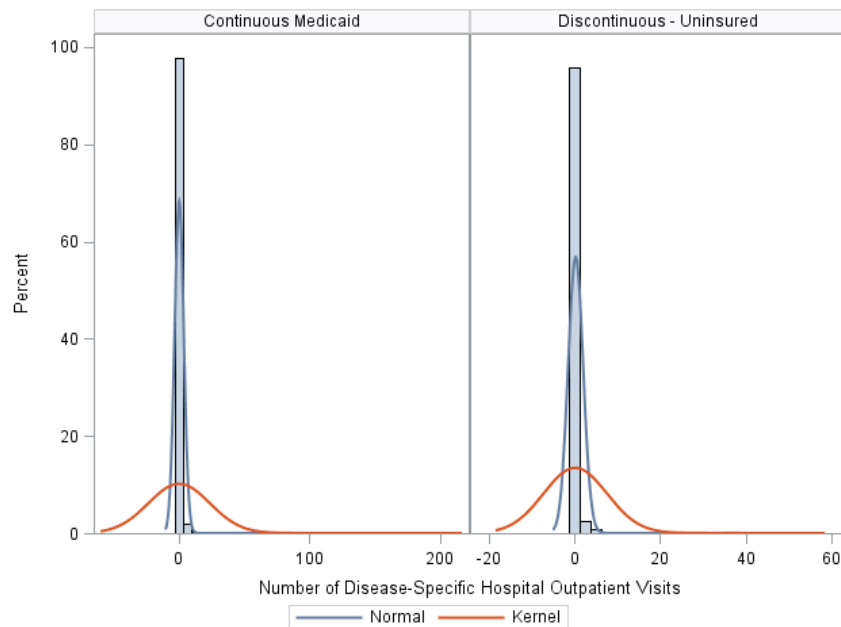


Figure B-11. Distribution of Disease-Specific Hospital Outpatient Visits among Continuous Medicaid and Discontinuous–Uninsured

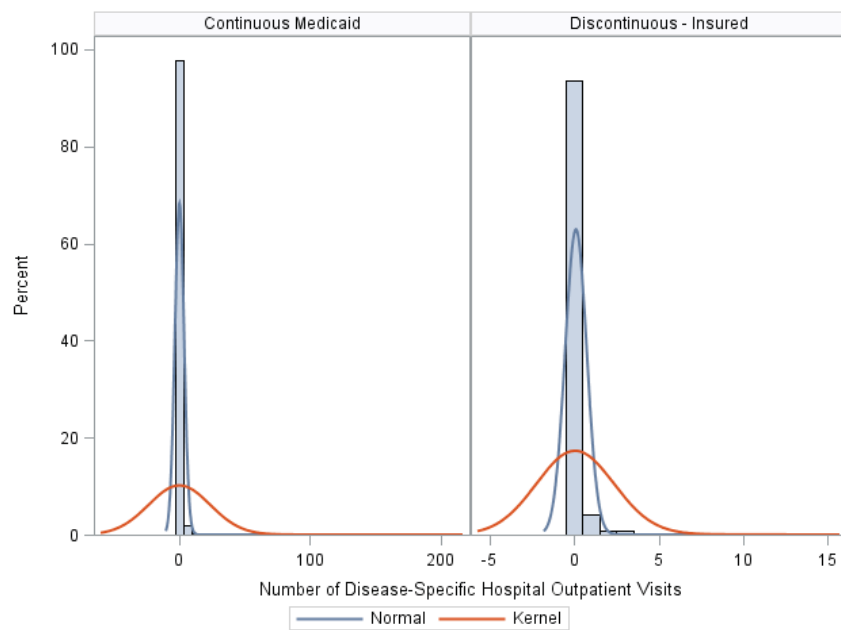


Figure B-12. Distribution of Disease-Specific Hospital Outpatient Visits among Continuous Medicaid and Discontinuous–Insured

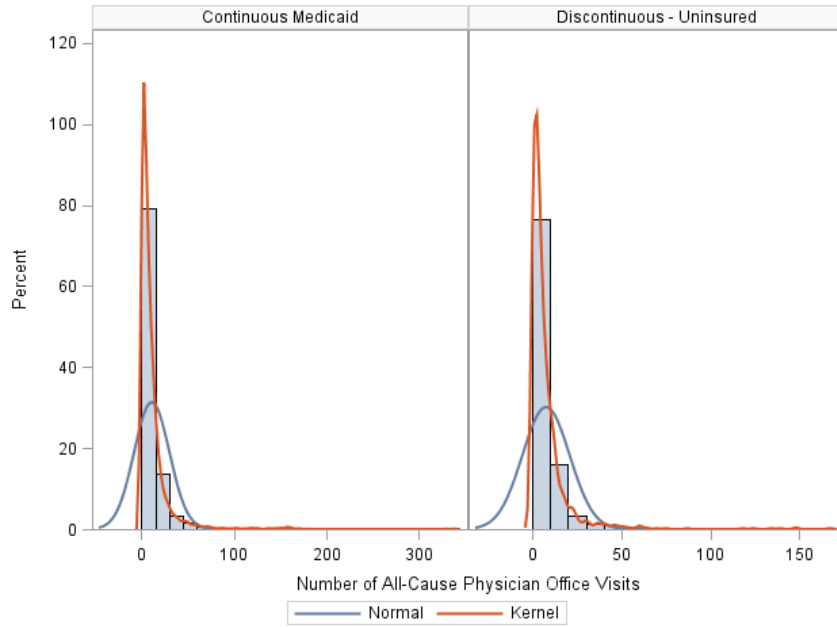


Figure B-13. Distribution of All-Cause Physician Office Visits among Continuous Medicaid and Discontinuous-Uninsured

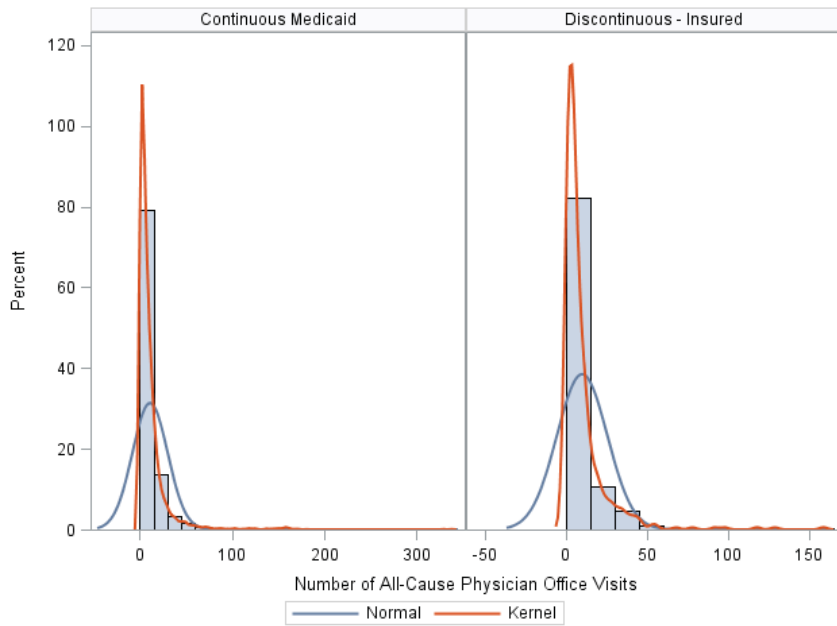


Figure B-14. Distribution of All-Cause Physician Office Visits among Continuous Medicaid and Discontinuous-Insured

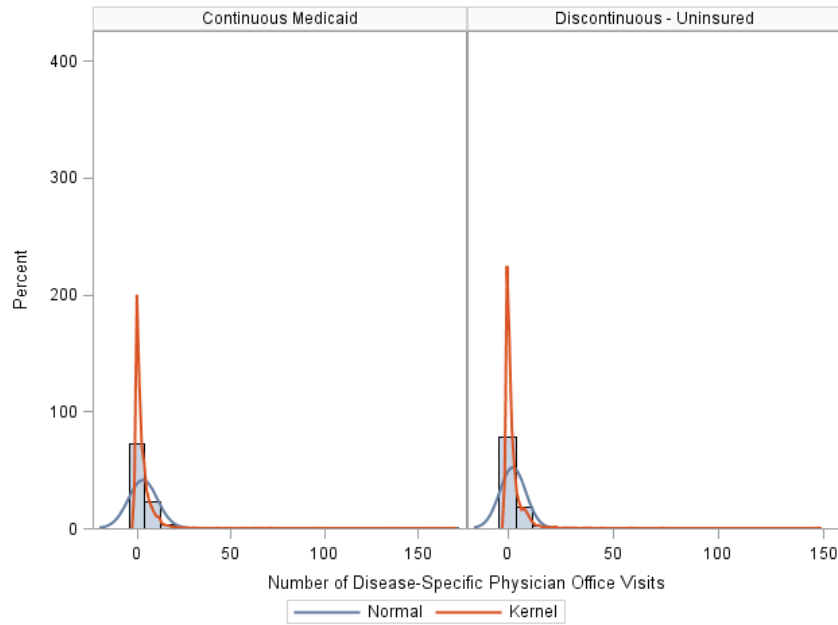


Figure B-15. Distribution of Disease-Specific Physician Office Visits among Continuous Medicaid and Discontinuous–Uninsured

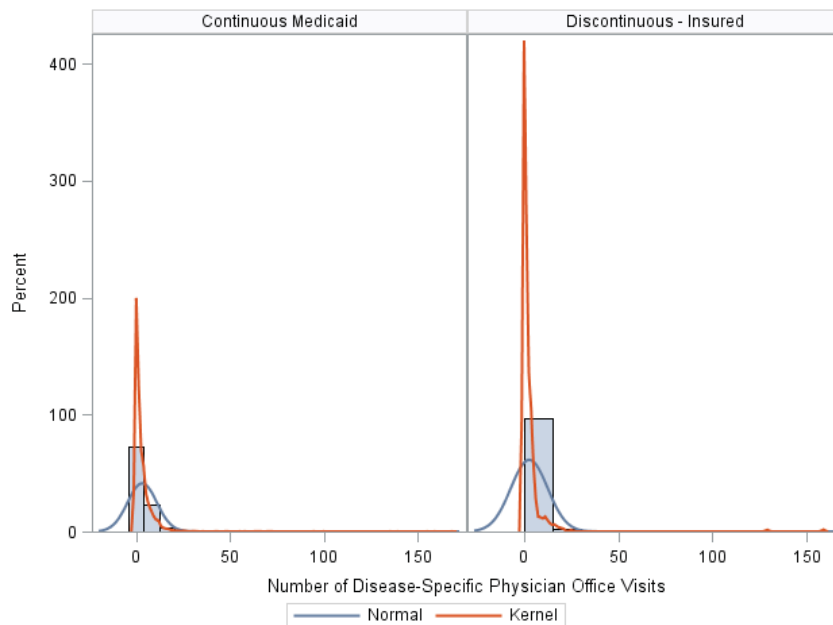


Figure B-16. Distribution of Disease-Specific Physician Office Visits among Continuous Medicaid and Discontinuous–Insured

VITA

Kiraat Munshi was born in Mumbai, India in 1984. He is a registered Pharmacist in India after having successfully completed his undergraduate education in Pharmaceutical Sciences from Manipal University in Karnataka, India in 2007. He moved to the United States to pursue his master's degree in Pharmacy Administration at St. John's University in Jamaica, NY, which he successfully obtained in 2010. In 2011 he joined the University of Tennessee Health Science Center (UTHSC) as a doctoral student in the Health Outcomes and Policy Research program, majoring in Health Policy.

During his time at UTHSC, Kiraat has made several presentations, including poster and podium, published papers in peer-reviewed journals, and been in leadership positions in his capacity as the President of the UTHSC Student Chapter of the International Society for Pharmacoeconomics and Outcomes Research (ISPOR) organization, and co-Chair of the ISPOR 2013-2014 Student Council's Education Committee. In his role as a Data Manager for the SafeMed program, Kiraat has successfully handled several responsibilities, including but not limited to, training the SafeMed staff in data entry, oversight of the SafeMed patient data and hospitalization data quality and integrity, periodic analyses of the SafeMed data for examining process measures for the dashboard run charts and quarterly reports, and coordinating with the SafeMed measurement team in obtaining Medicare and Medicaid claims data and designing the final SafeMed data analysis plan. After having successfully defended his dissertation in June 2015, Kiraat will be awarded his doctoral degree in December 2015.